

# Chapter 14

## Introduction to Randomized Algorithms: QuickSort and QuickSelect

OLD CS 473: Fundamental Algorithms, Spring 2015  
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### 14.1 Introduction to Randomized Algorithms

### 14.2 Introduction

#### 14.2.0.1 Randomized Algorithms

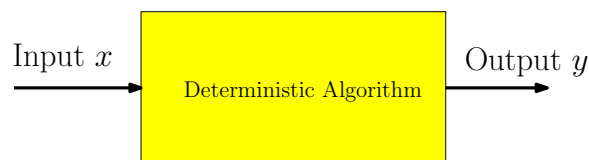
#### 14.2.0.2 Example: Randomized QuickSort

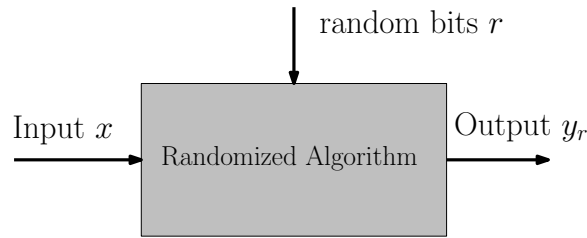
QuickSort **Hoare [1962]**

- (A) Pick a pivot element from array
- (B) Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
- (C) Recursively sort the subarrays, and concatenate them.

Randomized **QuickSort**

- (A) Pick a pivot element *uniformly at random* from the array
- (B) Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
- (C) Recursively sort the subarrays, and concatenate them.





### 14.2.0.3 Example: Randomized Quicksort

Recall: **QuickSort** can take  $\Omega(n^2)$  time to sort array of size  $n$ .

**Theorem 14.2.1.** *Randomized **QuickSort** sorts a given array of length  $n$  in  $O(n \log n)$  expected time.*

**Note:** On *every* input randomized **QuickSort** takes  $O(n \log n)$  time in expectation. On *every* input it may take  $\Omega(n^2)$  time with some small probability.

### 14.2.0.4 Example: Verifying Matrix Multiplication

Problem Given three  $n \times n$  matrices  $A, B, C$  is  $AB = C$ ?

Deterministic algorithm:

- (A) Multiply  $A$  and  $B$  and check if equal to  $C$ .
- (B) Running time?  $O(n^3)$  by straight forward approach.  $O(n^{2.37})$  with fast matrix multiplication (complicated and impractical).

### 14.2.0.5 Example: Verifying Matrix Multiplication

Problem Given three  $n \times n$  matrices  $A, B, C$  is  $AB = C$ ?

Randomized algorithm:

- (A) Pick a random  $n \times 1$  vector  $r$ .
- (B) Return the answer of the equality  $ABr = Cr$ .
- (C) Running time?  $O(n^2)$ !

**Theorem 14.2.2.** *If  $AB = C$  then the algorithm will always say YES. If  $AB \neq C$  then the algorithm will say YES with probability at most  $1/2$ . Can repeat the algorithm 100 times independently to reduce the probability of a false positive to  $1/2^{100}$ .*

### 14.2.0.6 Why randomized algorithms?

- (A) Many applications: algorithms, data structures and CS.
- (B) In some cases only known algorithms are randomized or randomness is provably necessary.
- (C) Often randomized algorithms are (much) simpler and/or more efficient.
- (D) Several deep connections to mathematics, physics etc.
- (E) ...
- (F) Lots of fun!

### 14.2.0.7 Where do I get random bits?

**Question:** Are true random bits available in practice?

- (A) Buy them!
- (B) CPUs use physical phenomena to generate random bits.
- (C) Can use pseudo-random bits or semi-random bits from nature. Several fundamental unresolved questions in complexity theory on this topic. Beyond the scope of this course.
- (D) In practice pseudo-random generators work quite well in many applications.
- (E) The model is interesting to think in the abstract and is very useful even as a theoretical construct. One can *derandomize* randomized algorithms to obtain deterministic algorithms.

### 14.2.0.8 Average case analysis vs Randomized algorithms

**Average case analysis:**

- (A) Fix a deterministic algorithm.
- (B) Assume inputs comes from a probability distribution.
- (C) Analyze the algorithm's *average* performance over the distribution over inputs.

**Randomized algorithms:**

- (A) Algorithm uses random bits in addition to input.
- (B) Analyze algorithms *average* performance over the given input where the average is over the random bits that the algorithm uses.
- (C) On each input behaviour of algorithm is random. Analyze worst-case over all inputs of the (average) performance.

## 14.3 Basics of Discrete Probability

### 14.3.0.9 Discrete Probability

We restrict attention to finite probability spaces.

**Definition 14.3.1.** A discrete probability space is a pair  $(\Omega, \mathbf{Pr})$  consists of finite set  $\Omega$  of **elementary events** and function  $p : \Omega \rightarrow [0, 1]$  which assigns a probability  $\mathbf{Pr}[\omega]$  for each  $\omega \in \Omega$  such that  $\sum_{\omega \in \Omega} \mathbf{Pr}[\omega] = 1$ .

**Example 14.3.2.** An unbiased coin.  $\Omega = \{H, T\}$  and  $\mathbf{Pr}[H] = \mathbf{Pr}[T] = 1/2$ .

**Example 14.3.3.** A 6-sided unbiased die.  $\Omega = \{1, 2, 3, 4, 5, 6\}$  and  $\mathbf{Pr}[i] = 1/6$  for  $1 \leq i \leq 6$ .

