

36) **Channel with memory.**

Consider the discrete memoryless channel  $Y_i = Z_i X_i$  with input alphabet  $X_i \in \{-1, 1\}$ .

- a) What is the capacity of this channel when  $\{Z_i\}$  is i.i.d. with

$$Z_i = \begin{cases} 1, & p = 0.5 \\ -1, & p = 0.5 \end{cases} ? \quad (579)$$

Now consider the channel with memory. Before transmission begins,  $Z$  is randomly chosen and fixed for all time. Thus  $Y_i = Z X_i$ .

- b) What is the capacity if

$$Z = \begin{cases} 1, & p = 0.5 \\ -1, & p = 0.5 \end{cases} ? \quad (580)$$

36) **Channel with memory solution.**

- a) This is a BSC with cross over probability 0.5, so  $C = 1 - H(p) = 0$ .  
 b) Consider the coding scheme of sending  $X^n = (1, b_1, b_2, \dots, b_{n-1})$  where the first symbol is always a zero and the rest of the  $n - 1$  symbols are  $\pm 1$  bits. For the first symbol  $Y_1 = Z$ , so the receiver knows  $Z$  exactly. After that the receiver can recover the remaining bits error free. So in  $n$  symbol transmissions  $n$  bits are sent, for a rate  $R = \frac{n-1}{n} \rightarrow 1$ . The capacity  $C$  is bounded by  $\log |\mathcal{X}| = 1$ , therefore the capacity is 1 bit per symbol.

1) **Differential entropy.** Evaluate the differential entropy  $h(X) = - \int f \ln f$  for the following:

- a) The exponential density,  $f(x) = \lambda e^{-\lambda x}$ ,  $x \geq 0$ .  
 b) The Laplace density,  $f(x) = \frac{1}{2} \lambda e^{-\lambda|x|}$ .  
 c) The sum of  $X_1$  and  $X_2$ , where  $X_1$  and  $X_2$  are independent normal random variables with means  $\mu_i$  and variances  $\sigma_i^2$ ,  $i = 1, 2$ .

1) *Differential Entropy.*

- a) Exponential distribution.

$$h(f) = - \int_0^{\infty} \lambda e^{-\lambda x} [\ln \lambda - \lambda x] dx \quad (690)$$

$$= - \ln \lambda + 1 \text{ nats.} \quad (691)$$

$$= \log \frac{e}{\lambda} \text{ bits.} \quad (692)$$

- b) Laplace density.

$$h(f) = - \int_{-\infty}^{\infty} \frac{1}{2} \lambda e^{-\lambda|x|} \left[ \ln \frac{1}{2} + \ln \lambda - \lambda|x| \right] dx \quad (693)$$

$$= - \ln \frac{1}{2} - \ln \lambda + 1 \quad (694)$$

$$= \ln \frac{2e}{\lambda} \text{ nats.} \quad (695)$$

$$= \log \frac{2e}{\lambda} \text{ bits.} \quad (696)$$

- c) Sum of two normal distributions.

The sum of two normal random variables is also normal, so applying the result derived the class for the normal distribution, since  $X_1 + X_2 \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$ ,

$$h(f) = \frac{1}{2} \log 2\pi e(\sigma_1^2 + \sigma_2^2) \text{ bits.} \quad (697)$$

- 2) **Concavity of determinants.** Let  $K_1$  and  $K_2$  be two symmetric nonnegative definite  $n \times n$  matrices. Prove the result of Ky Fan [?]:

$$|\lambda K_1 + \bar{\lambda} K_2| \geq |K_1|^\lambda |K_2|^{\bar{\lambda}}, \quad \text{for } 0 \leq \lambda \leq 1, \quad \bar{\lambda} = 1 - \lambda,$$

where  $|K|$  denotes the determinant of  $K$ .

*Hint:* Let  $\mathbf{Z} = \mathbf{X}_\theta$ , where  $\mathbf{X}_1 \sim N(0, K_1)$ ,  $\mathbf{X}_2 \sim N(0, K_2)$  and  $\theta = \text{Bernoulli}(\lambda)$ . Then use  $h(\mathbf{Z} | \theta) \leq h(\mathbf{Z})$ .

