

Week 11 and 12.

Independence of random variables.

• X & Y are independent if & only if $F_{X,Y}(u,v) = F_X(u)F_Y(v) \quad \forall u,v \in \mathbb{R}$

• If X & Y are discrete type, they are independent if & only if $P_{X,Y}(u,v) = P_X(u)P_Y(v) \quad \forall u,v \in \mathbb{R}$

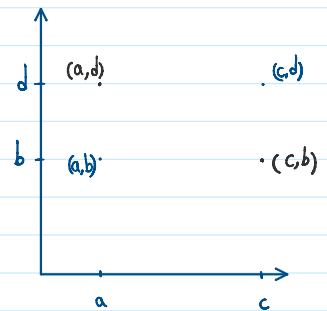
• If X & Y are jointly continuous, they are independent if & only if $f_{X,Y}(u,v) = f_X(u)f_Y(v) \quad \forall u,v \in \mathbb{R}$

Determining from a joint pdf whether independence holds

To ensure independence: for every $u \in \mathbb{R}$, either $f_X(u) = 0$ or $f_{Y|X}(v|u) = \frac{f_{X,Y}(u,v)}{f_X(u)}$ only depends of v , i.e., $f_{Y|X}(v|u) = f_Y(v)$.

To roll-out independence: if X & Y are independent then support of $f_{X,Y}$ should satisfy swap property, i.e.,

$$\begin{aligned} (a,b) \in \text{support of } f_{X,Y} & \Rightarrow (a,d) \in \text{support of } f_{X,Y} \\ (c,d) \in \text{support of } f_{X,Y} & \Rightarrow (b,c) \in \text{support of } f_{X,Y} \end{aligned}$$



Function of pair of jointly continuous random variables.

$Z = g(X,Y)$ & the joint pdf of (X,Y) is given by $f_{X,Y}$. Find distribution of Z .

Step 1: identify type of Z (continuous-type or discrete-type) and support of Z (c s.t. $f_Z(c) > 0$)

if Z is continuous-type: Step 2: find its cdf. $F_Z(c) = P(Z \leq c) = P(g(X,Y) \leq c) = \iint_{(u,v): g(u,v) \leq c} f_{X,Y}(u,v) du dv$

Step 3: take derivative of $F_Z(c)$ to derive pdf of Z :

$$f_Z(c) = \frac{dF_Z(c)}{dc}$$

if Z is discrete-type: Step 2: calculate pmf of Z . $P(Z=k) = P(g(X,Y)=k) = \iint_{(u,v): g(u,v)=k} f_{X,Y}(u,v) du dv$

Sum of random variables.

Suppos that X & Y are discrete-type, and integer valued. Let $S = X+Y$.

$$P_{S,1} \quad P(S=1) \quad \tau \quad P(X=1, Y=1) \quad \tau \quad P_{S,1} \quad (1,1)$$

$$P_S(k) = P(S=k) = \sum_j P(X=j, Y=k-j) = \sum_j P_{X,Y}(j, k-j)$$

If X and Y are independent, then

$$P_S(k) = P(S=k) = \sum_j P_X(j) P_Y(k-j) =: P_X * P_Y(k) \text{ if } X \text{ \& } Y \text{ are independent \& } S=X+Y$$

③ Suppose that X and Y are jointly continuous. Let $S=X+Y$.

$$f_S(c) = \int_{-\infty}^{+\infty} f_{X,Y}(u, c-u) du = \int_{-\infty}^{+\infty} f_{X,Y}(c-v, v) dv$$

If X and Y are independent:

$$f_S(c) = \int_{-\infty}^{+\infty} f_X(u) f_Y(c-u) du = \int_{-\infty}^{+\infty} f_X(c-v) f_Y(v) dv = f_X * f_Y(c) \text{ if } X \text{ \& } Y \text{ are independent \& } S=X+Y$$

Week 13:

joint pdf of functions of random variables.

Suppose that X & Y are jointly continuous random variables with joint pdf $f_{X,Y}$.

Suppose that $W=g_1(X,Y)$ and $Z=g_2(X,Y)$. Define $g: \mathbb{R}^2 \rightarrow \mathbb{R}^2$ as follows.