### 14.3.1 Discrete Probability

#### 14.3.1.1 And more examples

**Example 14.3.4.** A biased coin.  $\Omega = \{H, T\}$  and  $\mathbf{Pr}[H] = 2/3, \mathbf{Pr}[T] = 1/3$ .

**Example 14.3.5.** Two independent unbiased coins.  $\Omega = \{HH, TT, HT, TH\}$  and  $\Pr[HH] = \Pr[TT] = \Pr[HT] = \Pr[TH] = 1/4$ .

**Example 14.3.6.** A pair of (highly) correlated dice.

$\Omega = \{(i, j) \mid 1 \leq i \leq 6, 1 \leq j \leq 6\}$ .

$\Pr[i, i] = 1/6$  for  $1 \leq i \leq 6$  and  $\Pr[i, j] = 0$  if  $i \neq j$ .

### 14.3.1.2 Events

**Definition 14.3.7.** Given a probability space  $(\Omega, \Pr)$  an **event** is a subset of  $\Omega$ . In other words an event is a collection of elementary events. The probability of an event  $A$ , denoted by  $\Pr[A]$ , is  $\sum_{\omega \in A} \Pr[\omega]$ .

**Definition 14.3.8.** The **complement event** of an event  $A \subseteq \Omega$  is the event  $\Omega \setminus A$  frequently denoted by  $\bar{A}$ .

## 14.3.2 Events

### 14.3.2.1 Examples

**Example 14.3.9.** A pair of independent dice.  $\Omega = \{(i, j) \mid 1 \leq i \leq 6, 1 \leq j \leq 6\}$ .

(A) Let  $A$  be the event that the sum of the two numbers on the dice is even.

Then  $A = \{(i, j) \in \Omega \mid (i + j) \text{ is even}\}$ .

$\Pr[A] = |A|/36 = 1/2$ .

(B) Let  $B$  be the event that the first die has 1. Then  $B = \{(1, 1), (1, 2), (1, 3), (1, 4), (1, 5), (1, 6)\}$ .

$\Pr[B] = 6/36 = 1/6$ .

### 14.3.2.2 Independent Events

**Definition 14.3.10.** Given a probability space  $(\Omega, \Pr)$  and two events  $A, B$  are **independent** if and only if  $\Pr[A \cap B] = \Pr[A] \Pr[B]$ . Otherwise they are dependent. In other words  $A, B$  independent implies one does not affect the other.

**Example 14.3.11.** Two coins.  $\Omega = \{HH, TT, HT, TH\}$  and  $\Pr[HH] = \Pr[TT] = \Pr[HT] = \Pr[TH] = 1/4$ .

(A)  $A$  is the event that the first coin is heads and  $B$  is the event that second coin is tails.  $A, B$  are independent.

(B)  $A$  is the event that the two coins are different.  $B$  is the event that the second coin is heads.  $A, B$  independent.

## 14.3.3 Independent Events

### 14.3.3.1 Examples

**Example 14.3.12.**  $A$  is the event that both are not tails and  $B$  is event that second coin is heads.  $A, B$  are dependent.

## 14.3.4 Union bound

14.3.4.1 The probability of the union of two events, is  $\leq$  the probability of the sum of their probabilities.

**Lemma 14.3.13.** For any two events  $\mathcal{E}$  and  $\mathcal{F}$ , we have that  $\Pr[\mathcal{E} \cup \mathcal{F}] \leq \Pr[\mathcal{E}] + \Pr[\mathcal{F}]$ .

*Proof:* Consider  $\mathcal{E}$  and  $\mathcal{F}$  to be a collection of elementary events (which they are). We have

$$\begin{aligned}\Pr[\mathcal{E} \cup \mathcal{F}] &= \sum_{x \in \mathcal{E} \cup \mathcal{F}} \Pr[x] \\ &\leq \sum_{x \in \mathcal{E}} \Pr[x] + \sum_{x \in \mathcal{F}} \Pr[x] = \Pr[\mathcal{E}] + \Pr[\mathcal{F}].\end{aligned}$$

■

## 14.3.4.2 Random Variables

**Definition 14.3.14.** Given a probability space  $(\Omega, \Pr)$  a (real-valued) random variable  $X$  over  $\Omega$  is a function that maps each elementary event to a real number. In other words  $X : \Omega \rightarrow \mathbb{R}$ .

**Example 14.3.15.** A 6-sided unbiased die.  $\Omega = \{1, 2, 3, 4, 5, 6\}$  and  $\Pr[i] = 1/6$  for  $1 \leq i \leq 6$ .

(A)  $X : \Omega \rightarrow \mathbb{R}$  where  $X(i) = i \bmod 2$ .

(B)  $Y : \Omega \rightarrow \mathbb{R}$  where  $Y(i) = i^2$ .

**Definition 14.3.16.** A **binary random variable** is one that takes on values in  $\{0, 1\}$ .

## 14.3.4.3 Indicator Random Variables

Special type of random variables that are quite useful.

**Definition 14.3.17.** Given a probability space  $(\Omega, \Pr)$  and an event  $A \subseteq \Omega$  the indicator random variable  $X_A$  is a binary random variable where  $X_A(\omega) = 1$  if  $\omega \in A$  and  $X_A(\omega) = 0$  if  $\omega \notin A$ .

**Example 14.3.18.** A 6-sided unbiased die.  $\Omega = \{1, 2, 3, 4, 5, 6\}$  and  $\Pr[i] = 1/6$  for  $1 \leq i \leq 6$ . Let  $A$  be the even that  $i$  is divisible by 3. Then  $X_A(i) = 1$  if  $i = 3, 6$  and 0 otherwise.

