## Lecture 15: A tutorial on hidden Markov models and selected applications in speech recognition, part 2

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ECE 537, Fall 2022
(1) Review: Hidden Markov Models
(2) Maximum-Likelihood Training of an HMM
(3) Baum-Welch Re-Estimation

4 Gaussian Observation Probabilities
(5) Summary

## Outline

(1) Review: Hidden Markov Models
(2) Maximum-Likelihood Training of an HMM

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## Hidden Markov Model


(1) Start in state $q_{t}=i$ with pmf $\pi_{i}$.
(2) Generate an observation, $\vec{o}$, with pdf $b_{i}(\vec{o})$.
(3) Transition to a new state, $q_{t+1}=j$, according to pmf $a_{i j}$.
(9) Repeat.

## The Three Problems for an HMM

(1) Recognition: Given two different HMMs, $\lambda_{1}$ and $\lambda_{2}$, and an observation sequence $O$. Which HMM was more likely to have produced $O$ ? In other words, $p\left(O \mid \lambda_{1}\right)>p\left(O \mid \lambda_{2}\right)$ ?
(2) Segmentation: What is $p\left(q_{t}=i \mid O, \lambda\right)$ ?
(3) Training: Given an initial HMM $\lambda$, and an observation sequence $O$, can we find $\bar{\lambda}$ such that $p(O \mid \bar{\lambda})>p(O \mid \lambda)$ ?

## The Forward Algorithm

Definition: $\alpha_{t}(i) \equiv p\left(\vec{o}_{1}, \ldots, \vec{o}_{t}, q_{t}=i \mid \lambda\right)$. Computation:
(1) Initialize:

$$
\alpha_{1}(i)=\pi_{i} b_{i}\left(\vec{o}_{1}\right), \quad 1 \leq i \leq N
$$

(2) Iterate:

$$
\alpha_{t}(j)=\sum_{i=1}^{N} \alpha_{t-1}(i) a_{i j} b_{j}\left(\vec{o}_{t}\right), \quad 1 \leq j \leq N, 2 \leq t \leq T
$$

(3) Terminate:

$$
p(O \mid \lambda)=\sum_{i=1}^{N} \alpha_{T}(i)
$$

## The Backward Algorithm

Definition: $\beta_{t}(i) \equiv p\left(\vec{o}_{t+1}, \ldots, \vec{o}_{T} \mid q_{t}=i, \lambda\right)$. Computation:
(1) Initialize:

$$
\beta_{T}(i)=1, \quad 1 \leq i \leq N
$$

(2) Iterate:

$$
\beta_{t}(i)=\sum_{j=1}^{N} a_{i j} b_{j}\left(\vec{o}_{t+1}\right) \beta_{t+1}(j), \quad 1 \leq i \leq N, 1 \leq t \leq T-1
$$

(3) Terminate:

$$
p(O \mid \lambda)=\sum_{i=1}^{N} \pi_{i} b_{i}\left(\vec{o}_{1}\right) \beta_{1}(i)
$$

## Segmentation

(1) The State Posterior:

$$
\gamma_{t}(i)=p\left(q_{t}=i \mid O, \lambda\right)=\frac{\alpha_{t}(i) \beta_{t}(i)}{\sum_{k=1}^{N} \alpha_{t}(k) \beta_{t}(k)}
$$

(2) The Segment Posterior:

$$
\begin{aligned}
\xi_{t}(i, j) & =p\left(q_{t}=i, q_{t+1}=j \mid O, \lambda\right) \\
& =\frac{\alpha_{t}(i) a_{i j} b_{j}\left(\vec{o}_{t+1}\right) \beta_{t+1}(j)}{\sum_{k=1}^{N} \sum_{\ell=1}^{N} \alpha_{t}(k) a_{k \ell} b_{\ell}\left(\vec{o}_{t+1}\right) \beta_{t+1}(\ell)}
\end{aligned}
$$

