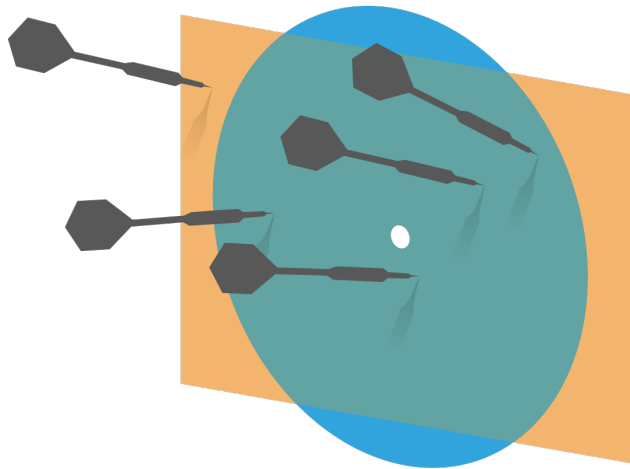


Probability and Statistics for Computer Science



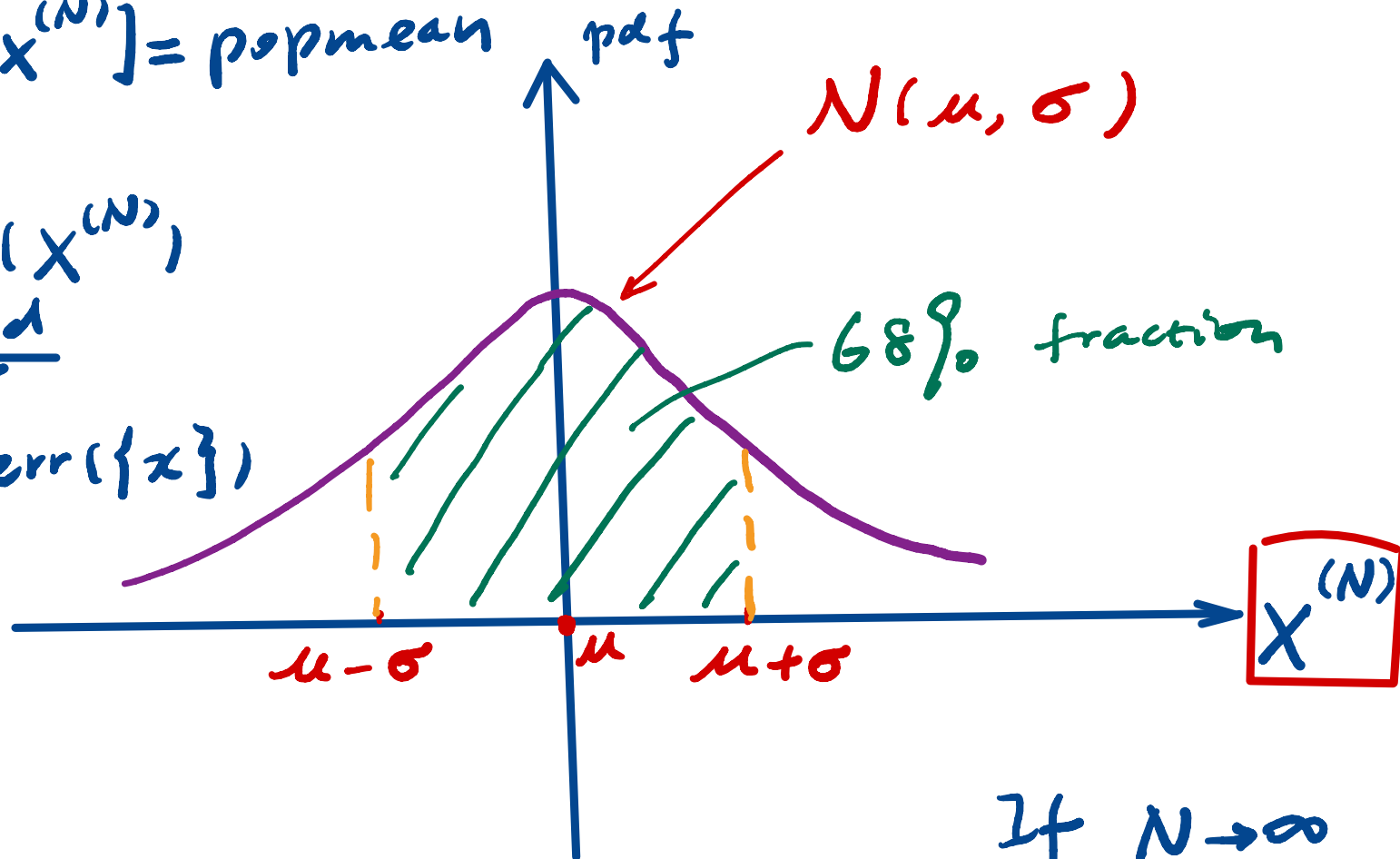
"Statistical thinking will one day be as necessary for efficient citizenship as the ability to read and write." H. G. Wells