## 14.3.4.4 Expectation

**Definition 14.3.19.** For a random variable  $X$  over a probability space  $(\Omega, \Pr)$  the **expectation** of  $X$  is defined as  $\sum_{\omega \in \Omega} \Pr[\omega] X(\omega)$ . In other words, the expectation is the average value of  $X$  according to the probabilities given by  $\Pr[\cdot]$ .

**Example 14.3.20.** A 6-sided unbiased die.  $\Omega = \{1, 2, 3, 4, 5, 6\}$  and  $\Pr[i] = 1/6$  for  $1 \leq i \leq 6$ .

(A)  $X : \Omega \rightarrow \mathbb{R}$  where  $X(i) = i \bmod 2$ . Then  $\mathbf{E}[X] = 1/2$ .

(B)  $Y : \Omega \rightarrow \mathbb{R}$  where  $Y(i) = i^2$ . Then  $\mathbf{E}[Y] = \sum_{i=1}^6 \frac{1}{6} \cdot i^2 = 91/6$ .

### 14.3.4.5 Expectation

**Proposition 14.3.21.** For an indicator variable  $X_A$ ,  $\mathbf{E}[X_A] = \mathbf{Pr}[A]$ .

*Proof:*

$$\begin{aligned}\mathbf{E}[X_A] &= \sum_{y \in \Omega} X_A(y) \mathbf{Pr}[y] \\ &= \sum_{y \in A} 1 \cdot \mathbf{Pr}[y] + \sum_{y \in \Omega \setminus A} 0 \cdot \mathbf{Pr}[y] \\ &= \sum_{y \in A} \mathbf{Pr}[y] \\ &= \mathbf{Pr}[A].\end{aligned}$$

■

### 14.3.4.6 Linearity of Expectation

**Lemma 14.3.22.** Let  $X, Y$  be two random variables (not necessarily independent) over a probability space  $(\Omega, \mathbf{Pr})$ . Then  $\mathbf{E}[X + Y] = \mathbf{E}[X] + \mathbf{E}[Y]$ .

*Proof:*

$$\begin{aligned}\mathbf{E}[X + Y] &= \sum_{\omega \in \Omega} \mathbf{Pr}[\omega] (X(\omega) + Y(\omega)) \\ &= \sum_{\omega \in \Omega} \mathbf{Pr}[\omega] X(\omega) + \sum_{\omega \in \Omega} \mathbf{Pr}[\omega] Y(\omega) = \mathbf{E}[X] + \mathbf{E}[Y].\end{aligned}$$

■

**Corollary 14.3.23.**  $\mathbf{E}[a_1 X_1 + a_2 X_2 + \dots + a_n X_n] = \sum_{i=1}^n a_i \mathbf{E}[X_i]$ .

## 14.4 Analyzing Randomized Algorithms

### 14.4.0.7 Types of Randomized Algorithms

Typically one encounters the following types:

- (A) **Las Vegas randomized algorithms:** for a given input  $x$  output of algorithm is always correct but the running time is a random variable. In this case we are interested in analyzing the *expected* running time.
- (B) **Monte Carlo randomized algorithms:** for a given input  $x$  the running time is deterministic but the output is random; correct with some probability. In this case we are interested in analyzing the *probability* of the correct output (and also the running time).
- (C) Algorithms whose running time and output may both be random variables.

### 14.4.0.8 Analyzing Las Vegas Algorithms

*Deterministic* algorithm  $Q$  for a problem  $\Pi$ :

- (A) Let  $Q(x)$  be the time for  $Q$  to run on input  $x$  of length  $|x|$ .
- (B) Worst-case analysis: run time on worst input for a given size  $n$ .

$$T_{wc}(n) = \max_{x:|x|=n} Q(x).$$

*Randomized* algorithm  $R$  for a problem  $\Pi$ :

- (A) Let  $R(x)$  be the time for  $Q$  to run on input  $x$  of length  $|x|$ .
- (B)  $R(x)$  is a random variable: depends on random bits used by  $R$ .
- (C)  $\mathbf{E}[R(x)]$  is the expected running time for  $R$  on  $x$
- (D) Worst-case analysis: expected time on worst input of size  $n$

$$T_{rand-wc}(n) = \max_{x:|x|=n} \mathbf{E}[Q(x)].$$

### 14.4.0.9 Analyzing Monte Carlo Algorithms

*Randomized* algorithm  $M$  for a problem  $\Pi$ :

