Joint Distribution Functions, Independent Random Variables

Probability with Engineering Applications

Lecture 16

Professor Ravi K. Iyer

Dept. of Electrical and Computer Engineering

University of Illinois at Urbana Champaign

Announcements

- Midterm next Tuesday, October 22
 11:00am 12:20pm, in class
 - All topics covered in Lectures 1 to 15
 - Homework 1-6, In-class projects 1-3, and Mini-Projects 1-2
- Be on time, exam starts at 11:00am sharp.
- You are allowed to bring only one 8"x11" sheet of notes
- Review Session Today, 5:00pm 7:00pm, CSL 141

Additional TA Office hours on Friday, 2:00pm – 5pm, CSL 249.

Today's Topics

- Quick Review on Joint Distribution Functions
 - Example
- Independence of Random Variables
- Review of Material for the Midterm Exam

Joint Distribution Functions

- We have concerned ourselves with the probability distribution of a single random variable
- Often interested in probability statements concerning two or more random variables
- Define, for any two random variables *X* and *Y*, the *joint* cumulative probability distribution function of *X* and *Y* by

$$F(a,b) = P\{X \le a, Y \le b\}, -\infty < a, b < \infty$$

• The distribution of *X* can be obtained from the joint distribution of *X* and *Y* as follows:

$$F_{X}(a) = P\{X \le a\}$$

$$= P\{X \le a, Y < \infty\}$$

$$= F(a, \infty)$$

- Similarly, $F_Y(b) = P\{Y \le b\} = F(\infty, b)$ Where X and Y are both discrete random variables it is convenient to define the *joint* probability mass function of X and Y by $p(x, y) = P\{X = x, Y = y\}$
- Probability mass function of X $p_X(x) = \sum_{y:p(x,y)>0} p(x,y)$

$$p_{Y}(y) = \sum_{x:p(x,y)>0} p(x,y)$$

 We say that X and Y are jointly continuous defined for all real x and y

$$P\{X \in A, Y \in B\} = \int_{B} \int_{A} f(x, y) dx dy$$

 Called the *joint probability density function* of X and Y. The probability density of X

$$P\{X \in A\} = P\{X \in A, Y \in (-\infty, \infty)\}$$
$$= \int_{-\infty}^{\infty} \int_{A} f(x, y) dx dy$$
$$= \int_{A} f_{X}(x) dx$$

 $f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$ is thus the probability density function of X

• Similarly the probability density function of Y is $f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$ because

$$F(a,b) = P(X \le a, Y \le b) = \int_{-\infty}^{a} \int_{-\infty}^{b} f(x,y) dy dx$$

 Proposition: if X and Y are random variables and g is a function of two variables, then

$$E[g(X,Y)] = \sum_{y} \sum_{x} g(x,y) p(x,y)$$
$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f(x,y) dx dy$$

• For example, if g(X,Y)=X+Y, then, in the continuous case

$$E[X,Y] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x+y)f(x,y)dx dy$$
$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xf(x,y)dx dy + \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} yf(x,y)dx dy$$
$$= E[X] + E[Y]$$

- Where the first integral is evaluated by using the foregoing Proposition with g(x,y)=x and the second with g(x,y)=y
- In the discrete case E[aX + bY] = aE[X] + bE[Y]
- Joint probability distributions may also be defined for n random variables. If $X_1, X_2, ..., X_n$ are n random variables, then for any n constants $a_1, a_2, ..., a_n$

$$E[a_1X_1 + a_2X_2 + ...a_nX_n] = a_1E[X_1] + a_2E[X_2] + ... + a_nE[X_n]$$

Example 3

Suppose that the joint probability mass function of *X* and *Y* is

$$P(X = i, Y = j) = {j \choose i} e^{-2\lambda} \lambda^j / j!, \quad 0 \le i \le j$$

- (a) Find the probability mass function of Y.
- (b) Find the probability mass function of X.
- (c) Find the probability mass function of Y X.

