University of Illinois at Urbana-Champaign

ECE 534: RANDOM PROCESSES Fall 2007 Final Exam

Tuesday, December 11, 2007

• This is a closed-book exam. You may consult both sides of threfont size 10 or equivalent handwriting size.	e sheets of notes, typed in							
• Each part is worth 4 points.								
• Calculators, laptop computers, Palm Pilots, two-way email pagers, etc. may not be us								
• Write your answers in the space provided.								
• Please show all of your work. Answers without appropriate justified ittle credit.	tification will receive very							
Score: 1:								
Total(or points)								

Problem 1. Note the following properties of the sinc function might be useful:

- $\operatorname{sinc}(u) \triangleq \begin{cases} \frac{\sin(\pi u)}{\pi u}, & u \neq 0\\ 1, & u = 0 \end{cases}$
- $\sum_{j} \operatorname{sinc}\left(\frac{t_1 jT}{T}\right) \operatorname{sinc}\left(\frac{t_2 jT}{T}\right) = \operatorname{sinc}\left(\frac{t_2 t_1}{T}\right)$
- For any integer $k \neq 0$: $\operatorname{sinc}(kT) = 0$.

Consider a baseband random process Z(t), bandlimited to [-W, W], given by

$$Z(t) = \sum_{j} Z_{j} \operatorname{sinc}\left(\frac{t - jT}{T}\right),$$

where $T = \frac{1}{2W}$ and the random variables $\{Z_j\}$ are i.i.d and $\mathcal{N}(0,1)$.

a) Is Z a Gaussian random process?

b) Is Z a wide sense stationary (WSS) random process?

c) Is Z a stationary random process?

Now consider the random process W defined by

$$W(t) = \sum_{j} W_{2j} \operatorname{sinc}\left(\frac{t - 2jT}{T}\right) + W_{2j+1} \operatorname{sinc}\left(\frac{t - (2j + 1)T}{T}\right),$$

where the random variables indexed at even integer times, $\{W_{2j}\}$, are i.i.d. Gaussian $\mathcal{N}(0,1)$ and the random variables indexed at odd integer times, $\{W_{2j+1}\}$, are given by

$$W_{2i+1} = A_{2i}W_{2i}$$

where $\{A_{2j}\}$ are i.i.d. equiprobable ± 1 and independent of the random variables $\{W_{2j}\}$.

- d) What is the probability distribution of W_{2j+1} for any j?
- e) Is W a WSS process?

f) Is W a Gaussian process?

Problem 2.

a) Consider a Poisson process N with jumps given by times T_1, T_2, \ldots Define $T_0 = t_0 = 0$. Show that for any integer $k \geq 1$, and any $t > t_{k-1}$:

$$P(T_k > t | T_{k-1} = t_{k-1}, \dots, T_1 = t_1) = e^{-\lambda(t - t_{k-1})}.$$

(you need not write much here).

b) Suppose we are interested in characterizing the probability density of the jump times on [0, T]. Define

$$f_{N(T),T_1,\ldots,T_n}(n,t_1,\ldots,t_n) \triangleq \lim_{\Delta \to 0} \frac{P\left(T_1 \in [t_1,t_1+\Delta],\ldots,T_n \in [t_n,t_n+\Delta],N(T)=n\right)}{\Delta^n}.$$

Use part a), along with a baby Bernoulli approach splitting [0, T] into small intervals of length Δ , to show that

$$f_{N(T),T_1,...,T_n}(n,t_1,...,t_n) = \left[\prod_{i=1}^n \lambda e^{-\lambda(t_i-t_{i-1})}\right] e^{-\lambda(T-t_n)} = \lambda^n e^{-\lambda T}.$$

c) Use the Dirac delta function - along with the fact that on [0, T], any sample path counting process N_t with n jumps is completely defined by the jump times t_1, t_2, \ldots, t_n - to show that (b) can be expressed as: $f_{N(T),T_1,\ldots,T_n}(n,t_1,\ldots,t_n) = \exp\left\{\int_0^T \log \lambda dN_t - \lambda dt\right\}.$