Credit: wikipedia

Meaning of #% Confidence Interval

$$\mu = E[X^{(N)}] = \text{popmean}$$

$$\begin{aligned}\sigma &= \text{std}(X^{(N)}) \\ &= \frac{\text{popstd}}{\sqrt{N}} \\ &\doteq \text{stderr}(\{x\})\end{aligned}$$



Meaning of #% Confidence Interval

$$\mu = E[X^{(N)}] = \text{popmean}$$

$$\text{if } \mu - \sigma \leq x_0 \leq \mu + \sigma$$

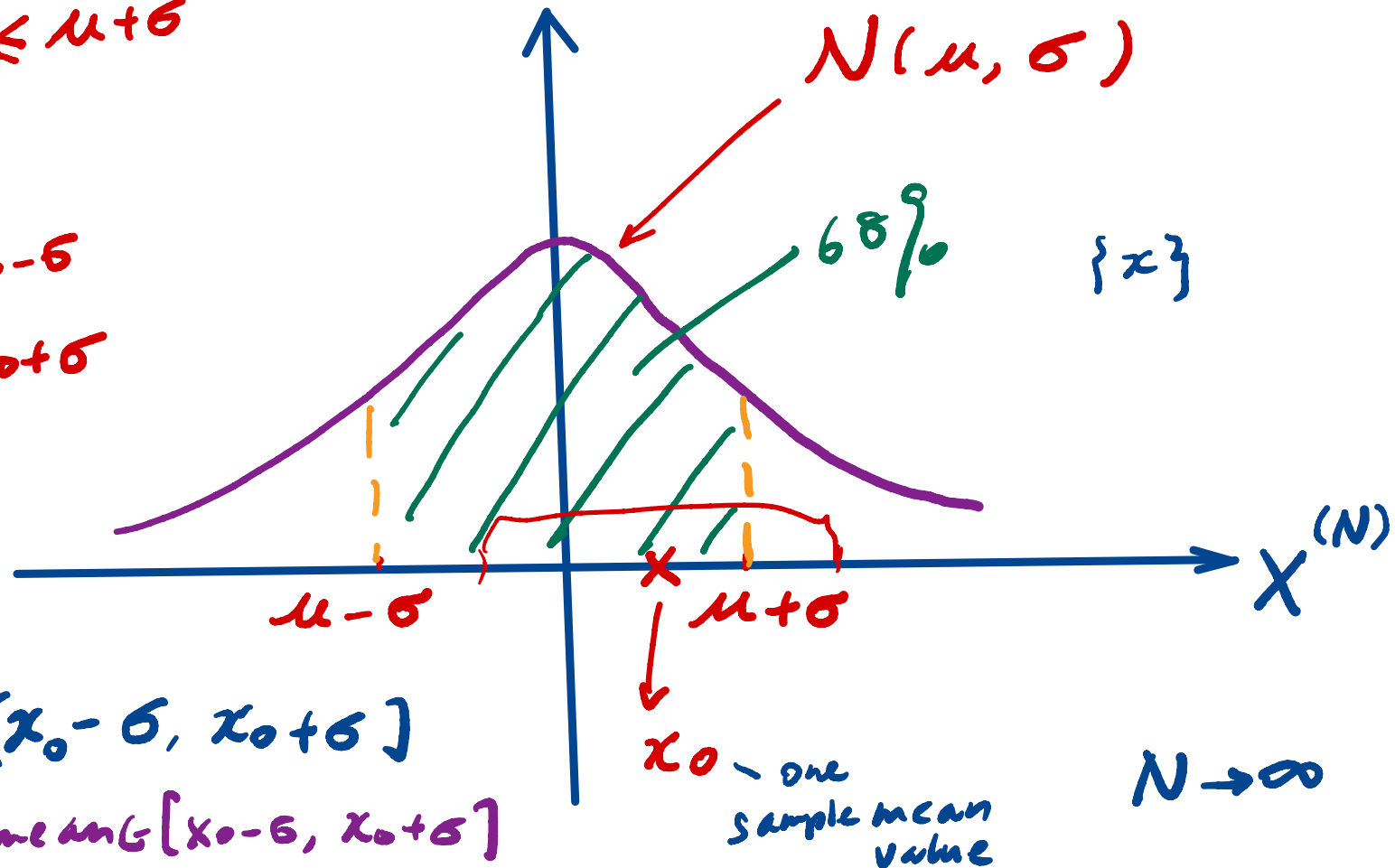


$$\mu \geq x_0 - \sigma$$

$$\mu \leq x_0 + \sigma$$