Example 3 (Cont'd)

 $P(Y=j) = \sum_{i=0}^{J} {j \choose i} e^{-2\lambda} \lambda^{j} / j!$ Marginal PDF of Y: a) $= e^{-2\lambda} \frac{\lambda^j}{j!} \sum_{i=0}^{j} {j \choose i} 1^i 1^{j-i}$ $P(X=i) = \sum_{i=i}^{\infty} {j \choose i} e^{-2\lambda} \lambda^{j} / j!$ b) Marginal PDF of *X:* $= \frac{1}{i!}e^{-2\lambda}\sum_{i=i}^{\infty} \frac{\lambda^{j}}{(j-i)!}$ $= \frac{\lambda^i}{i!} e^{-2\lambda} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!}$

Example 3 (Cont'd)

c) We first calculate the joint density function of X and Y-X

$$P(X = i, Y - X = k) = P(X = i, Y = k + i)$$

$$= {\binom{k+i}{i}} e^{-2\lambda} \frac{\lambda^{k+i}}{(k+i)!}$$

$$= e^{-2\lambda} \frac{\lambda^k}{k!} \frac{\lambda^i}{i!}.$$

• Then summing up with respect to i, we get the marginal distribution of Y - X, which is for k:

$$P(Y - X = k) = \sum_{i=0}^{\infty} P(X = i, Y - X = k)$$

$$= e^{-2\lambda} \frac{\lambda^k}{k!} \sum_{i=0}^{\infty} \frac{\lambda^i}{i!}$$

$$= e^{-2\lambda} \frac{\lambda^k}{k!} e^{-\lambda}$$

$$= e^{-\lambda} \frac{\lambda^k}{k!}.$$

Iyer - Lecture 16

Independent Random Variables

- We define two random variables X and Y to be independent if: $F(x,y) = F_X(x)F_Y(y), -\infty < x < \infty, -\infty < y < \infty$
- Independence of random variables X and Y implies that their joint CDF factors into the product of the marginal CDFs.
- Applies to all types of random variables
- In case X and Y are discrete, the preceding definition of independence is equivalent to $p(x, y) = P_X(x)P_Y(y)$
- If *X* and *Y* are continuous, the preceding definition of independence is equivalent to the condition

$$f(x, y) = f_X(x) f_Y(y), -\infty < x < \infty, -\infty < y < \infty$$

- assuming that f(x,y) exists.
- The joint distribution of X and Y when one of them is a discrete random variable while the other is a continuous random variable

Independent Random Variables Cont'd

If X is discrete and Y is continuous their independence becomes:

$$P(X = x, Y \le y) = p_X(x)f_Y(y)$$
, all x and y

- The definition of joint distribution, joint density, and independence of two random variables can be easily generalized to a set of n random variables, $X_1, X_2, ..., X_n$.
- Example (Independent R.V.)
- Assume that the lifetime X and the brightness Y of a light bulb are being modeled as continuous random variables. Let the joint pdf be given by $f(x,y) = \lambda_1 \lambda_2 e^{-(\lambda_1 x + \lambda_2 y)}, 0 < x < \infty, 0 < y < \infty$
- This is known as the bivariate exponential density.
- The marginal density of X is $f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$ $= \int_{0}^{\infty} \lambda_1 \lambda_2 e^{-(\lambda_1 x + \lambda_2 y)} dy$ $= \lambda_1 e^{-\lambda_1 x}, 0 < x < \infty$

Independent Random Variables Cont'd

- Similarly $f_Y(y) = \lambda_2 e^{-\lambda_2 y}, 0 < y < \infty$
- It follows that *X* and *Y* are independent random variables:

$$f(xy) = f(x)f(y)$$

The joint distribution function can be computed to be

$$F(x,y) = \int_{-\infty}^{x} \int_{-\infty}^{y} f(u,v) \, dv \, du$$

$$= \int_{0}^{x} \int_{0}^{y} \lambda_{1} \lambda_{2} e^{-(\lambda_{1}u + \lambda_{2}v)} \, dv \, du$$

$$= (1 - e^{-\lambda_{1}x})(1 - e^{-\lambda_{2}y}), 0 < x < \infty, 0 < y < \infty.$$

Basic Concepts:

- **Random experiment** is an experiment the outcome of which is not certain
- Sample Space (S) is the totality of the possible outcomes of a random experiment
- Discrete (countable) sample space is a sample space which is either
 - finite, i.e., the set of all possible outcomes of the experiment is finite
 - *countably infinite, i.e.*, the set of all outcomes can be put into a one-to-one correspondence with the natural numbers
- Continuous sample space is a sample space for which all elements constitute a continuum, such as all the points on a line, all the points in a plane
- An *event* is a collection of certain sample points, i.e., a subset of the sample space
 - *Universal event* is the entire sample space S
 - The null set \emptyset is a null or impossible event

Algebra of Events

- The *intersection* of E_1 and E_2 is given by:
 - $E_1 \cap E_2 = \{s \in S \mid s \text{ is an element of both } E_1 \text{ and } E_2\}$
- The *union* E_1 and E_2 is given by:
 - $E_1 \cup E_2 = \{ s \in S \mid \text{either } s \in E_1 \text{ or } s \in E_2 \text{ or both} \}$
- In general: $|E_1 \cup E_2| \le |E_1| + |E_2|$
 - where |A| = the number of elements in the set (Cardinality)
- Definition of union and intersection extend to any finite number of sets:

$$\bigcap_{i=1}^{n} E_{i} = E_{1} \cup E_{2} \cup E_{3} \cup ... \cup E_{n}$$

$$\bigcap_{i=1}^{n} E_{i} = E_{1} \cap E_{2} \cap E_{3} \cap ... \cap E_{n}$$

• Mutually exclusive or disjoint events are two events for which

$$A \cap B = \emptyset$$

- A list of events A₁, A₂, ..., A_n is said to be
 - composed of *mutually exclusive events* iff:

$$A_i \cap A_j = \begin{cases} A_i & \text{if } i = j \\ \emptyset & \text{otherwise} \end{cases}$$

- collectively **exhaustive** iff: $A_1 \cup A_2 \cup ... \cup A_n = S$

Probability Axioms

- Let S be a sample space of a random experiment and P(A) be the probability of the event A
- The probability function P(.) must satisfy the three following axioms:
- (A1) For any event A, $P(A) \ge 0$ (probabilities are nonnegative real numbers)
- (A2) P(S) = 1 (probability of a certain event, an event that must happen is equal 1)
- (A3) $P(A \cup B) = P(A) + P(B)$, whenever A and B are mutually exclusive events, i.e., $A \cap B = \emptyset$ (probability function must be additive)
- (A3') For any countable sequence of events $A_1, A_2, ..., A_n$..., that are mutually exclusive (that is $A_i \cap A_k = \emptyset$ whenever $j \neq k$)

$$P(\bigcup_{n=1}^{\infty}A_n)=\sum_{n=1}^{\infty}P(A_n)$$

- (Ra) For any event A, $P(\overline{A}) = 1 P(A)$
- (Rb) If \varnothing is the impossible event, then $P(\varnothing) = 0$
- **(Rc)** If A and B are any events, not necessarily mutually exclusive, then

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

• (Rd)(generalization of Rc) If A₁, A₂, ..., A_n are any events, then

$$P(\bigcup_{i=1}^{n} A_i) = P(A_1 \cup A_2 \cup ... \cup A_n) = \sum_{i} P(A_i) - \sum_{1 \le i < j \le n} P(A_i \cap A_j)$$

$$+ \sum_{1 \le i < j < k \le n} P(A_i \cap A_j \cap A_k) + \dots + (-1)^{n-1} P(A_1 \cap A_2 \cap \dots \cap A_n)$$

where the successive sums are over all possible events, pairs of events, triples of events, and so on.

(Can prove this relation by induction (see class web site))

Combinatorial Problems

- Permutations with replacement:
 - Ordered samples of size k, with replacement P(n, k)
- Permutations without replacement
 - Ordered Samples of size k, without replacement

$$n(n-1)....(n-k+1) = \frac{n!}{(n-k)!}$$
 $k = 1, 2, ..., n$

- Combinations
 - Unordered sample of size k, without replacement

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$

Binomial Theorem

$$(x+y)^n = \sum_{k=0}^n \binom{n}{k} x^k y^{n-k}$$

 Conditional Probability of A given B (P(A|B)) defines the conditional probability of the event A given that the event B occurs and is given by:

$$P(A \mid B) = \frac{P(A \cap B)}{P(B)}$$

if $P(B) \neq 0$ and is undefined otherwise.

