

LECTURE 10: THE LAW OF TOTAL PROBABILITY, BAYES' THEOREM, AND INDEPENDENCE OF EVENTS

• TOPICS TO COVER (BASED ON CH 2.10)

→ THE LAW OF TOTAL PROBABILITY

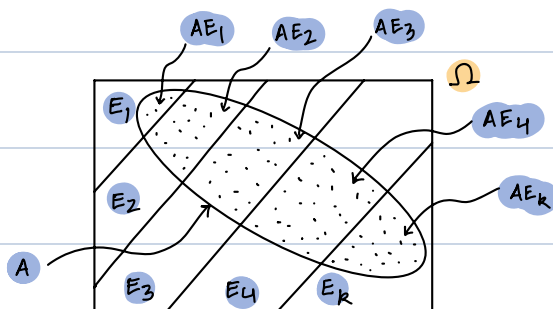
→ BAYES' THEOREM

→ INDEPENDENCE OF EVENTS

→ THE LAW OF TOTAL PROBABILITY

CONSIDER EVENTS $E_1, \dots, E_k$ SUCH THAT $E_i \cap E_j = \emptyset \quad \forall \quad i, j = 1, \dots, k \quad i \neq j$ PARTITION OF Ω

$$\bigcup_{i=1}^k E_i = \Omega$$



CONSIDER: $P(\Omega) = P\left(\bigcup_{i=1}^k E_i\right) = P(E_1) + P(E_2) + \dots + P(E_k)$ AXIOM P.2

$$\Rightarrow P(E_1) + P(E_2) + \dots + P(E_k) = 1$$

NOTICE FROM THE ABOVE VENN DIAGRAM:

$$A = AE_1 \cup AE_2 \cup \dots \cup AE_k, \quad AE_1, AE_2, \dots, AE_k \text{ ARE DISJOINT AS } E_1, E_2, \dots, E_k \text{ ARE}$$

$$\Rightarrow P(A) = P(AE_1) + P(AE_2) + \dots + P(AE_k) \quad \text{AXIOM P.2}$$

$$= P(A|E_1) \cdot P(E_1) + P(A|E_2) \cdot P(E_2) + \dots + P(A|E_k) \cdot P(E_k) \quad \begin{array}{l} \text{DEF. OF} \\ \text{COND. PROB.} \end{array}$$

$$P(A) = \sum_{i=1}^k P(A|E_i) \cdot P(E_i) : \begin{array}{l} \text{THE LAW OF TOTAL PROBABILITY} \\ \text{TOTAL PROBABILITY FORMULA (TPF)} \end{array} \dots (i)$$

RECALL, IF $P(A) \neq 0$:

$$\begin{aligned} P(E_i | A) &:= \frac{P(A|E_i)}{P(A)} && \because \text{DEF. OF COND. PROB.} \\ &= \frac{P(A|E_i) \cdot P(E_i)}{P(A)} && \because \text{DEF. OF COND. PROB.} \end{aligned}$$

$$P(E_i | A) = \frac{P(A|E_i) \cdot P(E_i)}{\sum_{i=1}^k P(A|E_i) \cdot P(E_i)} \quad \because \text{TPF} : \text{BAYES' THEOREM}$$

EXAMPLE: SEQUENTIAL EXPERIMENT

E_i : COIN i IS CHOSEN $i = 1, 2$ $P(E_1) = \frac{1}{5}$, $P(E_2) = \frac{4}{5}$

A : HEAD $P(A|E_1) = \frac{1}{3}$, $P(A|E_2) = \frac{1}{4}$

$$\begin{aligned} P(A) &= P(\text{HEAD}) \stackrel{\text{TPF}}{=} P(A|E_1) \cdot P(E_1) + P(A|E_2) \cdot P(E_2) \\ &= \underbrace{\frac{1}{3} \cdot \frac{1}{5}}_{(ii)} + \frac{1}{4} \cdot \frac{4}{5} \quad \dots (iii) \end{aligned}$$

$$\begin{aligned} P(E_1 | A) &= P(\text{YOU CHOSE COIN 1} \mid \text{EXP. RESULTED IN A HEAD}) \\ &= \frac{P(E_1 | A)}{P(A)} && \because \text{DEF. OF COND. PROB.} \end{aligned}$$

$$\begin{aligned} &= \frac{P(A|E_1) \cdot P(E_1)}{P(A|E_1) \cdot P(E_1) + P(A|E_2) \cdot P(E_2)} && \because \text{BAYES' RULE} \\ &= \frac{(ii)}{(iii)} \end{aligned}$$

INDEPENDENCE OF EVENTS

RE $\leadsto$ $(\Omega, \mathcal{F}, P)$, TWO EVENTS A and B.