We can write.

$$g\left(\begin{pmatrix} u \\ v \end{pmatrix}\right) = \begin{pmatrix} g_1(u,v) \\ g_2(u,v) \end{pmatrix}$$

$$\begin{pmatrix} W \\ Z \end{pmatrix} = g\left(\begin{pmatrix} X \\ Y \end{pmatrix}\right)$$

Suppose that (X,Y) is in $u-v$ plane & (W,Z) is in $\alpha-\beta$ plane.

Let $J(u,v)$ denote the Jacobian matrix of g at point (u,v) .

$$J(u,v) = \begin{bmatrix} \frac{\partial g_1(u,v)}{\partial u} & \frac{\partial g_1(u,v)}{\partial v} \\ \frac{\partial g_2(u,v)}{\partial u} & \frac{\partial g_2(u,v)}{\partial v} \end{bmatrix}$$

Proposition: Suppose that $\begin{pmatrix} W \\ Z \end{pmatrix} = g\left(\begin{pmatrix} X \\ Y \end{pmatrix}\right)$, where $\begin{pmatrix} X \\ Y \end{pmatrix}$ has pdf $f_{X,Y}$, and g is one to one mapping from support of $f_{X,Y}$ to \mathbb{R}^2 . Suppose that the Jacobian matrix J of g exists, is continuous, and has nonzero determinant everywhere. Then for all (α, β) in the support of $f_{W,Z}$ we have

$$f_{W,Z}(\alpha, \beta) = \frac{1}{|\det(J)|} f_{X,Y}(g^{-1}(\alpha, \beta))$$

Important remark: The jacobian matrix J depends on point, i.e., $J = J(u,v)$; however $f_{W,Z}(\alpha, \beta)$ is in terms of α & β . You should calculate J in terms of u and v , and then write it in terms of α and β .

Example:

$$\begin{pmatrix} W \\ Z \end{pmatrix} = A \begin{pmatrix} X \\ Y \end{pmatrix}, \quad A = \begin{pmatrix} a & b \\ c & d \end{pmatrix}. \quad A \text{ is invertible}$$

$$f_{W,Z}(\alpha, \beta) = \frac{1}{|\det(A)|} f_{X,Y}(A^{-1}(\alpha, \beta))$$

Correlation, Covariance, Correlation Coefficient:

Let X and Y be random variables with finite second moments.

the correlation: $E[XY]$

the covariance: $\text{Cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$

the correlation coefficient: $\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}} = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y}$

Scaling and linearity of Covariance:

(i) $\text{Cov}(X+Y, U+V) = \text{Cov}(X, U) + \text{Cov}(X, V) + \text{Cov}(Y, U) + \text{Cov}(Y, V)$

(ii) $\text{Cov}(aX+c, cY+d) = ac \text{Cov}(X, Y)$

Important def. & relations.

① uncorrelated, negatively correlated, positively correlated

② $\text{Var}(X) = \text{Cov}(X, X)$. In particular if X and Y are uncorrelated

$$\text{Var}(X+Y) = \text{Cov}(X+Y, X+Y) = \text{Cov}(X, X) + \text{Cov}(Y, Y) = \text{Var}(X) + \text{Var}(Y)$$

③ If X & Y are independent, they are uncorrelated. The reverse does not hold in general.

(Note: If X & Y are jointly gaussian & uncorrelated, they are independent.)

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Week 14 & 15.

Estimation & Minimum Mean Square Error (MMSE).

① Sample mean & variance of data set

Given n independent & identically distributed random variables $X_1, X_2, X_3, \dots, X_n$, the sample mean & variance are unbiased estimator and are given as follows.

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i, \quad \hat{\sigma}^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \hat{\mu})^2$$

② Minimum mean square error estimators of random variable Y , after observing random variable X

. Constant estimator: $\hat{S}^* = E[Y]$, $MMSE = \sigma_Y^2$

. Linear estimator: $L^*(X) = \hat{E}[Y|X] = \mu_Y + \sigma_Y \rho_{X,Y} \left(\frac{X - \mu_X}{\sigma_X} \right)$, $MMSE = \sigma_Y^2 (1 - \rho_{X,Y}^2)$

. Unconstrained estimator: $g^*(X) = E[Y|X]$, $MMSE = \sigma_Y^2 - E[(E[Y|X])^2]$

Limit Theorems.