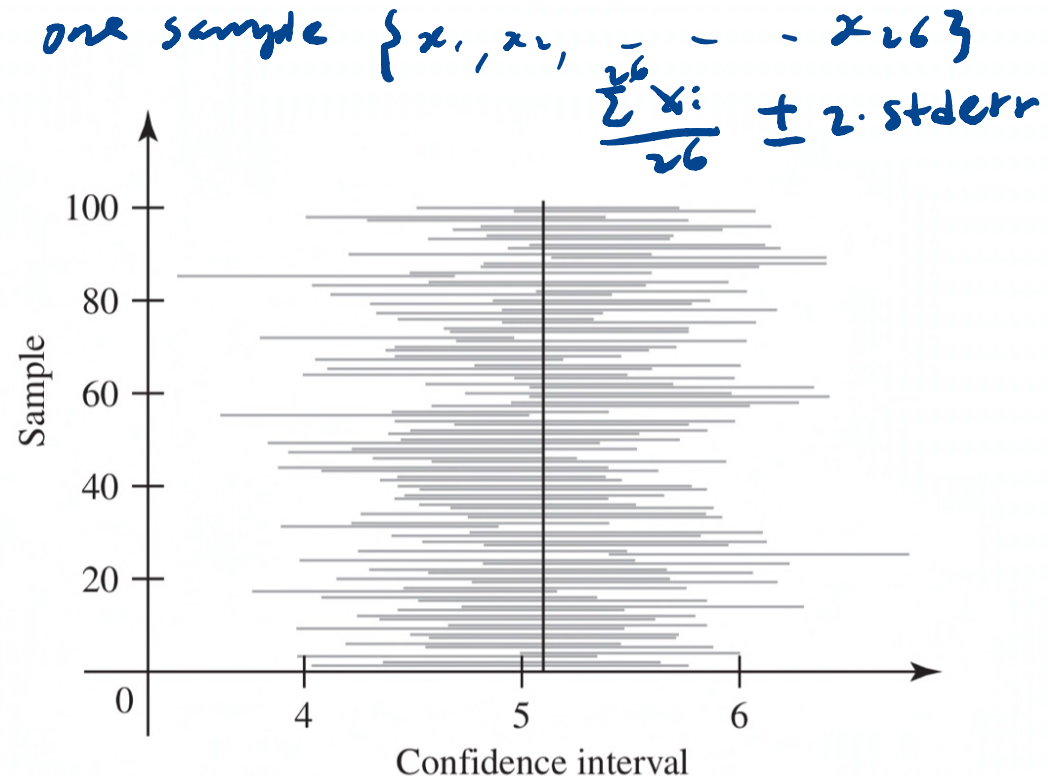
$$\mu \in [x_0 - \sigma, x_0 + \sigma]$$

$$\downarrow \text{popmean} \in [x_0 - \sigma, x_0 + \sigma]$$



Meaning of #% Confidence Interval

Figure 8.5 A sample of one hundred observed 95% confidence intervals based on samples of size 26 from the normal distribution with mean $\mu = 5.1$ and standard deviation $\sigma = 1.6$. In this figure, 94% of the intervals contain the value of μ .



