## CS 473 **♦** Fall 2024

## ∽ Homework 5 へ

## Due Wednesday, October 16, 2024 at 9pm Central Time

Unless a problem specifically states otherwise, you may assume a function RANDOM that takes a positive integer k as input and returns an integer chosen uniformly and independently at random from  $\{1, 2, ..., k\}$  in O(1) time. For example, to model a fair coin flip, you could call RANDOM(2).

1. Recall that a *priority search tree* is a binary tree in which every node has both a *search key* and a *priority*, arranged so that the tree is simultaneously a binary search tree for the keys and a min-heap for the priorities. A *heater* is a priority search tree in which the *priorities* are given by the user, and the *search keys* are distributed uniformly and independently at random in the real interval [0, 1]. Intuitively, a heater is a sort of anti-treap.

The following problems consider an *n*-node heater *T* whose priorities are the integers from 1 to *n*. We identify nodes in *T* by their *priorities*; thus, "node 5" means the node in *T* with *priority* 5. For example, the min-heap property implies that node 1 is the root of *T*. Finally, let *i* and *j* be integers with  $1 \le i < j \le n$ .

- (a) What is the *exact* expected depth of node *j* in an *n*-node heater? Answering the following subproblems will help you:
  - i. Prove that in a random permutation of the (i + 1)-element set  $\{1, 2, ..., i, j\}$ , elements *i* and *j* are adjacent with probability 2/(i + 1).
  - ii. Prove that node *i* is an ancestor of node *j* with probability 2/(i + 1). [Hint: Use the previous question!]
  - iii. What is the probability that node *i* is a *descendant* of node *j*? [*Hint: Do* **not** *use the previous question*!]
- (b) Describe an algorithm to insert a new item into a heater. Analyze the expected running time as a function of the number of nodes.
- (c) Describe an algorithm to delete the minimum-priority item (the root) from an *n*-node heater. What is the expected running time of your algorithm?
- Suppose we generate a bit-string *w* by flipping a fair coin *n* times. Thus, each bit in *w* is equal to 0 or 1 with equal probability, and the bits in *w* are fully independent. A *run of length* ℓ in *w* is a substring of length ℓ in which all bits are equal. For example, the string 01000011101 contains *three* runs of length 3, starting at the third, fourth, and seventh bits
  - (a) Suppose *n* is a power of 2. Show that the expected number of runs of length  $\lg n + 1$  is 1 o(1). (Here "lg" is standard shorthand for log-base-2.)
  - (b) Show that, for sufficiently large *n*. the probability that *every* run in *w* has length less than [lg *n*−2lg lg *n*] is less than 1/*n*. [*Hint: Break w into disjoint substrings of length* [lg *n*−2lg lg *n*] and use the following fact: The event that all bits in one substring are equal is independent of the event that all bits in any other substring are equal.]

3. Suppose we are given a coin that may or may not be biased, and we would like to compute an accurate *estimate* of the probability of heads. Specifically, if the actual unknown probability of heads is p, we would like to compute an estimate  $\tilde{p}$  such that

$$\Pr[|\tilde{p} - p| > \varepsilon] < \delta$$

where  $\varepsilon$  is a given *accuracy* or *error* parameter, and  $\delta$  is a given *confidence* parameter.

The following algorithm is a natural first attempt; here FLIP() returns the result of an independent flip of the unknown coin.

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\frac{\text{MEANESTIMATE}(\varepsilon):}{count \leftarrow 0}
for i \leftarrow 1 to N
if FLIP() = HEADS
count \leftarrow count + 1
return count/N
```

- (a) Let  $\tilde{p}$  denote the estimate returned by MEANESTIMATE( $\varepsilon$ ). Prove that  $E[\tilde{p}] = p$ .
- (b) Prove that if we set  $N = \lceil \alpha/\varepsilon^2 \rceil$  for some appropriate constant  $\alpha$ , then we have  $\Pr[|\tilde{p} p| > \varepsilon] < 1/4$ . [*Hint: Use Chebyshev's inequality.*]
- (c) We can increase the previous estimator's confidence (for the same accuracy) by running it multiple times, independently, and returning the *median* of the resulting estimates.

$$\frac{\text{MEDIANOFMEANSESTIMATE}(\delta, \varepsilon):}{\text{for } j \leftarrow 1 \text{ to } K}$$
$$estimate[j] \leftarrow \text{MEANESTIMATE}(\varepsilon)$$
$$return \text{MEDIAN}(estimate[1..K])$$

Let  $p^*$  denote the estimate returned by MEDIANOFMEANSESTIMATE( $\delta, \varepsilon$ ). Prove that if we set  $N = \lceil \alpha/\varepsilon^2 \rceil$  (inside MEANESTIMATE) and  $K = \lceil \beta \ln(1/\delta) \rceil$ , for some appropriate constants  $\alpha$  and  $\beta$ , then  $\Pr[|p^* - p| > \varepsilon] < \delta$ . [Hint: Use Chernoff bounds.]