- 2) *Concavity of Determinants.* Let  $X_1$  and  $X_2$  be normally distributed  $n$ -vectors,  $\mathbf{X}_i \sim \phi_{K_i}(\mathbf{x})$ ,  $i = 1, 2$ . Let the random variable  $\theta$  have distribution  $\Pr\{\theta = 1\} = \lambda$ ,  $\Pr\{\theta = 2\} = 1 - \lambda$ ,  $0 \leq \lambda \leq 1$ . Let  $\theta$ ,  $\mathbf{X}_1$ , and  $\mathbf{X}_2$  be independent and let  $\mathbf{Z} = \mathbf{X}_\theta$ . Then  $\mathbf{Z}$  has covariance  $K_Z = \lambda K_1 + (1 - \lambda) K_2$ . However,  $\mathbf{Z}$  will not be multivariate normal. However, since a normal distribution maximizes the entropy for a given variance, we have

$$\frac{1}{2} \ln(2\pi e)^n |\lambda K_1 + (1 - \lambda) K_2| \geq h(\mathbf{Z}) \geq h(\mathbf{Z} | \theta) = \lambda \frac{1}{2} \ln(2\pi e)^n |K_1| + (1 - \lambda) \frac{1}{2} \ln(2\pi e)^n |K_2|.$$

Thus

$$|\lambda K_1 + (1 - \lambda) K_2| \geq |K_1|^\lambda |K_2|^{1-\lambda}, \quad (698)$$

as desired.

- 7) **Differential entropy bound on discrete entropy:** Let  $X$  be a discrete random variable on the set  $\mathcal{X} = \{a_1, a_2, \dots\}$  with  $\Pr(X = a_i) = p_i$ . Show that

$$H(p_1, p_2, \dots) \leq \frac{1}{2} \log(2\pi e) \left( \sum_{i=1}^{\infty} p_i i^2 - \left( \sum_{i=1}^{\infty} i p_i \right)^2 + \frac{1}{12} \right). \quad (688)$$

Moreover, for every permutation  $\sigma$ ,

$$H(p_1, p_2, \dots) \leq \frac{1}{2} \log(2\pi e) \left( \sum_{i=1}^{\infty} p_{\sigma(i)} i^2 - \left( \sum_{i=1}^{\infty} i p_{\sigma(i)} \right)^2 + \frac{1}{12} \right). \quad (689)$$

*Hint:* Construct a random variable  $X'$  such that  $\Pr(X' = i) = p_i$ . Let  $U$  be an uniform(0,1] random variable and let  $Y = X' + U$ , where  $X'$  and  $U$  are independent. Use the maximum entropy bound on  $Y$  to obtain the bounds in the problem. This bound is due to Massey (unpublished) and Willems(unpublished).

7) *Differential entropy bound on discrete entropy*

Of all distributions with the same variance, the normal maximizes the entropy. So the entropy of the normal gives a good bound on the differential entropy in terms of the variance of the random variable.

Let  $X$  be a discrete random variable on the set  $\mathcal{X} = \{a_1, a_2, \dots\}$  with

$$\Pr(X = a_i) = p_i. \quad (717)$$

$$H(p_1, p_2, \dots) \leq \frac{1}{2} \log(2\pi e) \left( \sum_{i=1}^{\infty} p_i i^2 - \left( \sum_{i=1}^{\infty} i p_i \right)^2 + \frac{1}{12} \right). \quad (718)$$

Moreover, for every permutation  $\sigma$ ,

$$H(p_1, p_2, \dots) \leq \frac{1}{2} \log(2\pi e) \left( \sum_{i=1}^{\infty} p_{\sigma(i)} i^2 - \left( \sum_{i=1}^{\infty} i p_{\sigma(i)} \right)^2 + \frac{1}{12} \right). \quad (719)$$

Define two new random variables. The first,  $X_0$ , is an integer-valued discrete random variable with the distribution

$$\Pr(X_0 = i) = p_i. \quad (720)$$

Let  $U$  be a random variable uniformly distributed on the range  $[0, 1]$ , independent of  $X_0$ . Define the continuous random variable  $\tilde{X}$  by

$$\tilde{X} = X_0 + U. \quad (721)$$

It is clear that  $H(X) = H(X_0)$ , since discrete entropy depends only on the probabilities and not on the values of the outcomes. Now

$$H(X_0) = - \sum_{i=1}^{\infty} p_i \log p_i \quad (722)$$

$$= - \sum_{i=1}^{\infty} \left( \int_i^{i+1} f_{\tilde{X}}(x) dx \right) \log \left( \int_i^{i+1} f_{\tilde{X}}(x) dx \right) \quad (723)$$

$$= - \sum_{i=1}^{\infty} \int_i^{i+1} f_{\tilde{X}}(x) \log f_{\tilde{X}}(x) dx \quad (724)$$

$$= - \int_1^{\infty} f_{\tilde{X}}(x) \log f_{\tilde{X}}(x) dx \quad (725)$$

$$= h(\tilde{X}), \quad (726)$$

since  $f_{\tilde{X}}(x) = p_i$  for  $i \leq x < i + 1$ .