 A rearrangement of the above definition gives the following multiplication rule (MR)

$$P(A \cap B) = \begin{cases} P(B)P(A \mid B) & \text{if } P(B) \neq 0 \\ P(A)P(B \mid A) & \text{if } P(A) \neq 0 \\ 0 & \text{otherwise} \end{cases}$$

Or:

$$P(A \cap B) = P(A)P(B|A) = P(B)P(A|B)$$

- Theorem of Total Probability
- Any event A can be partitioned into two disjoint subsets:

$$A = (A \cap B) \cup (A \cap \overline{B})$$

• Then:

$$P(A) = P(A \cap B) \cup P(A \cap \overline{B})$$
$$= P(A \mid B)P(B) + P(A \mid \overline{B})P(\overline{B})$$

In general:

$$P(A) = \sum_{i=1}^{n} P(A \mid B_i) P(B_i)$$

Bayes Formula:

$$P(B_{j} | A) = \frac{P(B_{j} \cap A)}{P(A)} = \frac{P(A | B_{j})P(B_{j})}{\sum_{i} P(A | B_{i})P(B_{i})}$$

- Independence of Events:
- Two events A and B are independent if and only if:

$$P(A|B)=P(A)$$

Or events A and be are said to be independent if:

$$P(A \cap B) = P(A)P(B)$$

Reliability Applications:

- Recovery blocks
- Series and parallel systems:

• Series System:
$$R_s = P$$
 ("The system is functioning properly.")
$$= P(A_1 \cap A_2 \cdots \cap A_n)$$

$$= P(A_1)P(A_2)\cdots P(A_n)$$

$$= \prod_{i=1}^n R_i$$
 (2.1)

• Parallel System:

$$R_p = 1 - F_p = 1 - \prod_{i=1}^{n} (1 - R_i)$$

• In general:
$$R_{sp} = \prod_{i=1}^{n} [1 - (1 - R_i)^{n_i}]$$

Bayes formula in example non series parallel systems

Bernoulli Trials

- The probability of obtaining exactly k successes in n trials is:

$$p(k) = {n \choose k} p^k q^{n-k}$$
 $k = 0, 1, ..., n$

NMR System:

$$R_{m|n} = P("m or more components functioning properly")$$

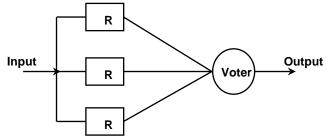
$$= P(\bigcup_{i=m}^{n} \{"exactly i components functioning properly"\})$$

$$= \sum_{i=m}^{n} P("exactly i components functioning properly")$$

$$= \sum_{i=m}^{n} p(i) = \sum_{i=m}^{n} {n \choose i} R^{i} (1-R)^{n-i}$$

• TMR System:

$$R_{TMR} = 3R^2 - 2R^3$$
 Input



Random Variables:

- A random variable X on a sample space S is a function X: $S \to \Re$ that assigns a real number X(s) to each sample point $s \in S$.
- Discrete random variables: The random variables which are either finite or countable.
 - Bernoulli
 - Binomial
 - Poisson
 - Geometric
 - Modified Geometric
- Continuous random variables: The random variables that take on a continuum of possible values.
 - Uniform
 - Normal
 - Exponential

- Cumulative distribution function (cdf) (or distribution function) $F(\cdot)$ of a random variable X is defined for any real number $b, -\infty < b < \infty$, by $F(b) = P\{X \le b\}$
- F(b) denotes the probability that the random variable X takes on a value that is less than or equal to b.
- Some properties of cdf F are:
 - i. F(b) is a non-decreasing function of b,
 - ii. $\lim_{b\to +\infty} F(b) = F(\infty) = 1$,
 - iii. $\lim_{b\to -\infty} F(b) = F(-\infty) = 0$.
- All probability questions about X can be answered in terms of $cdf \ F(\cdot)$. e.g.: $P\{a \le X \le b\} = F(b) F(a)$ for all a < b