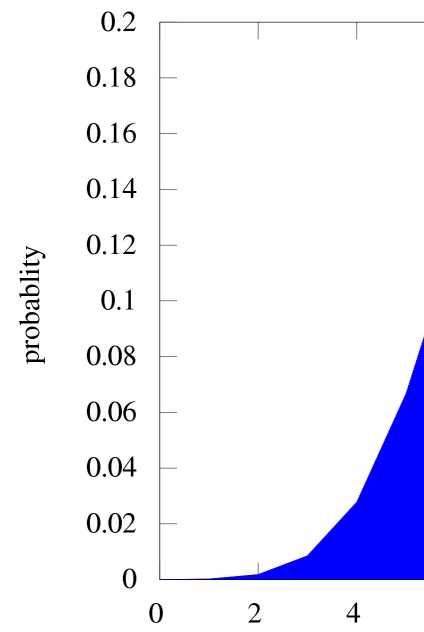
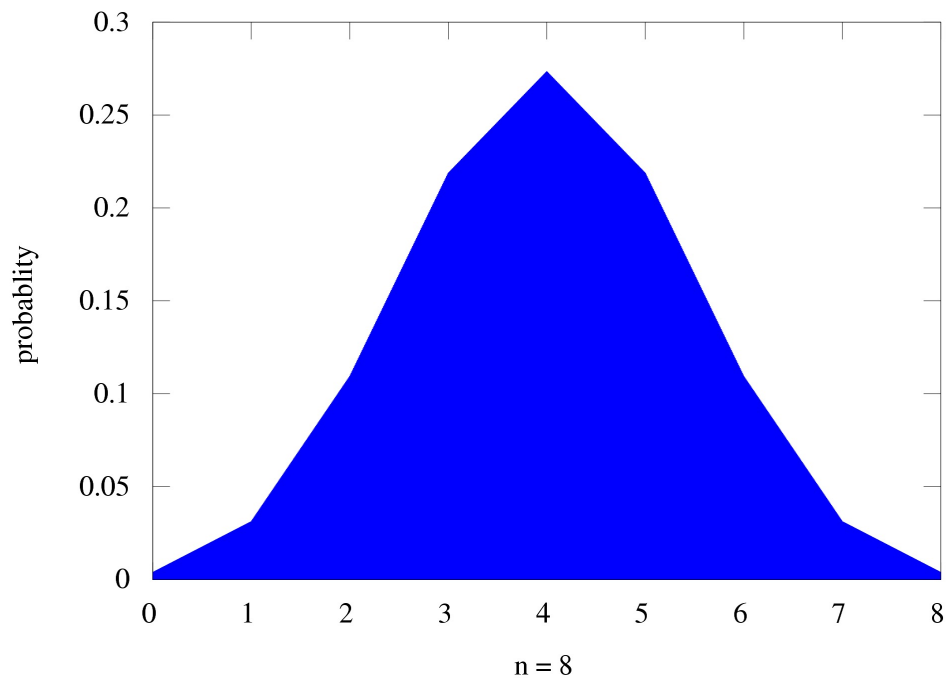
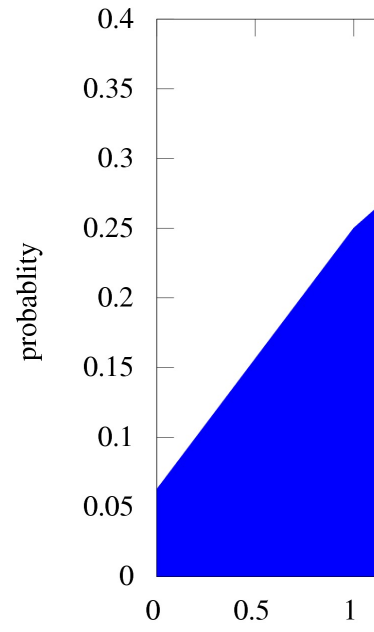
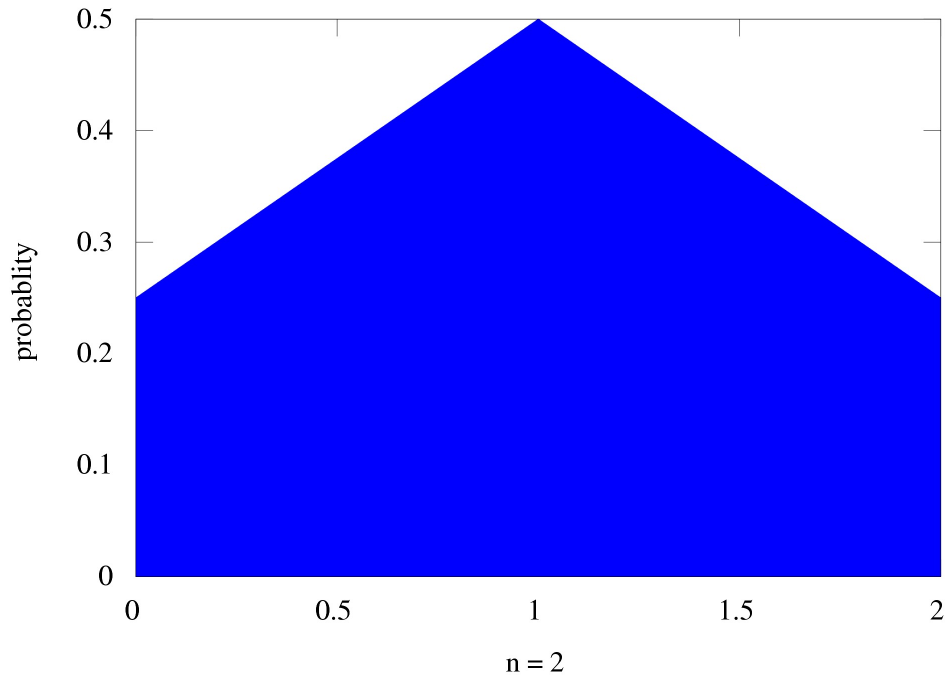
- (A) Let  $M(x)$  be the time for  $M$  to run on input  $x$  of length  $|x|$ . For Monte Carlo, assumption is that run time is deterministic.
- (B) Let  $\mathbf{Pr}[x]$  be the probability that  $M$  is correct on  $x$ .
- (C)  $\mathbf{Pr}[x]$  is a random variable: depends on random bits used by  $M$ .
- (D) Worst-case analysis: success probability on worst input