## The Three Problems for an HMM

(1) Recognition: Given two different HMMs, $\lambda_{1}$ and $\lambda_{2}$, and an observation sequence $O$. Which HMM was more likely to have produced $O$ ? In other words, $p\left(O \mid \lambda_{1}\right)>p\left(O \mid \lambda_{2}\right)$ ?
(2) Segmentation: What is $p\left(q_{t}=i \mid O, \lambda\right)$ ?
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## Maximum Likelihood Training

Suppose we're given several observation sequences of the form $O=\left[\vec{o}_{1}, \ldots, \vec{o}_{T}\right]$. Suppose, also, that we have some initial guess about the values of the model parameters (our initial guess doesn't have to be very good). Maximum likelihood training means we want to compute a new set of parameters, $\bar{\lambda}=\left\{\bar{\pi}_{i}, \bar{a}_{i j}, \bar{b}_{j}(\vec{o})\right\}$ that maximize $p(O \mid \bar{\lambda})$.
(1) Initial State Probabilities: Find values of $\bar{\pi}_{i}, 1 \leq i \leq N$, that maximize $p(O \mid \bar{\lambda})$.
(2) Transition Probabilities: Find values of $\bar{a}_{i j}, 1 \leq i, j \leq N$, that maximize $p(O \mid \bar{\lambda})$.
(3) Observation Probabilities: Learn $\bar{b}_{j}(\vec{o})$. What does that mean, actually?

## Learning the Observation Probabilities

There are four typical ways of modeling the observations:
(1) Discrete: Vector quantize $\vec{o}$, using some VQ method. Suppose $\vec{o}$ is the $k^{\text {th }}$ codevector; then we just need to learn $b_{j}(k)$ such that

$$
b_{j}(k) \geq 0, \quad \sum_{k=0}^{K-1} b_{j}(k)=1
$$

(2) Gaussian: Model $b_{j}(k)$ as a Gaussian or mixture Gaussian, and learn its parameters.
(3) Neural Net: Model $b_{j}(k)$ as a neural net, and learn its parameters.
For now, assume discrete observations.

## Maximum Likelihood Training

Given discrete observations, we need to learn the following parameters:
(1) Initial State Probabilities: $\bar{\pi}_{i}$ such that

$$
\bar{\pi}_{i} \geq 0, \quad \sum_{i=1}^{N} \bar{\pi}_{i}=1
$$

(2) Transition Probabilities: $\bar{a}_{i j}$ such that

$$
\bar{a}_{i j} \geq 0, \quad \sum_{j=1}^{N} \bar{a}_{i j}=1
$$

(3) Observation Probabilities: $\bar{b}_{j}(k)$ such that

$$
\bar{b}_{j}(k) \geq 0, \quad \sum_{k=1}^{K} \bar{b}_{j}(k)=1
$$

## Maximum Likelihood Training with Known State Sequence

Impossible assumption: Suppose that we actually know the state sequences, $Q=\left[q_{1}, \ldots, q_{T}\right]$, matching with each observation sequence $O=\left[\vec{o}_{1}, \ldots, \vec{o}_{T}\right]$. Then what would be the maximum-likelihood parameters?

## Maximum Likelihood Training with Known State Sequence

Our goal is to find $\lambda=\left\{\pi_{i}, a_{i j}, b_{j}(k)\right\}$ in order to maximize

$$
\begin{aligned}
\mathcal{L}(\lambda) & =\ln p(Q, O \mid \lambda) \\
& =\ln \pi_{q_{1}}+\ln b_{q_{1}}\left(o_{1}\right)+\ln a_{q_{1}, q_{2}}+b_{q_{2}}\left(o_{2}\right)+\ldots \\
& =\ln \pi_{q_{1}}+\sum_{i=1}^{N}\left(\sum_{j=1}^{N} n_{i j} \ln a_{i j}+\sum_{k=1}^{K} m_{i k} \ln b_{i}(k)\right)
\end{aligned}
$$

where

- $n_{i j}$ is the number of times we saw $\left(q_{t}=i, q_{t+1}=j\right)$,
- $m_{i k}$ is the number of times we saw $\left(q_{t}=i, k_{t}=k\right)$