③ Law of large numbers

Suppose that X_1, X_2, \dots, X_n are uncorrelated, each one has the same mean μ , and their variance is bounded by C . Let $S_n = X_1 + X_2 + \dots + X_n$. Then, for any $\delta > 0$, we have

$$P\left(\left|\frac{S_n}{n} - \mu\right| \geq \delta\right) = P\left(\left\{\omega \in \Omega : \left|\frac{S_n(\omega)}{n} - \mu\right| \geq \delta\right\}\right) \leq \frac{C}{n\delta^2} \xrightarrow{\text{as } n \rightarrow \infty} 0$$

④ Central limit theorem

Suppose that X_1, X_2, \dots, X_n are independent, identically distributed, each with mean μ and variance σ^2 . Let $S_n = X_1 + X_2 + \dots + X_n$. Then, for any c , we have

$$\lim_{n \rightarrow \infty} P\left(\left\{\frac{S_n - n\mu}{\sqrt{n\sigma^2}} \leq c\right\}\right) = \lim_{n \rightarrow \infty} P\left(\left\{\tilde{S}_n \leq c\right\}\right) = \Phi(c)$$

where $\tilde{S}_n = \frac{S_n - n\mu}{\sqrt{n\sigma^2}} = \frac{S_n - E[S_n]}{\sqrt{\text{Var}(S_n)}}$ is the standardized version of S_n .

Jointly Gaussian random variables

We say random variables X and Y are jointly gaussian random variables if any linear combination of X and Y are gaussian random variables, i.e., for any $a, b \in \mathbb{R}$, $aX + bY$ is a gaussian random variables.

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X and Y are gaussian random variables, i.e., for any $a, b \in \mathbb{R}$, $aX + bY$ is a gaussian random variable.

① joint distribution of gaussian random variables. Suppose that X, Y are non-degenerate, i.e., they are not linearly related. Then their joint distribution is given by:

$$f_{X,Y}(u,v) = \frac{1}{2\pi \sigma_X \sigma_Y \sqrt{1-\rho^2}} \exp\left(-\frac{\left(\frac{u-\mu_X}{\sigma_X}\right)^2 + \left(\frac{v-\mu_Y}{\sigma_Y}\right)^2 - 2\rho\left(\frac{u-\mu_X}{\sigma_X}\right)\left(\frac{v-\mu_Y}{\sigma_Y}\right)}{2(1-\rho^2)}\right)$$

where

$$\mu_X = E[X], \sigma_X^2 = \text{Var}(X)$$

$$\mu_Y = E[Y], \sigma_Y^2 = \text{Var}(Y)$$

$$\rho = \rho_{X,Y} = \frac{\text{Cov}(X,Y)}{\sigma_X \sigma_Y}$$

This is called bivariate normal distribution.

② class of bivariate normal pdfs are preserved via linear transformation.

$$\begin{pmatrix} X \\ Y \end{pmatrix} \text{ has bivariate distribution} \Rightarrow A \begin{pmatrix} X \\ Y \end{pmatrix} \text{ has bivariate distribution}$$

This same as in definition of jointly Gaussian random variables.

$$X \text{ \& } Y \text{ are jointly Gaussian} \Rightarrow aX+bY \text{ and } cX+dY \text{ are jointly Gaussian}$$

③ If $\rho=0$ then $f_{X,Y}(u,v) = f_X(u)f_Y(v)$. Hence uncorrelated jointly gaussian random variables are independent

④ If X and Y are independent, then $\rho_{X,Y} = 0$ and their joint distribution $f_{X,Y}(u,v) = f_X(u)f_Y(v)$ is bivariate normal distribution. Hence, independent gaussian random variables are jointly gaussian.

⑤ If X and Y are gaussian, it does not mean that they are jointly gaussian.

⑥ If X and Y are jointly gaussian, then X has $\mathcal{N}(\mu_X, \sigma_X^2)$ distribution and Y has $\mathcal{N}(\mu_Y, \sigma_Y^2)$ distribution.