$$P_{rand-wc}(n) = \min_{x:|x|=n} \mathbf{Pr}[x].$$

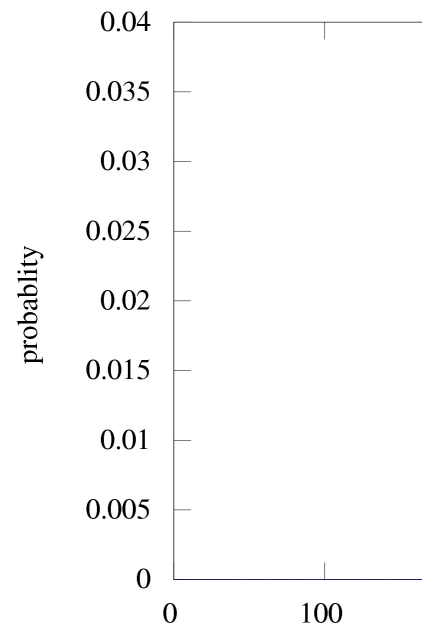
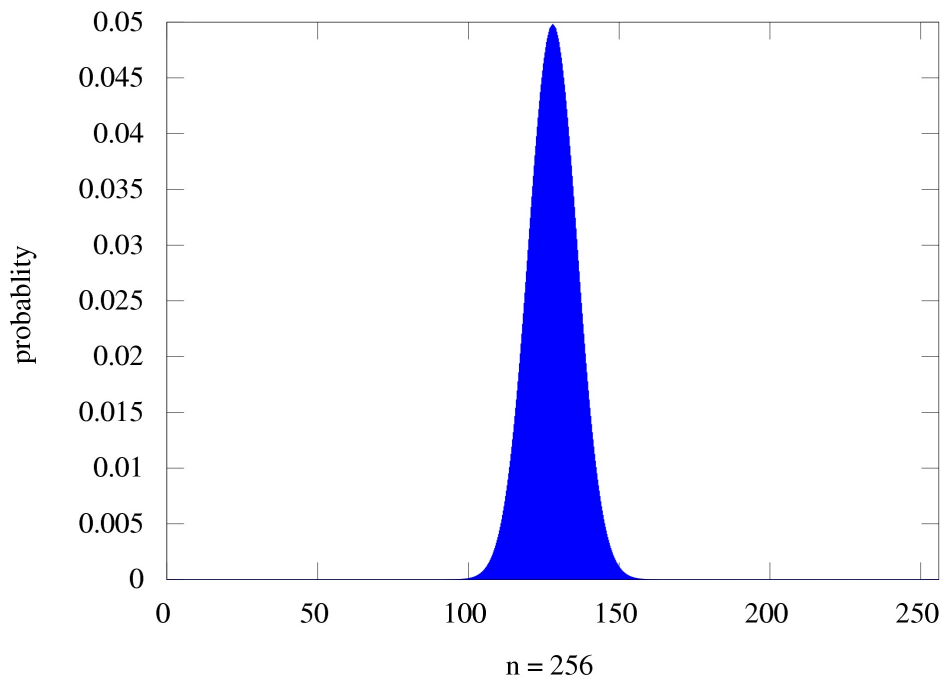
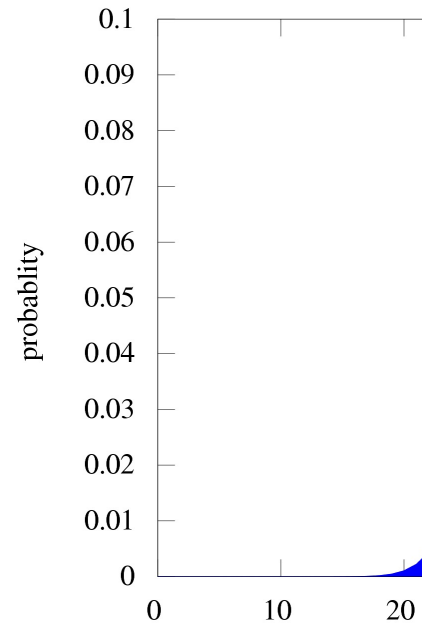
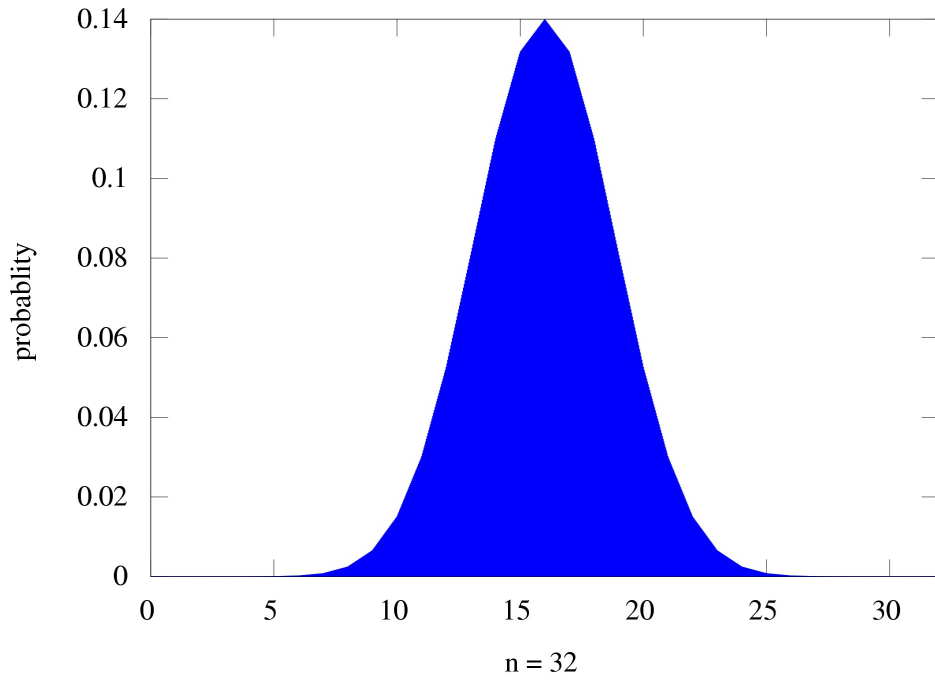
## 14.5 Why does randomization help?