## Maximum Likelihood Training with Known State Sequence

$$
\mathcal{L}(\lambda)=\ln \pi_{q_{1}}+\sum_{i=1}^{N}\left(\sum_{j=1}^{N} n_{i j} \ln a_{i j}+\sum_{k=1}^{K} m_{i k} \ln b_{i}(k)\right)
$$

When we differentiate that, we find the following derivatives:

$$
\begin{aligned}
\frac{\partial \mathcal{L}}{\partial \pi_{i}} & = \begin{cases}\frac{1}{\pi_{i}} & i=q_{1} \\
0 & \text { otherwise }\end{cases} \\
\frac{\partial \mathcal{L}}{\partial a_{i j}} & =\frac{n_{i j}}{a_{i j}} \\
\frac{\partial \mathcal{L}}{\partial b_{j}(k)} & =\frac{m_{j k}}{b_{j}(k)}
\end{aligned}
$$

These derivatives are never equal to zero! What went wrong?

## Maximum Likelihood Training with Known State Sequence

Here's the problem: we forgot to include the constraints
$\sum_{i} \pi_{i}=1, \sum_{j} a_{i j}=1$, and $\sum_{k} b_{j}(k)=1$ !
We can include the constraints using a Lagrangian optimization.

## Maximum Likelihood Training with Known State Sequence

The Lagrangian, $\mathcal{J}(\lambda)$, is the thing we want to optimize $(\mathcal{L}(\lambda))$, plus the things that should be zero, each of which is multiplied by an arbitrary constant called a Lagrange multiplier:

$$
\begin{gathered}
\mathcal{J}(\lambda)=\ln \pi_{q_{1}}+\sum_{i=1}^{N}\left(\sum_{j=1}^{N} n_{i j} \ln a_{i j}+\sum_{k=1}^{K} m_{i k} \ln b_{i}(k)\right) \\
+\kappa\left(1-\sum_{i=1}^{N} \pi_{i}\right)+\sum_{i=1}^{N} \mu_{i}\left(1-\sum_{j=1}^{N} a_{i j}\right)+\sum_{i=1}^{N} \nu_{i}\left(1-\sum_{k=1}^{M} b_{j}(k)\right)
\end{gathered}
$$

(1) First solve for the parameters as functions of the Lagrange multipliers.
(2) Second, set the Lagrange multipliers equal to whatever value will zero out the constraints.

## Maximum Likelihood Training with Known State Sequence

Step 1: Solve for the parameters as functions of the Lagrange multipliers. If we set

$$
\frac{\partial \mathcal{J}(\lambda)}{\partial \pi_{i}}=\frac{\partial \mathcal{J}(\lambda)}{\partial a_{i j}}=\frac{\partial \mathcal{J}(\lambda)}{\partial b_{j k}}=0
$$

we get:

$$
\bar{\pi}_{i}=\left\{\begin{array}{ll}
\frac{1}{\kappa} & i=q_{1} \\
0 & \text { otherwise }
\end{array}, \quad \bar{a}_{i j}=\frac{n_{i j}}{\mu_{i}}, \quad \bar{b}_{j}(k)=\frac{m_{j k}}{\nu_{j}}\right.
$$

## Maximum Likelihood Training with Known State Sequence

Step 2: Set the Lagrange multipliers to whatever value zeros out the constraints:

$$
\begin{gathered}
\bar{\pi}_{i}= \begin{cases}1 & i=q_{1} \\
0 & \text { otherwise }\end{cases} \\
\bar{a}_{i j}=\frac{n_{i j}}{\sum_{j=1}^{N} n_{i j}} \\
\bar{b}_{j}(k)=\frac{m_{j k}}{\sum_{k=1}^{M} m_{j k}}
\end{gathered}
$$