Deyoung Pg 487

A tale of two statisticians

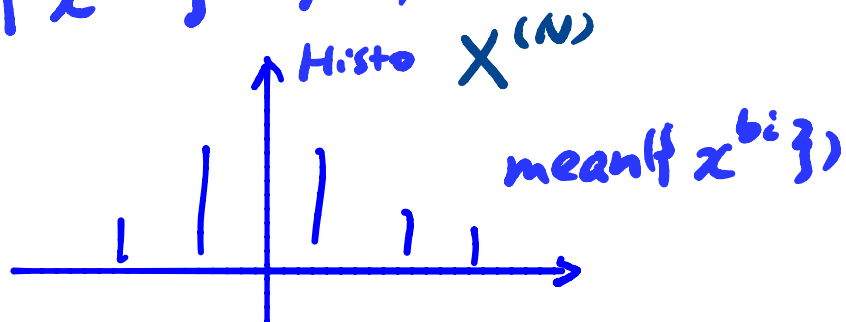
$$\{X\} = \{1, 2, 3, \dots, 12\} \quad N_p = 12$$

$$\{x^b\} = \{1, 4, 5, 7, 11\}$$

$$\{x^{b1}\} = \{1, 1, 4, 5, 7\} \quad x^{b1} = \frac{18}{5}$$

$$\{x^{b2}\} = \{4, 5, 7, 7, 1\}$$

$$\{x^{b3}\} = \{5, 5, 5, 5, 5\}$$



$$\{x\} = \{1, 4, 5, 7, 11\} \quad N=5$$

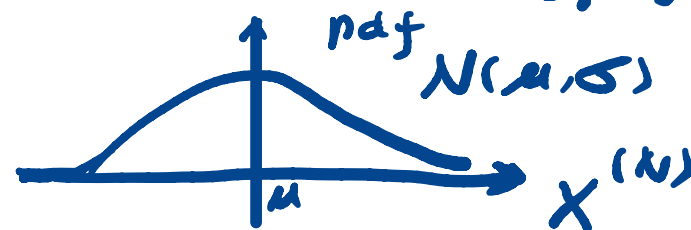
i.i.d. $x^{(i)}$

$$\text{if } N \rightarrow \infty$$

$$X^{(N)} \sim N(\mu, \sigma)$$

$$\mu = E[X^{(N)}] = \text{mean}(\{x\})$$

$$\sigma = \text{std}[X^{(N)}] = \text{stderr}(\{x\})$$



Objectives

- ✱ Hypothesis test
- ✱ Chi-square test
- ✱ Maximum Likelihood Estimation

A hypothesis

- Ms. Smith's vote percentage is 55% *Simple $\theta = \theta_0$*
This is what we want to test, often called null hypothesis H_0
 H_1 : perct $\neq 55\%$

| | | DATES | POLLSTER | SAMPLE | RESULT | | NET RESULT |
|-------------|-------|--------------|--------------------|----------|----------|----------------|---------------|
| U.S. Senate | Miss. | NOV 25, 2018 | C+ Change Research | 1,211 LV | Espy 46% | 51% Hyde-Smith | Hyde-Smith +5 |

51%

- Should we reject this hypothesis given the poll data?

Rejection region of null hypothesis H_0

✱ Assuming the hypothesis H_0 is true

pop mean = μ_0

✱ Define a test statistic

mean($\{x\}$)

$$x = \frac{(\text{sample mean}) - (\text{hypothesized value})}{\text{standard error } \text{i+der($\{x\}$)}}$$