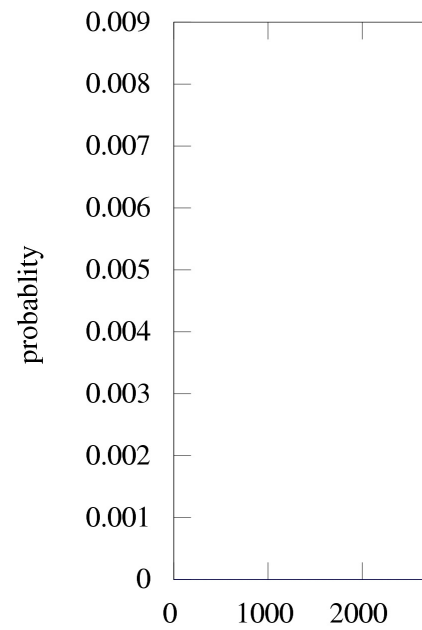
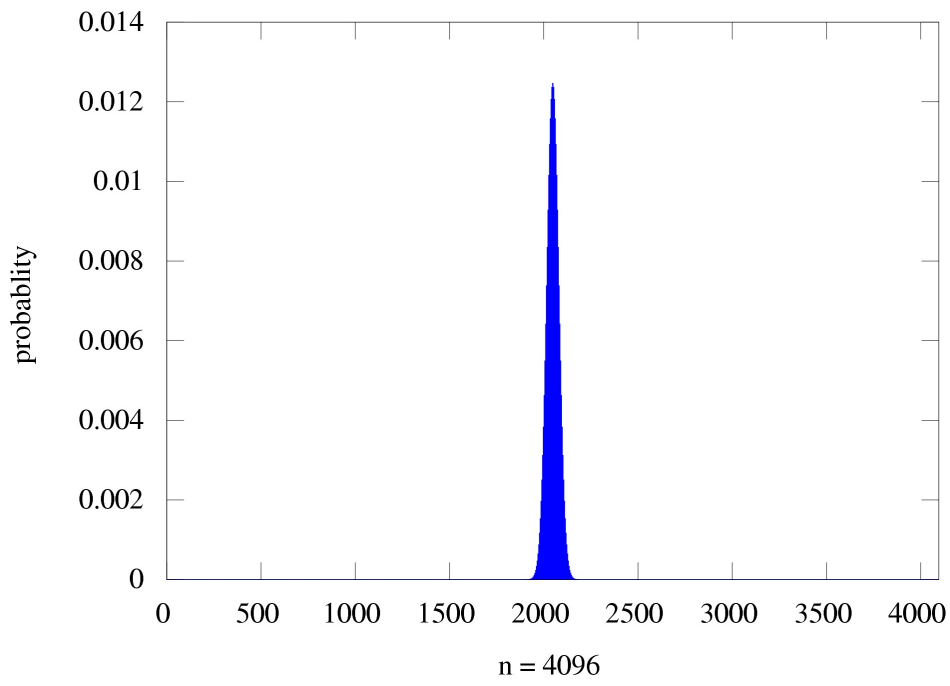
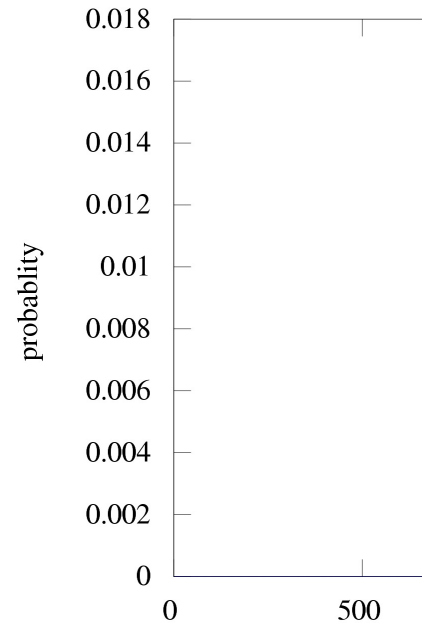
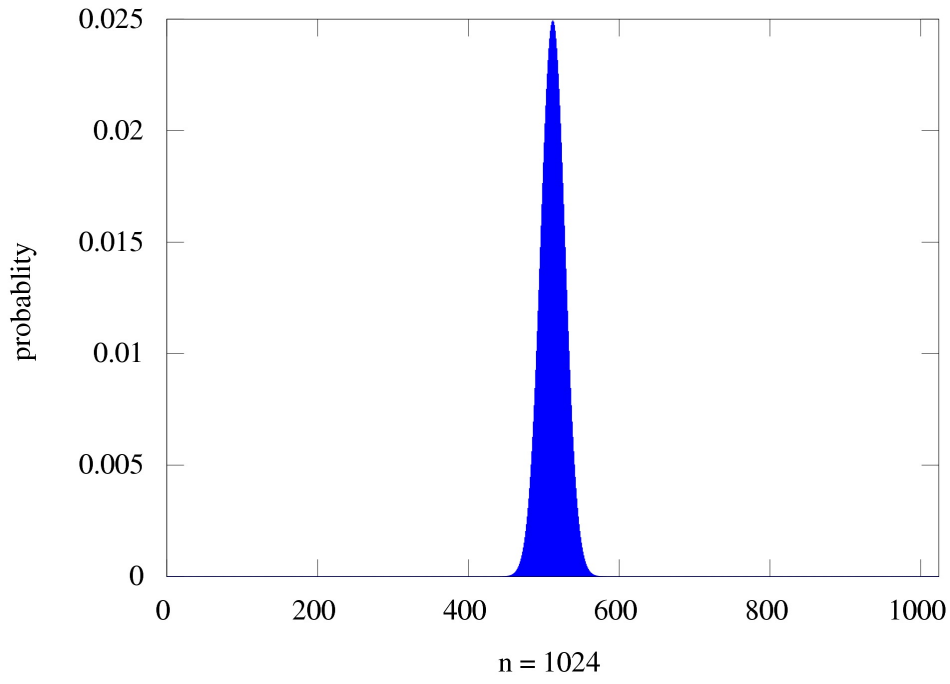
### 14.5.0.10 Massive randomness.. Is not that random.