## Maximum Likelihood Training with Known State Sequence

Using the Lagrange multiplier method, we can show that the maximum likelihood parameters for the HMM are:
(1) Initial State Probabilities:

$$
\bar{\pi}_{i}=\frac{\# \text { state sequences that start with } q_{1}=i}{\# \text { state sequences in training data }}
$$

(2) Transition Probabilities:

$$
\bar{a}_{i j}=\frac{\# \text { frames in which } q_{t-1}=i, q_{t}=j}{\# \text { frames in which } q_{t-1}=i}
$$

(3) Observation Probabilities:

$$
\bar{b}_{j}(k)=\frac{\# \text { frames in which } q_{t}=j, k_{t}=k}{\# \text { frames in which } q_{t}=j}
$$

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## Expectation Maximization

When the true state sequence is unknown, then we can't maximize the likelihood $p(O, Q \mid \bar{\lambda})$ directly. Instead, we maximize Baum's auxilary function:

$$
Q(\lambda, \bar{\lambda})=\sum_{Q} p(Q \mid O, \lambda) \ln p(O, Q \mid \bar{\lambda})
$$

This method has two key advantages:

- The maximizer of $Q(\lambda, \bar{\lambda})$ can be computed analytically.
- Baum proved that, regardless of the value of $\lambda$,

$$
\max _{\bar{\lambda}} Q(\lambda, \bar{\lambda}) \quad \Rightarrow \quad P(O \mid \bar{\lambda}) \geq P(O \mid \lambda)
$$

## Baum-Welch Re-Estimation: Overview

(1) Start out by setting $\lambda$ to any arbitrary initial value.
(2) Iterate:
(1) Find $\bar{\lambda}=\operatorname{argmax} Q(\lambda, \bar{\lambda})$
(2) Set $\lambda=\bar{\lambda}$
(3) Stop when $P(O \mid \lambda)$ stops (quickly) increasing.

## Calculating the Baum Auxiliary

The Baum auxiliary is:

$$
\begin{aligned}
Q(\lambda, \bar{\lambda}) & =\sum_{Q} p(Q \mid O, \lambda) \ln p(O, Q \mid \bar{\lambda}) \\
& =\sum_{i=1}^{N} p\left(q_{1}=i \mid O, \lambda\right) \ln \bar{\pi}_{i} \\
& +\sum_{t=1}^{T-1} \sum_{i=1}^{N} \sum_{j=1}^{N} p\left(q_{t}=i, q_{t+1}=j \mid O, \lambda\right) \ln \bar{a}_{i j} \\
& +\sum_{t=1}^{T} \sum_{i=1}^{N} \sum_{k=1}^{M} p\left(q_{t}=i, o_{t}=k \mid O, \lambda\right) \ln \bar{b}_{j}(k)
\end{aligned}
$$

Now we need to find those three probabilities.

## Calculating the Baum Auxiliary

First: $p\left(q_{1}=i \mid O, \lambda\right)$. We already know this one! It's

$$
p\left(q_{1}=i \mid O, \lambda\right)=\gamma_{1}(i)
$$

## Calculating the Baum Auxiliary

Second: $p\left(q_{t}=i, q_{t+1}=i \mid O, \lambda\right)$. This one is a two-step state posterior, calculated similar to $\gamma$. Rabiner uses the letter $\xi$ for this probability:

$$
\begin{aligned}
p\left(q_{t}=i, q_{t+1}=j \mid O, \lambda\right) & =\frac{p\left(q_{t}=i, q_{t+1}=j, O \mid \lambda\right)}{P(O \mid \lambda)} \\
& =\frac{\alpha_{t}(i) a_{i j} b_{j}\left(\vec{o}_{t+1}\right) \beta_{t+1}(j)}{P(O \mid \lambda)} \\
& \equiv \xi_{t}(i, j)
\end{aligned}
$$

## Calculating the Baum Auxiliary

Finally: $p\left(q_{t}=i, o_{t}=k \mid O, \lambda\right)$.