⑦ If X and Y are jointly Gaussian, for estimation of Y from X , $L^*(X) = g^*(X)$. Equivalently, $E[Y|X] = \hat{E}[Y|X]$, i.e., the best linear estimator of Y using X , is the best unconstrained estimator.

$$g^*(X) = E[Y|X] = \hat{E}[Y|X] = \mu_Y + \sigma_Y \rho_{X,Y} \left(\frac{X-\mu_X}{\sigma_X}\right)$$

⑧ If X and Y are jointly Gaussian, the conditional distribution of Y given $X=u$ is $\mathcal{N}(\hat{E}[Y|X=u], \sigma_e^2)$

⑧ If X and Y are jointly Gaussian, the conditional distribution of Y given $X=u$ is $N(\hat{E}[Y|X=u], \sigma_e^2)$ where σ_e^2 is MSE for $\hat{E}[Y|X]$.

$$\sigma_e^2 = \sigma_Y^2 (1 - \rho_{X,Y}^2)$$

10. [30 points] (3 points per answer) → Spring 2015
 In order to discourage guessing, 3 points will be deducted for each incorrect answer (no penalty or gain for blank answers). A net negative score will reduce your total exam score.

(a) Suppose X and Y are independent, continuous-type random variables. $W = \max(X, Y)$, $Z = \min(X, Y)$.

TRUE FALSE

$F_W(t) = F_X(t)F_Y(t)$

$F_Z(t) = (1 - F_X(t))(1 - F_Y(t))$

$f_Z(t) = f_X(t)(1 - F_Y(t)) + f_Y(t)(1 - F_X(t))$

(b) Let A, B, C be independent events with $0 < P(A), P(B), P(C) < 1$.

TRUE FALSE

$P(AC|B) = P(AC|B^c)$

$P(AB|C) \leq P(C|AB)P(AB)$

(c) Let X, Y be two independent normal random variables with $\mu_X = \mu_Y$ and with $\sigma_X > \sigma_Y$.

TRUE FALSE

$\mathbb{P}(Y \leq X) > 1/2$.

$\mathbb{P}(X = Y) > 0$

(d) X and Y are jointly distributed discrete random variables. They are uncorrelated if:

TRUE FALSE

$\text{Var}(X + Y) = \text{Var}(X - Y)$

$E[XY] = 0$

$P(X = u, Y = v) = P(X = u)P(Y = v)$ for every pair (u, v)

7. [7+16+7 points] Suppose $X \sim N(1, 1)$ and $Y \sim N(1, 4)$ are independent Gaussian random variables. Define the random variables $Z = 2X + Y$ and $W = X - Y$. ~> spring 2019

- (a) Find the unconstrained MMSE estimator of Y given X , and the resulting MSE.
- (b) Find the unconstrained MMSE estimator of Z given W , and the resulting MSE.
- (c) If instead $W = X - aY$ for some real a and $E[Z|W] = E[Z]$, find a .

10. [10+10 points] Suppose X and Y are jointly Gaussian with the following parameters: $\mu_x = 0$, $\mu_y = 0$, $\sigma_x^2 = 1$, $\sigma_y^2 = 2^2$, $\rho = 1/8$.

- (a) Find $P\{2X + Y \geq 3\}$. Express your answer using the Q function.
- (b) Find $E[Y^2|X = 2]$

7. [15 points] The two parts of the problem are unrelated. ~> spring 2018

- (a) Suppose a fair die is rolled 100 times. What is a rough approximation to the sum of the numbers showing, based on the law of large numbers?
- (b) Suppose each of 1200 real numbers are rounded to the nearest integer and then added. Assume the individual roundoff errors are independent and uniformly distributed over the interval $[-0.5, 0.5]$. The random variable equal to the sum is denoted by S . Using the CLT, find the approximate probability that the absolute value of the sum of the errors is greater than 5.

3. [8+4 points] Suppose X and Y are independent random variables with joint pdf:

$$f_{X,Y}(u, v) = \begin{cases} 2e^{-u}e^{-2v} & \text{if } u \geq 0, v \geq 0 \\ 0 & \text{else.} \end{cases} \quad \text{~> spring 2019}$$

- (a) Find the joint pdf of $S = X + Y$ and $W = Y - X$.
- (b) Are S and W independent? Explain.