Consider flipping a fair coin  $n$  times independently, head given 1, tail gives zero. How many heads? ...we get a binomial distribution.

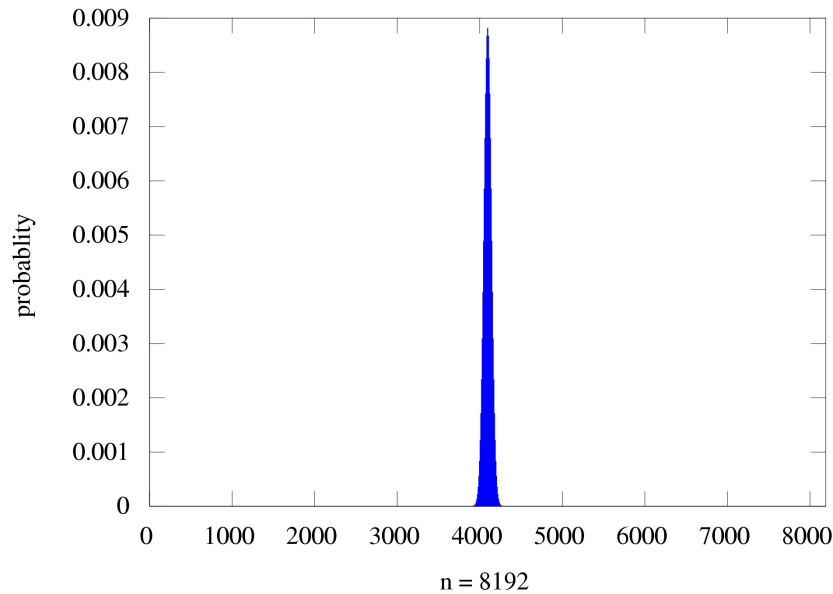








### 14.5.0.11 Massive randomness.. Is not that random.



This is known as *concentration of mass*.

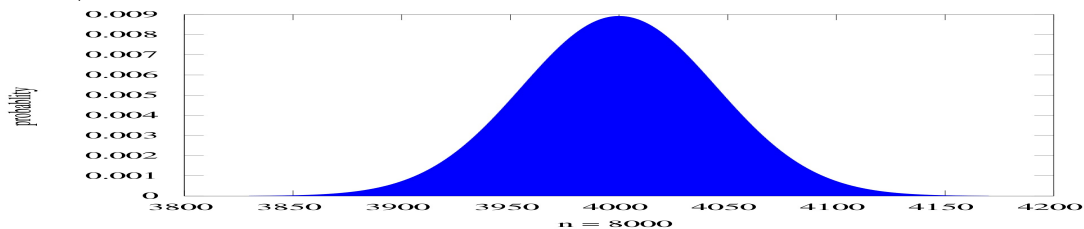
This is a very special case of the *law of large numbers*.

## 14.5.1 Side note...

### 14.5.1.1 Law of large numbers (weakest form)...

#### Informal statement of law of large numbers

For  $n$  large enough, the middle portion of the binomial distribution looks like (converges to) the normal/Gaussian distribution.



### 14.5.1.2 Massive randomness.. Is not that random.

#### Intuitive conclusion

Randomized algorithm are unpredictable in the tactical level, but very predictable in the strategic level.

### 14.5.1.3 Binomial distribution

$X_n$  = numbers of heads when flipping a coin  $n$  times.

## Claim

$$\Pr[X_n = i] = \frac{\binom{n}{i}}{2^n}.$$

$$\text{Where: } \binom{n}{k} = \frac{n!}{(n-k)!k!}.$$

Indeed,  $\binom{n}{i}$  is the number of ways to choose  $i$  elements out of  $n$  elements (i.e., pick which  $i$  coin flip come up heads).

Each specific such possibility (say 0100010...) had probability  $1/2^n$ .

We are interested in the bad event  $\Pr[X_n \leq n/4]$  (way too few heads). We are going to prove this probability is tiny.

## 14.5.2 Binomial distribution

### 14.5.2.1 Playing around with binomial coefficients

**Lemma 14.5.1.**  $n! \geq (n/e)^n$ .

*Proof:*

$$\frac{n^n}{n!} \leq \sum_{i=0}^{\infty} \frac{n^i}{i!} = e^n,$$

by the Taylor expansion of  $e^x = \sum_{i=0}^{\infty} \frac{x^i}{i!}$ . This implies that  $(n/e)^n \leq n!$ , as required. ■