$$
\begin{aligned}
p\left(q_{t}=i, o_{t}=k \mid O, \lambda\right) & = \begin{cases}p\left(q_{t}=i \mid O, \lambda\right) & o_{t}=k \\
0 & \text { otherwise }\end{cases} \\
& = \begin{cases}\gamma_{t}(i) & o_{t}=k \\
0 & \text { otherwise }\end{cases}
\end{aligned}
$$

## Calculating the Baum Auxiliary

Putting it all together,

$$
\begin{aligned}
Q(\lambda, \bar{\lambda}) & =\sum_{Q} p(Q \mid O, \lambda) \ln p(O, Q \mid \bar{\lambda}) \\
& =\sum_{i=1}^{N} \gamma_{1}(i) \ln \bar{\pi}_{i} \\
& +\sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T-1} \xi_{t}(i, j) \ln \bar{a}_{i j} \\
& +\sum_{i=1}^{N} \sum_{k=1}^{M} \sum_{t: o_{t}=k} \gamma_{t}(i) \ln \bar{b}_{j}(k)
\end{aligned}
$$

## Maximizing the Baum Auxiliary

Now let's create a Lagrangian:

$$
\begin{aligned}
& \mathcal{J}(\lambda)=\sum_{i=1}^{N} \gamma_{1}(i) \ln \bar{\pi}_{i}+\sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T-1} \xi_{t}(i, j) \ln \bar{a}_{i j}+\sum_{i=1}^{N} \sum_{k=1}^{M} \sum_{t: O_{t}=k} \gamma_{t}(i) \ln \bar{b} \\
& +\kappa\left(1-\sum_{i=1}^{N} \bar{\pi}_{i}\right)+\sum_{i=1}^{N} \mu_{i}\left(1-\sum_{j=1}^{N} \bar{a}_{i j}\right)+\sum_{i=1}^{N} \nu_{i}\left(1-\sum_{k=1}^{M} \bar{b}_{j}(k)\right)
\end{aligned}
$$

...differentiate it, and set the derivative equal to zero.

## Maximizing the Baum Auxiliary

Here's the result of that differentiation:
(1) Initial State Probabilities:

$$
\bar{\pi}_{i}=\frac{\gamma_{1}(i)}{\sum_{i^{\prime}=1}^{N} \gamma_{1}\left(i^{\prime}\right)}
$$

(2) Transition Probabilities:

$$
\bar{a}_{i j}=\frac{\sum_{t=1}^{T-1} \xi_{t}(i, j)}{\sum_{j^{\prime}=1}^{N} \sum_{t=1}^{T-1} \xi_{t}\left(i, j^{\prime}\right)}
$$

(3) Observation Probabilities:

$$
\bar{b}_{j}(k)=\frac{\sum_{t: o_{t}=k} \gamma_{t}(i)}{\sum_{i^{\prime}=1}^{N} \sum_{t: o_{t}=k} \gamma_{t}\left(i^{\prime}\right)}
$$

## Maximizing the Baum Auxiliary

If you look closely at the equations on the previous slide, you will see that they are just like the known-state case, except that instead of counting known state frequencies, we now compute expected state frequencies!

## (1) Initial State Probabilities:

$$
\bar{\pi}_{i}=\frac{E\left[\# \text { state sequences that start with } q_{1}=i\right]}{\# \text { state sequences in training data }}
$$

(2) Transition Probabilities:

$$
\bar{a}_{i j}=\frac{E\left[\# \text { frames in which } q_{t-1}=i, q_{t}=j\right]}{E\left[\# \text { frames in which } q_{t-1}=i\right]}
$$

(3) Observation Probabilities:

$$
\bar{b}_{j}(k)=\frac{E\left[\# \text { frames in which } q_{t}=j, o_{t}=k\right]}{E\left[\# \text { frames in which } q_{t}=j\right]}
$$