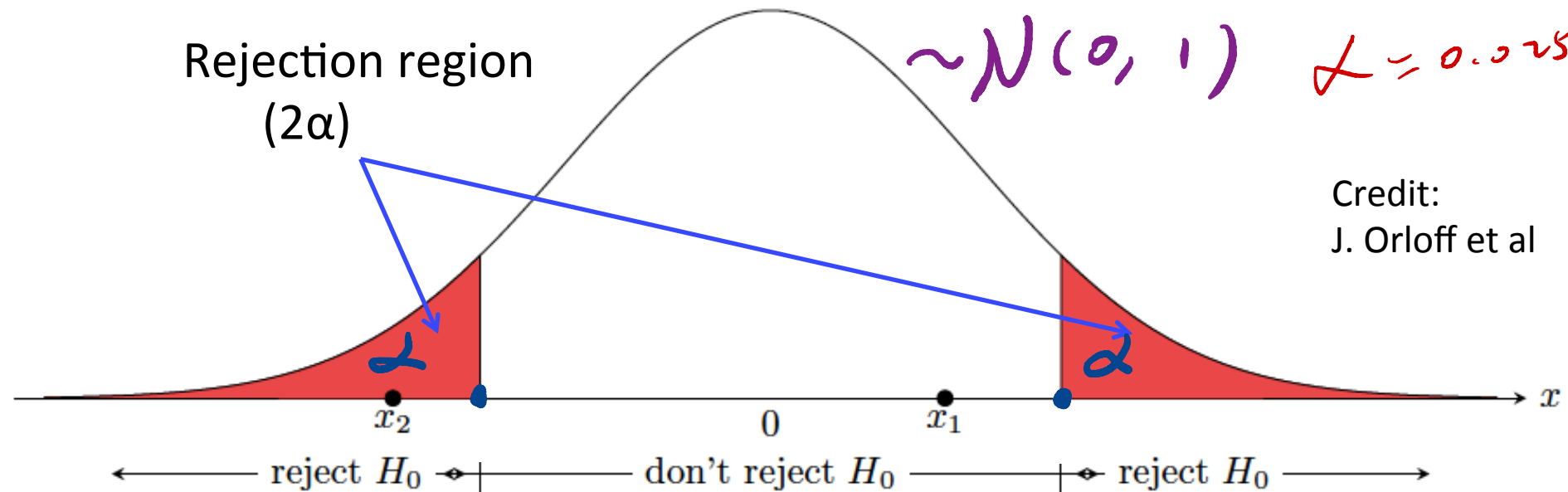
✱ Since $N > 30$, assume x comes from a standard normal

$\sim N(0, 1)$

$\alpha = 0.025$

Rejection region
(2α)

Credit:
J. Orloff et al



Fraction of “less extreme” statistic

✱ Assuming the hypothesis H_0 is true $\{X\} = \{10, 20, \dots, 50\}$

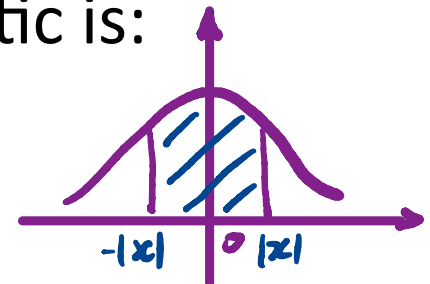
✱ Define a statistic for the test $\{x\} = \{10, 20, 30\}$
 $\rightarrow \text{mean}(\{x\}) = 20$

$$x = \frac{(\text{sample mean}) - (\text{hypothesized value})}{\text{standard error}} \quad \mu_0 = 50$$

✱ Since $N > 30$, we assume x comes from a standard normal

✱ So, the fraction of “less extreme” statistic is:

$$f = \frac{1}{\sqrt{2\pi}} \int_{-|x|}^{|x|} \exp\left(-\frac{u^2}{2}\right) du$$



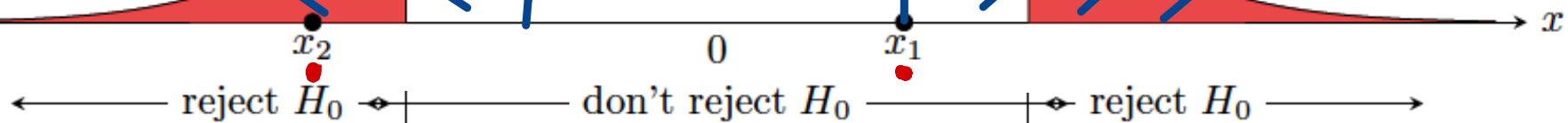
P-value: Rejection region- “The extreme fraction”

- ✱ It is conventional to report the p-value

$$p = 1 - f = 1 - \frac{1}{\sqrt{2\pi}} \int_{-|x|}^{|x|} \exp\left(-\frac{u^2}{2}\right) du$$

Rejection region
(2α)