Now suppose we have a point process that no longer has the independent increments property that is special to Poisson. Rather than there being independent coin flips in $[t, t + \Delta]$, now the probability of heads of any coin flip in the interval $[t, t + \Delta]$ is allowed to depend upon the previous coin flips in [0, t]. Define the history at time t to be what has occurred up to time t: $H_t = \{N_\tau\}_{\tau < t}$. In this more general context, let us define the conditional intensity function $\lambda(t|H_t)$ as: $\lambda(t|H_t) \triangleq \lim_{\Delta \to 0} \frac{P(N(t, t + \Delta) = 1|H_t)}{\Delta}.$

Let also assume, as for the Poisson case, that the process is regular: $\lim_{\Delta \to 0} \frac{P(N(t,t+\Delta)>1|H_t)}{\Delta} = 0$.

d) With this more general point process definition, show that $P\left(T_{k} > t_{k} \middle| T_{k-1} = t_{k-1}, \dots T_{1} = t_{1}\right) = P\left(T_{k} > t_{k} \middle| H_{t_{k-1}}\right) = \exp\left\{-\int_{t_{k-1}}^{t_{k}} \lambda\left(\tau \middle| H_{\tau}\right) d\tau\right\}$

 $\left(\text{hint: Just as for a fixed } \beta > 0, \lim_{\Delta \to 0} \left[1 - \Delta \beta \right]^{t/\Delta} = \exp(-\beta t), \text{ this can be generalized for an integrable function } g(t) \ge 0 \text{ as } \lim_{\Delta \to 0} \prod_{j=1}^{t/\Delta} \left[1 - \Delta g(j\Delta) \right] = \exp\left(- \int_0^t g(\tau) d\tau \right) \right).$

e) Show that the probability density of the jump times on [0, T] is given by

$$f_{N(T),T_{1},...,T_{n}}(n,t_{1},...,t_{n}) = \left[\prod_{j=1}^{n} \lambda\left(t_{j}|H_{t_{j}}\right)\right] \exp\left\{-\int_{0}^{T} \lambda\left(t|H_{t}\right) dt\right\}$$
$$= \exp\left\{\int_{0}^{T} \log \lambda\left(t|H_{t}\right) dN_{t} - \lambda\left(t|H_{t}\right) dt\right\}.$$

f) Assume that the process N satisfies $\lambda(t|H_t) > 0$ for any t and any H_t . Define $T_0 = 0$. Show the Time-Rescaling Theorem for Point Processes, which states that the time-rescaled random variables $\{Z_i\}$, given by

$$Z_k = \int_{T_{k-1}}^{T_k} \lambda(t|H_t)dt,$$

form the inter-arrival times of a unit-rate Poisson process. (Hint: show that for any integrable function h > 0, the function $g(x) = \int_0^x h(s)ds$ is monotonically increasing and thus invertible. Using this fact, equivalence of events, and careful reasoning involving the result of part a requires minimal space to show the result.)

Problem 3. Suppose we observe Y which is related to X through multiplicative noise:

$$Y_t = X_t(1 + N_t)$$

where X and N are real, independent, and wide sense stationary (WSS) with $E[N_t] = 0$.

a) Show that Y is wide sense stationary (WSS).

Now suppose that $S_X(\omega) = \frac{2}{1+\omega^2}$ and $S_N(\omega) = \frac{4}{4+\omega^2}$.

b) Find the (possibly noncausal) Wiener filter for X_t .

c)	Find	the	causal	Wiener	filter	for	X_t	given	observations	$\{Y_{\tau}\}_{\tau \leq t}$	

d) Suppose that X_t and N_t are jointly Gaussian random processes. Is Y_t a Gaussian process? Is Y_t a Gauss-Markov process?