1. $P(A) > 0$: $P(B|A) = \frac{P(AB)}{P(A)}$ DEF. OF COND. PRDB.

• WE CALL B TO BE INDEP. OF A IF $P(B|A) = P(B)$.
(iv)
KNOWLEDGE OF A HAS NO EFFECT ON B

NOW, $P(B|A) = \frac{P(AB)}{P(A)} = P(B)$ UNDER (iv) : B IS INDEP. OF A

$$\Rightarrow P(AB) = P(A) \cdot P(B) \quad \dots (11)$$

• SIMILARLY, WE CALL A TO BE INDEP. OF B IF $P(A|B) = P(A)$.
ASSUMING $P(B) > 0$
(v)
KNOWLEDGE OF B HAS NO EFFECT ON A

NOW, $P(A|B) = \frac{P(AB)}{P(B)} = P(A)$ UNDER (v) : A IS INDEP. OF B

$$\Rightarrow P(AB) = P(A) \cdot P(B) \quad \dots (12)$$

11 and 12 are exactly the same, i.e., UNDER $P(A) > 0$ AND $P(B) > 0$:

IND. OF A FROM B (v) : $P(A|B) = P(A) \Rightarrow P(AB) = P(A) \cdot P(B) \quad \dots (11)$

(iv) IND. OF B FROM A : $P(B|A) = P(B) \Rightarrow P(AB) = P(A) \cdot P(B) \quad \dots (12)$

D1

ALSO:

D2

IF

$$P(AB) = P(A) \cdot P(B)$$

THEN

IND. OF A FROM B

(v)

$$P(A|B) = P(A)$$

and

IND. OF B FROM A

(iv)

$$P(B|A) = P(B)$$

PROOF:

EXERCISE!

IN

SUMMARY:

IND. OF A FROM B

(v)

$$P(A|B) = P(A)$$

OR

IND. OF B FROM A

(iv)

$$P(B|A) = P(B)$$

D1

D2

$$P(AB) = P(A) \cdot P(B)$$

SO FAR

DEF.

GIVEN

$$P(A) > 0$$

AND

$$P(B) > 0:$$

A

and

B

ARE

'MUTUALLY'

INDEPENDENT

IF

$$P(AB) = P(A) \cdot P(B)$$

2.

$$P(A) = 0$$

$$P(B|A)$$

= UNDEFINED

CONSIDER

$$A^c$$

:

$$P(A^c)$$

$$= 1 - P(A)$$

$$= 1$$

DEF. OF COND. PROB.

$$P(B|A^c)$$

=

$$\frac{P(BA^c)}{P(A^c)}$$

$$P(A^c)$$

=

$$\frac{P(BA^c)}{1}$$

$$1$$

... (vi)

NOTE THAT

$$B$$

=

$$BA$$

U

$$BA^c$$

$\Rightarrow$

AXIOM P.2

$$P(B)$$

=

$$P(BA)$$

+

$$P(BA^c)$$

$$\therefore (vi) \Rightarrow P(B|A^c) = P(B) - P(AB)$$

$$= P(B) - P(A)P(B)$$

IF WE USE THE DEF. OF INDEP. SO FAR

$$\Rightarrow P(B|A^c) = P(B) \quad \therefore P(A) = 0$$

ALSO, $P(A^c|B) = P(A^c)$ (EXERCISE)

IN GENERAL $[P(A) \geq 0 \text{ AND } P(B) \geq 0]$

DEF. A and B ARE 'MUTUALLY' INDEPENDENT IF $P(AB) = P(A) \cdot P(B)$.