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## Baum-Welch with Gaussian Probabilities

The requirement that we vector-quantize the observations is a problem. It means that we can't model the observations very precisely.
It would be better if we could model the observation likelihood, $b_{j}(\vec{o})$, as a probability density in the space $\vec{o} \in \Re^{D}$. One way is to use a parameterized function that is guaranteed to be a properly normalized pdf. For example, a Gaussian:

$$
b_{i}(\vec{o})=\mathcal{N}\left(\vec{o} ; \vec{\mu}_{i}, \Sigma_{i}\right)
$$

## Calculating the Baum Auxiliary

The Baum auxiliary is now:

$$
\begin{aligned}
Q(\lambda, \bar{\lambda}) & =\sum_{Q} p(Q \mid O, \lambda) \ln p(O, Q \mid \bar{\lambda}) \\
& =\sum_{i=1}^{N} \gamma_{1}(i) \ln \bar{\pi}_{i} \\
& +\sum_{i=1}^{N} \sum_{j=1}^{N} \sum_{t=1}^{T-1} \xi_{t}(i, j) \ln \bar{a}_{i j} \\
& +\sum_{i=1}^{N} \sum_{t=1}^{T} \gamma_{t}(i) \ln \mathcal{N}\left(\vec{o}_{t} ; \vec{\mu}_{i}, \Sigma_{i}\right)
\end{aligned}
$$

## Maximizing the Baum Auxiliary

When we maximize the Baum auxiliary, we get:

$$
\begin{gathered}
\bar{\mu}_{i}=\frac{\sum_{t=1}^{T} \gamma_{t}(i) \vec{o}_{t}}{\sum_{t=1}^{T} \gamma_{t}(i)} \\
\bar{\Sigma}_{i}=\frac{\sum_{t=1}^{T} \gamma_{t}(i)\left(\vec{o}_{t}-\vec{\mu}_{i}\right)\left(\vec{o}_{t}-\vec{\mu}_{i}\right)^{T}}{\sum_{t=1}^{T} \gamma_{t}(i)}
\end{gathered}
$$

## Maximizing the Baum Auxiliary

Notice the similarity to what we would do if the states were known:

- Known states: $\mu_{i}$ is the sample mean of the observations, $\Sigma_{i}$ is their sample variance.
- Known states:
- $\mu_{i}$ is the weighted average, where the weights are $\gamma_{t}(i)$.
- $\Sigma_{i}$ is the weighted sample variance, where the weights are $\gamma_{t}(i)$.


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# The Baum-Welch Algorithm: Initial and Transition Probabilities 

(1) Initial State Probabilities:

$$
\bar{\pi}_{i}=\frac{\sum_{\text {sequences }} \gamma_{1}(i)}{\# \text { sequences }}
$$

(2) Transition Probabilities:

$$
\bar{a}_{i j}=\frac{\sum_{t=1}^{T-1} \xi_{t}(i, j)}{\sum_{j=1}^{N} \sum_{t=1}^{T-1} \xi_{t}(i, j)}
$$

## The Baum-Welch Algorithm: Observation Probabilities

(1) Discrete Observation Probabilities:

$$
\bar{b}_{j}(k)=\frac{\sum_{t: \vec{o}_{t}=k} \gamma_{t}(j)}{\sum_{t} \gamma_{t}(j)}
$$

(2) Gaussian Observation PDFs:

$$
\begin{gathered}
\bar{\mu}_{i}=\frac{\sum_{t=1}^{T} \gamma_{t}(i) \vec{o}_{t}}{\sum_{t=1}^{T} \gamma_{t}(i)} \\
\bar{\Sigma}_{i}=\frac{\sum_{t=1}^{T} \gamma_{t}(i)\left(\vec{o}_{t}-\vec{\mu}_{i}\right)\left(\vec{o}_{t}-\vec{\mu}_{i}\right)^{T}}{\sum_{t=1}^{T} \gamma_{t}(i)}
\end{gathered}
$$