## 14.5.3 Binomial distribution

### 14.5.3.1 Playing around with binomial coefficients

**Lemma 14.5.2.** For any  $k \leq n$ , we have  $\binom{n}{k} \leq \left(\frac{ne}{k}\right)^k$ .

*Proof:*

$$\begin{aligned} \binom{n}{k} &= \frac{n!}{(n-k)!k!} = \frac{n(n-1)(n-2)\dots(n-k+1)}{k!} \\ &\leq \frac{n^k}{k!} \leq \frac{n^k}{\left(\frac{k}{e}\right)^k} = \left(\frac{ne}{k}\right)^k. \end{aligned}$$

since  $k! \geq (k/e)^k$  (by previous lemma). ■

## 14.5.4 Binomial distribution

### 14.5.4.1 Playing around with binomial coefficients

$$\Pr\left[X_n \leq \frac{n}{4}\right] = \sum_{k=0}^{n/4} \frac{1}{2^n} \binom{n}{k} = \frac{1}{2^n} \sum_{k=0}^{n/4} \binom{n}{k} \leq \frac{1}{2^n} 2 \cdot \binom{n}{n/4}$$

For  $k \leq n/4$  the above sequence behave like a geometric variable.

$$\begin{aligned} \binom{n}{k+1} / \binom{n}{k} &= \frac{n!}{(k+1)!(n-k-1)!} / \frac{n!}{k!(n-k)!} \\ &= \frac{n-k}{k+1} \geq \frac{(3/4)n}{n/4+1} \geq 2. \end{aligned}$$

## 14.5.5 Binomial distribution

### 14.5.5.1 Playing around with binomial coefficients

$$\begin{aligned} \Pr\left[X_n \leq \frac{n}{4}\right] &\leq \frac{1}{2^n} 2 \cdot \binom{n}{n/4} \leq \frac{1}{2^n} 2 \cdot \left(\frac{ne}{n/4}\right)^{n/4} \leq 2 \cdot \left(\frac{4e}{2^4}\right)^{n/4} \\ &\leq 2 \cdot 0.68^{n/4}. \end{aligned}$$

We just proved the following theorem.

**Theorem 14.5.3.** *Let  $X_n$  be the random variable which is the number of heads when flipping an unbiased coin independently  $n$  times. Then*

$$\Pr\left[X_n \leq \frac{n}{4}\right] \leq 2 \cdot 0.68^{n/4} \text{ and } \Pr\left[X_n \geq \frac{3n}{4}\right] \leq 2 \cdot 0.68^{n/4}.$$

## 14.6 Randomized Quick Sort and Selection

### 14.7 Randomized Quick Sort

#### 14.7.0.2 Randomized QuickSort

Randomized **QuickSort**

- (A) Pick a pivot element *uniformly at random* from the array.
- (B) Split array into 3 subarrays: those smaller than pivot, those larger than pivot, and the pivot itself.
- (C) Recursively sort the subarrays, and concatenate them.

#### 14.7.0.3 Example

- (A) array: 16, 12, 14, 20, 5, 3, 18, 19, 1

#### 14.7.0.4 Analysis via Recurrence

- (A) Given array  $A$  of size  $n$ , let  $Q(A)$  be number of comparisons of randomized **QuickSort** on  $A$ .
- (B) Note that  $Q(A)$  is a random variable.
- (C) Let  $A_{\text{left}}^i$  and  $A_{\text{right}}^i$  be the left and right arrays obtained if:  
pivot is of rank  $i$  in  $A$ .

$$Q(A) = n + \sum_{i=1}^n \Pr[\text{pivot has rank } i] (Q(A_{\text{left}}^i) + Q(A_{\text{right}}^i)).$$

Since each element of  $A$  has probability exactly of  $1/n$  of being chosen:

$$Q(A) = n + \sum_{i=1}^n \frac{1}{n} (Q(A_{\text{left}}^i) + Q(A_{\text{right}}^i)).$$

#### 14.7.0.5 Analysis via Recurrence

Let  $T(n) = \max_{A:|A|=n} \mathbf{E}[Q(A)]$  be the worst-case expected running time of randomized **QuickSort** on arrays of size  $n$ .

We have, for any  $A$ :

$$Q(A) = n + \sum_{i=1}^n \Pr[\text{pivot has rank } i] (Q(A_{\text{left}}^i) + Q(A_{\text{right}}^i))$$

Therefore, by linearity of expectation:

$$\begin{aligned} \mathbf{E}[Q(A)] &= n + \sum_{i=1}^n \Pr[\text{pivot is of rank } i] (\mathbf{E}[Q(A_{\text{left}}^i)] + \mathbf{E}[Q(A_{\text{right}}^i)]). \\ \Rightarrow \mathbf{E}[Q(A)] &\leq n + \sum_{i=1}^n \frac{1}{n} (T(i-1) + T(n-i)). \end{aligned}$$

#### 14.7.0.6 Analysis via Recurrence

Let  $T(n) = \max_{A:|A|=n} \mathbf{E}[Q(A)]$  be the worst-case expected running time of randomized **QuickSort** on arrays of size  $n$ .

We derived:

$$\mathbf{E}[Q(A)] \leq n + \sum_{i=1}^n \frac{1}{n} (T(i-1) + T(n-i)).$$

Note that above holds for any  $A$  of size  $n$ . Therefore

$$\max_{A:|A|=n} \mathbf{E}[Q(A)] = T(n) \leq n + \sum_{i=1}^n \frac{1}{n} (T(i-1) + T(n-i)).$$

### 14.7.0.7 Solving the Recurrence

$$T(n) \leq n + \sum_{i=1}^n \frac{1}{n} (T(i-1) + T(n-i))$$

with base case  $T(1) = 0$ .

**Lemma 14.7.1.**  $T(n) = O(n \log n)$ .

*Proof:* (Guess and) Verify by induction. ■





# Bibliography

Hoare, C. A. R. (1962). Quicksort. *Comput. J.*, 5(1):10–15.