Hence we have the following chain of inequalities:

$$H(X) = H(X_0) \quad (727)$$

$$= h(\tilde{X}) \quad (728)$$

$$\leq \frac{1}{2} \log(2\pi e) \text{Var}(\tilde{X}) \quad (729)$$

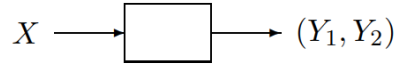
$$= \frac{1}{2} \log(2\pi e) (\text{Var}(X_0) + \text{Var}(U)) \quad (730)$$

$$= \frac{1}{2} \log(2\pi e) \left( \sum_{i=1}^{\infty} p_i i^2 - \left( \sum_{i=1}^{\infty} i p_i \right)^2 + \frac{1}{12} \right). \quad (731)$$

Since entropy is invariant with respect to permutation of  $p_1, p_2, \dots$ , we can also obtain a bound by a permutation of the  $p_i$ 's. We conjecture that a good bound on the variance will be achieved when the high probabilities are close together, i.e., by the assignment  $\dots, p_5, p_3, p_1, p_2, p_4, \dots$  for  $p_1 \geq p_2 \geq \dots$ . How good is this bound? Let  $X$  be a Bernoulli random variable with parameter  $\frac{1}{2}$ , which implies that  $H(X) = 1$ . The corresponding random variable  $X_0$  has variance  $\frac{1}{4}$ , so the bound is

$$H(X) \leq \frac{1}{2} \log(2\pi e) \left( \frac{1}{4} + \frac{1}{12} \right) = 1.255 \text{ bits.} \tag{732}$$

**2) The two-look Gaussian channel.**



Consider the ordinary Gaussian channel with two correlated looks at  $X$ , i.e.,  $Y = (Y_1, Y_2)$ , where

$$Y_1 = X + Z_1 \tag{780}$$

$$Y_2 = X + Z_2 \tag{781}$$

with a power constraint  $P$  on  $X$ , and  $(Z_1, Z_2) \sim \mathcal{N}_2(\mathbf{0}, K)$ , where

$$K = \begin{bmatrix} N & N\rho \\ N\rho & N \end{bmatrix}. \tag{782}$$

Find the capacity  $C$  for

- a)  $\rho = 1$
- b)  $\rho = 0$
- c)  $\rho = -1$

2) *The two look Gaussian channel.*

It is clear that the input distribution that maximizes the capacity is  $X \sim \mathcal{N}(0, P)$ . Evaluating the mutual information for this distribution,

$$C_2 = \max I(X; Y_1, Y_2) \quad (801)$$

$$= h(Y_1, Y_2) - h(Y_1, Y_2|X) \quad (802)$$

$$= h(Y_1, Y_2) - h(Z_1, Z_2|X) \quad (803)$$

$$= h(Y_1, Y_2) - h(Z_1, Z_2) \quad (804)$$

Now since

$$(Z_1, Z_2) \sim \mathcal{N}\left(\mathbf{0}, \begin{bmatrix} N & N\rho \\ N\rho & N \end{bmatrix}\right), \quad (805)$$

we have

$$h(Z_1, Z_2) = \frac{1}{2} \log(2\pi e)^2 |K_Z| = \frac{1}{2} \log(2\pi e)^2 N^2(1 - \rho^2). \quad (806)$$

Since  $Y_1 = X + Z_1$ , and  $Y_2 = X + Z_2$ , we have

$$(Y_1, Y_2) \sim \mathcal{N}\left(\mathbf{0}, \begin{bmatrix} P + N & P + \rho N \\ P + \rho N & P + N \end{bmatrix}\right), \quad (807)$$

and

$$h(Y_1, Y_2) = \frac{1}{2} \log(2\pi e)^2 |K_Y| = \frac{1}{2} \log(2\pi e)^2 (N^2(1 - \rho^2) + 2PN(1 - \rho)). \quad (808)$$

Hence the capacity is

$$C_2 = h(Y_1, Y_2) - h(Z_1, Z_2) \quad (809)$$

$$= \frac{1}{2} \log\left(1 + \frac{2P}{N(1 + \rho)}\right). \quad (810)$$

a)  $\rho = 1$ . In this case,  $C = \frac{1}{2} \log(1 + \frac{P}{N})$ , which is the capacity of a single look channel. This is not surprising, since in this case  $Y_1 = Y_2$ .

b)  $\rho = 0$ . In this case,

$$C = \frac{1}{2} \log\left(1 + \frac{2P}{N}\right), \quad (811)$$

which corresponds to using twice the power in a single look. The capacity is the same as the capacity of the channel  $X \rightarrow (Y_1 + Y_2)$ .

c)  $\rho = -1$ . In this case,  $C = \infty$ , which is not surprising since if we add  $Y_1$  and  $Y_2$ , we can recover  $X$  exactly.

Note that the capacity of the above channel in all cases is the same as the capacity of the channel  $X \rightarrow Y_1 + Y_2$ .