By convention:
 $2\alpha = 0.05$
That is:
If $p < 0.05$, reject H_0



p-value: election polling

✱ H_0 : Ms. Smith's vote percentage is 55% *population*

✱ The sample mean is 51% and stderr is 1.44%

✱ The test statistic $x = \frac{51 - 55}{1.44} = -2.7778$

✱ And the p-value for the test is:

$$p = 1 - \frac{1}{\sqrt{2\pi}} \int_{-2.7778}^{2.7778} \exp\left(-\frac{u^2}{2}\right) du = 0.00547 < \underline{0.05}$$

✱ So we reject the hypothesis

Convention

Hypothesis test if $N < 30$

- ✱ Q: what distribution should we use to test the hypothesis of sample mean if $N < 30$?
- A. Normal distribution
 - B. t-distribution with degree = 30
 - C. t-distribution with degree = N
 - ☒ D. t-distribution with degree = $N-1$

The use and misuse of p-value

- ✱ p-value use in scientific practice
 - ✱ Usually used to reject the null hypothesis that the data is random noise
 - ✱ Common practice is $p < 0.05$ is considered significant evidence for something interesting
- ✱ Caution about p-value hacking
 - ✱ Rejecting the null hypothesis doesn't mean the alternative is true
 - ✱ $P < 0.05$ is arbitrary and often is not enough for controlling false positive phenomenon

Chi-square distribution

- ✱ If Z_i 's are independent variables of standard normal distribution, $X = Z_1^2 + Z_2^2 + \dots + Z_m^2 = \sum_{i=1}^m Z_i^2$ has a Chi-square distribution with degree of freedom m , $X \sim \chi^2(m)$
- ✱ We can test the goodness of fit for a model using a statistic C against this distribution, where

$$C = \sum_{i=1}^m \frac{(f_o(\varepsilon_i) - f_t(\varepsilon_i))^2}{f_t(\varepsilon_i)}$$

$\varepsilon_i \rightarrow \text{event } i$

Independence analysis using Chi-square

- ✱ Given the two way table, test whether the ^{indep.} column and row are independent
- $$P(A \cap B) = P(A)P(B)$$

| | Boy | Girl | Total |
|---------|-----|------|-------|
| Grades | 117 | 130 | 247 |
| Popular | 50 | 91 | 141 |
| Sports | 60 | 30 | 90 |
| Total | 227 | 251 | 478 |

$$P(A|B) = P(A)$$

Independence analysis using Chi-square

- ✱ The theoretical expected values if independent

$$247 \times \frac{227}{478}$$

| | Boy | Girl | Total |
|---------|-----------|-----------|-------|
| Grades | 117.29916 | 129.70084 | 247 |
| Popular | 66.96025 | 74.03975 | 141 |
| Sports | 42.74059 | 47.25941 | 90 |
| Total | 227 | 251 | 478 |

The degree of the chi-square distribution for the two way table

- ✱ The degree of freedom for the chi-square distribution for a r by c table is

$$(r-1) \times (c-1) \text{ where } r > 1 \text{ and } c > 1$$

$$r = 3$$

$$c = 2$$

- ✱ Because the degree $df = n - 1 - p$ See textbook Pg 171-172

$$= rc - 1 - (r-1) - (c-1)$$

n is the number of cells of data;

p is the number of unknown parameters

$$= (r-1) \times (c-1)$$

$$= 2$$

Chi-square test for the popular kid data

✱ The Chi-statistic : 21.455

```
chisq.test(data_BG)
```

Pearson's Chi-squared test

data: data_BG

X-squared = 21.455, df = 2, p-value = 2.193e-05

✱ P-value: 2.193e-05

✱ It's very unlikely the two categories are independent

Q. What is the degree of freedom for this?

✱ The following 2-way table for chi-square test has a degree of freedom equal to:

Table 10.26 Data for Exercise 3

| | Number of lectures attended | | | | |
|------------|-----------------------------|----|----|---|----|
| | 0 | 1 | 2 | 3 | 4 |
| Freshmen | 10 | 16 | 27 | 6 | 11 |
| Sophomores | 14 | 19 | 20 | 4 | 13 |
| Juniors | 15 | 15 | 17 | 4 | 9 |
| Seniors | 19 | 8 | 6 | 5 | 12 |

$$r = 4$$

$$c = 5$$

$$(4-1)(5-1)$$

$$= 3 \times 4$$

A. 20

B. 9

☒ C. 12

D. 4

Chi-square test is very versatile

- ✱ Chi-square test is so versatile that it can be utilized in many ways either for discrete data or continuous data via intervals
- ✱ Please check out the worked-out examples in the textbook and read more about its applications.