Problem 4. Consider two jointly Wide Sense Stationary (jWSS) random processes Y and Z with power spectral densities S_Y , S_Z and cross-power spectral density S_{YZ} .

a) Suppose the processes Y and Z are put through a time-invariant linear transformation given by time-invariant functions h, g, k, l to produce the random processes U and V:

$$\left[\begin{array}{c} U \\ V \end{array}\right] = \left[\begin{array}{c} h & k \\ g & l \end{array}\right] \left[\begin{array}{c} Y \\ Z \end{array}\right].$$

That is:

$$U_t = \int_{-\infty}^{\infty} h(t-s)Y_s ds + \int_{-\infty}^{\infty} k(t-v)Z_v dv$$

$$V_{\tau} = \int_{-\infty}^{\infty} g(\tau-a)Y_a da + \int_{-\infty}^{\infty} l(\tau-b)Z_b db.$$

Show that

$$R_{UV}(t,\tau) = h * \tilde{g} * R_Y(t-\tau) + k * \tilde{g} * R_{ZY}(t-\tau) + h * \tilde{l} * R_{YZ}(t-\tau) + k * \tilde{l} * R_Z(t-\tau)$$

$$R_U(t,\tau) = h * \tilde{h} * R_Y(t-\tau) + h * \tilde{k} * R_{YZ}(t-\tau) + k * \tilde{h} * R_{ZY}(t-\tau) + k * \tilde{k} * R_Z(t-\tau)$$

$$R_V(t,\tau) = g * \tilde{g} * R_Y(t-\tau) + g * \tilde{l} * R_{YZ}(t-\tau) + l * \tilde{g} * R_{ZY}(t-\tau) + l * \tilde{l} * R_Z(t-\tau),$$

which equivalently means in the power spectral domain that

$$S_{UV} = HG^*S_Y + KG^*S_{ZY} + HL^*S_{YZ} + KL^*S_Z$$

$$S_U = |H|^2S_Y + HK^*S_{YZ} + KH^*S_{ZY} + |K|^2S_Z$$

$$S_V = |G|^2S_Y + GL^*S_{YZ} + LG^*S_{ZY} + |L|^2S_Z.$$

(Hint: remember that $\tilde{h}(v) \triangleq h^*(-v)$. Also, you need not show the structure of both R_U and R_V . You can argue the other via symmetry.).

(more room for calculations for 4a... Remember - in case you are in a time crunch - every question part is only worth 4 points...)

b) Suppose we would like to estimate a random process X that is jointly wide sense stationary (jWSS) with Y and Z. Consider using a linear estimator consisting of the LTI systems c and d of the form

$$\hat{X}_t = \int_{-\infty}^{\infty} c(t-s)Y_s + \int_{-\infty}^{\infty} d(t-v)Z_v dv$$

to optimally estimate X_t (possibly non-causally) from $\{Y_s\}_{\infty < s < \infty}$ and $\{Z_s\}_{\infty < s < \infty}$ in the linear MMSE sense. Find the functions c and d. Feel free to express this in terms of their Fourier transforms, $C(\omega)$ and $D(\omega)$.

c) Now suppose we would like to estimate X_t causally, using $\{Y_s\}_{s < t}$ and $\{Z_s\}_{s < t}$. Again, consider using LTI systems c and d (this time causal) to do so:

$$\hat{X}_t = \int_{-\infty}^t c(t-s)Y_s + \int_{-\infty}^t d(t-v)Z_v dv.$$

Use the orthogonality principle to express integral equations that must be satisfied on the LTI systems c and d. Also express this in the Fourier domain.

d) Now suppose that $S_Y = S_Z$ is rational and that $S_{YZ} = S_{ZY}$ is rational. Find the optimal causal filters $C(\omega)$ and $D(\omega)$. (hint: consider taking an intermediate step, consider innovations sequences, and consider using the result of part a).