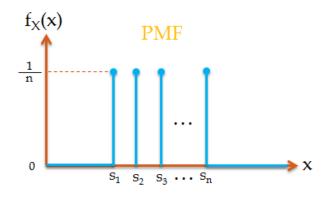
- Discrete Random Variables:
 - Probability mass function (pmf):

$$p(a) = P\{X = a\}$$

• Properties:

$$\begin{cases} p(x_i) > 0, & i = 1,2,... \\ p(x) = 0, & \text{for other values of } x \end{cases}$$

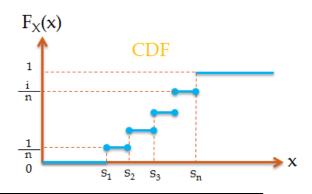
$$\sum_{i=1}^{\infty} p(x_i) = 1$$



Cumulative distribution function (CDF):

$$F(a) = \sum_{all \ x_{i \le a}} p(x_i)$$

A stair step function

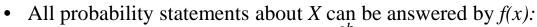


- Continuous Random Variables:
 - Probability distribution function (pdf):

$$P\{X \in B\} = \int_{B} f(x) dx$$

• Properties:

$$1 = P\{X \in (-\infty, \infty)\} = \int_{-\infty}^{\infty} f(x) dx$$

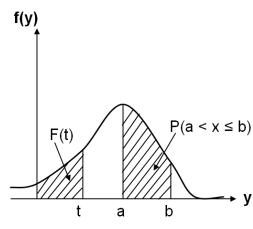


$$P\{a \le X \le b\} = \int_a^b f(x)dx$$
$$P\{X = a\} = \int_a^a f(x)dx = 0$$

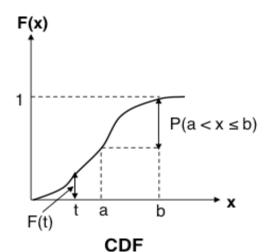
Cumulative distribution function (CDF):

$$F_{x}(x) = P(X \le x) = \int_{-\infty}^{x} f_{x}(t)dt , -\infty < x < \infty$$

- Properties: $\frac{d}{da}F(a) = f(a)$
- A continuous function



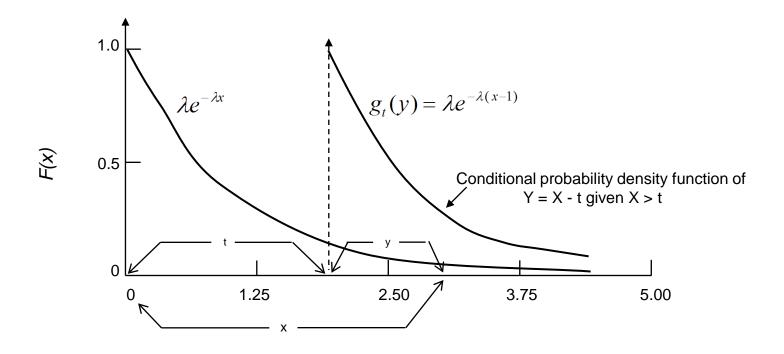
pdf



• Summary of important distributions:

Distribution	PDF or PMF	Mean	Variance
Bernoulli(p)	$\begin{cases} p, & \text{if } x = 1 \\ 1 - p, & \text{if } x = 0. \end{cases}$	p	p(1-p)
Binomial(n, p)	$\binom{n}{k} p^k (1-p)^{n-k}$ for $0 \le k \le n$	np	npq
Geometric(p)	$p(1-p)^{k-1}$ for $k = 1, 2, \dots$	$\frac{1}{p}$	$\frac{1-p}{p^2}$
$Poisson(\lambda)$	$e^{-\lambda}\lambda^x/x!$ for $k=1,2,\ldots$	λ	λ
Uniform(a,b)	$\frac{1}{b-a} \ \forall x \in (a,b)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
$Gaussian(\mu, \sigma^2)$	$\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	μ	σ^2
$Exponential(\lambda)$	$\lambda e^{-\lambda x} \ x \ge 0, \lambda > 0$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$