Maximum likelihood estimation

$$P(X=k) = \binom{N}{k} p^k (1-p)^{N-k} \quad k \geq 0$$

write $P(X=k)$ p is unknown
write p as θ

$$L(\theta) = \binom{N}{k} \theta^k (1-\theta)^{N-k}$$

Maximize $L(\theta)$, we get $\hat{\theta}$

$$\hat{\theta} = \underset{\theta}{\operatorname{Argmax}} L(\theta)$$

$$D: N, k$$

Motivation: Poisson example

- ✱ Suppose we have data on the number of babies born each hour in a large hospital

| hour | 1 | 2 | ... | N |
|-------------|-------|-------|-----|-------|
| # of babies | k_1 | k_2 | ... | k_N |

λ

- ✱ We can assume the data comes from a Poisson distribution
- ✱ What is your best estimate of the intensity λ ?

Maximum likelihood estimation (MLE)

- ✱ We write the probability of seeing the data D given parameter θ

$$L(\theta) = P(D|\theta)$$

- ✱ The **likelihood function** $L(\theta)$ is **not** a probability distribution

- ✱ The **maximum likelihood estimate (MLE)** of θ is

$$\hat{\theta} = \arg \max_{\theta} L(\theta)$$

Why is $L(\theta)$ not a probability distribution?

- A. It doesn't give the probability of all the possible θ values.
- B. Don't know whether the sum or integral of $L(\theta)$ for all possible θ values is one or not.
- ☒ C. Both.

Likelihood function: binomial example

- ✱ Suppose we have a coin with unknown probability of coming up heads ⑨
- ✱ We toss it ***N*** times and observe ***k*** heads
- ✱ We know that this data comes from a binomial distribution
- ✱ What is the likelihood function $L(\theta) = P(D|\theta)$?

$$L(\theta) = \binom{N}{k} \theta^k (1 - \theta)^{N-k}$$

⑨ = prob. of head

Likelihood function: binomial example

✱ Suppose we have a coin with unknown probability of θ coming up heads

✱ We toss it **10** times and observe **7** heads

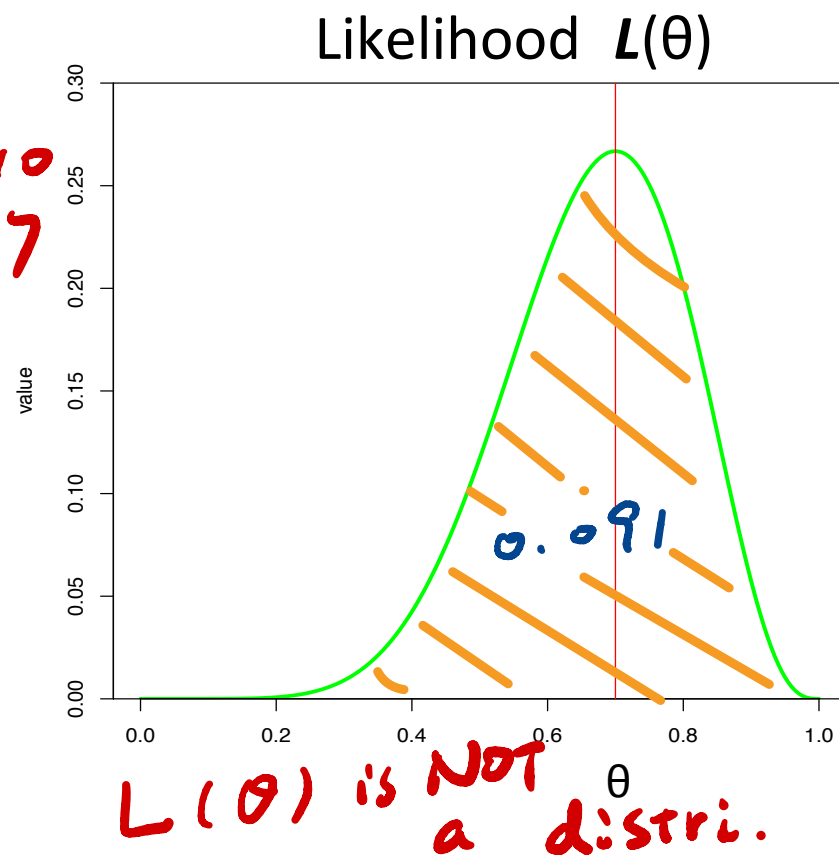
*D: N=10
k=7*

✱ The likelihood function is:

$$P(D|\theta) = \binom{10}{7} \theta^7 (1 - \theta)^3$$

✱ The MLE is

$$\hat{\theta} = 0.7$$



MLE derivation: binomial example

$$L(\theta) = \binom{N}{k} \theta^k (1 - \theta)^{N-k}$$

In order to find: $\hat{\theta} = \underset{\theta}{arg \max} L(\theta)$

We set: $\frac{dL(\theta)}{d\theta} = 0$

MLE derivation: binomial example

$$L(\theta) = \binom{N}{k} \theta^k (1 - \theta)^{N-k}$$

$(a(x)b(x))'$
 $a'b + ab'$

$$\frac{d}{d\theta} L(\theta) = \binom{N}{k} (k\theta^{k-1}(1-\theta)^{N-k} - \theta^k(N-k)(1-\theta)^{N-k-1}) = 0$$

$$k\theta^{k-1}(1-\theta)^{N-k} = \theta^k(N-k)(1-\theta)^{N-k-1}$$

$$k - k\theta = N\theta - k\theta$$

$$\hat{\theta} = \frac{k}{N}$$

The MLE of p

p : prob of seeing H

Likelihood function: geometric example

- ✱ Suppose we have a die with unknown probability of coming up six
- ✱ We roll it and it comes up six for the first time on the k th roll
- ✱ We know that this data comes from a geometric distribution
- ✱ What is the likelihood function $L(\theta) = P(D|\theta)$?
Assume θ is p .

MLE derivation: geometric example

$$L(\theta) = (1 - \theta)^{k-1} \theta$$

$P(D|\theta)$

what is the D ?

$D: k$

$$\hat{\theta} = \arg \max_{\theta} L(\theta)$$

Likelihood function: geometric example

- ✱ Suppose we have a die with unknown probability of coming up six
- ✱ We roll it and it comes up six for the first time on the k th roll
- ✱ We know that this data comes from a geometric distribution
- ✱ What is the likelihood function $L(\theta) = P(D|\theta)$?
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MLE derivation: geometric example

$$L(\theta) = (1 - \theta)^{k-1} \theta$$

MLE derivation: geometric example

$$L(\theta) = (1 - \theta)^{k-1} \theta$$

$$\frac{d}{d\theta} L(\theta) = (1 - \theta)^{k-1} - (k - 1)(1 - \theta)^{k-2} \theta = 0$$

MLE derivation: geometric example

$$L(\theta) = (1 - \theta)^{k-1} \theta$$

$$\frac{d}{d\theta} L(\theta) = (1 - \theta)^{k-1} - (k - 1)(1 - \theta)^{k-2} \theta = 0$$

$$(1 - \theta)^{k-1} = (k - 1)(1 - \theta)^{k-2} \theta$$

MLE derivation: geometric example

$$L(\theta) = (1 - \theta)^{k-1} \theta$$

$$\frac{d}{d\theta} L(\theta) = (1 - \theta)^{k-1} - (k - 1)(1 - \theta)^{k-2} \theta = 0$$

$$(1 - \theta)^{k-1} = (k - 1)(1 - \theta)^{k-2} \theta$$

$$1 - \theta = k\theta - \theta$$

MLE derivation: geometric example

$$L(\theta) = (1 - \theta)^{k-1} \theta$$

$$\frac{d}{d\theta} L(\theta) = (1 - \theta)^{k-1} - (k - 1)(1 - \theta)^{k-2} \theta = 0$$

$$(1 - \theta)^{k-1} = (k - 1)(1 - \theta)^{k-2} \theta$$

$$1 - \theta = k\theta - \theta$$

$$\hat{\theta} = \frac{1}{k}$$

The MLE of p

Assignments

- ✱ Finish Chapter 7 of the textbook
- ✱ Next time: Maximum likelihood estimate, Bayesian inference

See you next time

*See
You!*