Memory-less property of Exponential



- Expectation:
 - The Discrete Case $E[X] = \sum_{x:p(x)>0} xp(x)$
 - The Continuous Case $E[X] = \int_{-\infty}^{\infty} xf(x)dx$
- Expectation of function of a random variable

$$E[g(X)] = \sum_{x:p(x)>0} g(x)p(x)$$

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x)dx$$

Corollary:

$$E[aX + b] = aE[X] + b$$

Moments:

$$E[Y] = E[\phi(X)] = \begin{cases} \sum_{i} \phi(x_i) p_X(x_i), & \text{if } X \text{ is discrete,} \\ \int_{-\infty}^{\infty} \phi(x) f_X(x) dx, & \text{if } X \text{ is continuous,} \end{cases}$$

$$Y = X^n \implies E[X^n], \quad n \ge 1 \qquad \mu_k = E[(X - E[X])^k]$$

Variance:

$$Var[X] = \mu^{2} = \sigma^{2}x = \begin{cases} \sum_{i} (x_{i} - E[X])^{2} p(x_{i}) & \text{if } X \text{ is discrete} \\ \int_{-\infty}^{\infty} (x - E[X])^{2} f(x) dx & \text{if } X \text{ is continuous} \end{cases}$$

$$Var(X) = E[X^{2}] - (E[X])^{2}$$

Corollary:

$$Var[aX + b] = a^2 Var(X)$$

- Functions of random variables
 - Examples shown in the class
- Reliability function and Mean time to failure:
 - Let T denote the time to failure or lifetime of a component in the system

$$R(t) = P\{T > t\} = 1 - F(t)$$

$$E[T] = \int_{0}^{\infty} R(t)dt = MTTF$$

If the component lifetime is exponentially distributed, then:

$$E[T] = \int_{0}^{\infty} e^{-\lambda t} dt = \frac{1}{\lambda}$$

$$Var[T] = \int_{0}^{\infty} 2te^{-\lambda t} dt - \frac{1}{\lambda^{2}} = \frac{1}{\lambda^{2}}$$

Joint distribution functions:

For any two random variables X and Y, the *joint cumulative probability* distribution function of X and Y by:

$$F(a,b) = P\{X \le a, Y \le b\}, -\infty < a, b < \infty$$

- Discrete:
 - The *joint probability mass function* of *X* and *Y*

$$p(x, y) = P\{X = x, Y = y\}$$

• Marginal PMFs of X and Y:

$$p_{X}(x) = \sum_{y:p(x,y)>0} p(x,y)$$
$$p_{Y}(y) = \sum_{x:p(x,y)>0} p(x,y)$$

- Joint distribution functions:
 - Continuous:
 - The *joint probability density function* of *X* and *Y*:

$$P\{X \in A, Y \in B\} = \int_{B} \int_{A} f(x, y) dx dy$$

• Marginal PDFs of X and Y:

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

Relation between joint CDF and PDF

$$F(a,b) = P(X \le a, Y \le b) = \int_{-\infty}^{a} \int_{-\infty}^{b} f(x,y) dy dx$$

Function of two joint random variables

$$E[g(X,Y)] = \sum_{y} \sum_{x} g(x,y) p(x,y)$$
$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f(x,y) dx dy$$

• For example, if g(X,Y)=X+Y, then, in the continuous case

$$E[X+Y] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x+y)f(x,y)dx dy$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xf(x,y)dx dy + \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} yf(x,y)dx dy$$

$$= E[X] + E[Y]$$

• Independent Random Variables: Two random variables *X* and *Y* are said to be independent if:

$$F(x, y) = F_X(x)F_Y(y), -\infty < x < \infty, -\infty < y < \infty$$

If X and Y are continuous:

$$f(x, y) = f_X(x)f_Y(y), -\infty < x < \infty, -\infty < y < \infty$$

• If *X* is discrete and *Y* is continuous:

$$P(X = x, Y \le y) = p_X(x)f_Y(y)$$
, all x and y