FINAL DEFINITION

PROPERTY 1. IF A and B ARE INDEP. THEN A^c and B ARE ALSO INDEP.

PROOF. WE NEED TO SHOW : $P(A^cB) = P(A^c) \cdot P(B)$

CONSIDER: $B = AB \cup A^cB$

$$P(B) = P(AB) + P(A^cB) \quad \therefore \text{AXIOM P.2}$$

$$\Rightarrow P(A^cB) = P(B) - P(AB)$$

$$= P(B) - P(A) \cdot P(B) \quad \therefore \text{INDEP. OF A and B}$$

$$= (1 - P(A)) \cdot P(B)$$

$$\Rightarrow P(A^cB) = P(A^c) P(B)$$

HENCE THE PROPERTY.

PROPERTY 2 :

IF A and B ARE INDEP. $\Rightarrow$ A and B^c ARE INDEP.

PROPERTY 3 :

IF A and B ARE INDEP. $\Rightarrow$ A^c and B^c ARE INDEP.

DEFINITION:

A, B, AND C ARE PAIRWISE INDEP. IF

3 EVENTS

$\Rightarrow \binom{3}{2}$ PAIRS = 3

$$P(AB) = P(A) \cdot P(B)$$

$$P(BC) = P(B) \cdot P(C)$$

$$P(AC) = P(A) \cdot P(C)$$

STRONGER THAN PAIRWISE INDEP.

DEFINITION:

A, B, AND C ARE INDEP. IF

$$P(AB) = P(A) \cdot P(B)$$

$$P(BC) = P(B) \cdot P(C)$$

$$P(AC) = P(A) \cdot P(C)$$

$$P(ABC) = P(A) \cdot P(B) \cdot P(C)$$

OBSERVE THAT :

INDEP. $\Rightarrow$ PAIRWISE INDEP.

PAIRWISE INDEP. $\nRightarrow$ INDEP.

PROPERTY 4 :

(*) A, B, AND C ARE INDEP. THEN A IS INDEP. OF $B \cdot C$.

PROOF :

$$\rightarrow P(ABC) = P(A) \cdot P(B) \cdot P(C)$$

$$\Rightarrow P(ABC) = P(A) \cdot P(BC)$$

PROPERTY 5: ^(*) A, B, AND C ARE INDEP. THEN A AND BUC ARE ALSO INDEP.

PROOF: WE NEED TO SHOW: $P(A \cdot (BUC)) = P(A) \cdot P(BUC)$

CONSIDER: $P(A \cdot (BUC)) = P(AB \cup AC)$

$$= P(AB) + P(AC) - P(AB \cdot AC) \quad \text{ADDITION THEOREM}$$

$$= P(A) \cdot P(B) + P(A) \cdot P(C) - P(A) \cdot P(B) \cdot P(C) \quad (*)$$

$$= P(A) [P(B) + P(C) - P(BC)] \quad \text{ADDITION THEOREM}$$

$$\Rightarrow P(A \cdot (BUC)) = P(A) P(BUC)$$

FINAL OBSERVATION: IF A, B, AND C ARE INDEP. THEN A (or AC) IS INDEP. OF ANY EVENT FORMED USING SET OPERATIONS ON B and C.

DEFINITION. $A_1, A_2, \dots, A_n$ are INDEP. IF ^{STRONGER THAN PAIRWISE INDEP}

$$P(A_{i_1} A_{i_2} \dots A_{i_k}) = P(A_{i_1}) \cdot P(A_{i_2}) \cdot \dots \cdot P(A_{i_k})$$

$2 \leq k \leq n$ and $1 \leq i_1 < i_2 < \dots < i_k \leq n$

THERE ARE $\binom{n}{2} + \binom{n}{3} + \dots + \binom{n}{n}$ CONDITIONS

DEFINITION. $A_1, A_2, \dots, A_n$ are PAIRWISE INDEP. IF ^{WEAKER THAN INDEP.}

$$P(A_{i_1} A_{i_2}) = P(A_{i_1}) \cdot P(A_{i_2})$$

$1 \leq i_1 < i_2 \leq n$

THERE ARE $\binom{n}{2}$ CONDITIONS

INDEPENDENCE OF RANDOM VARIABLES

DEFINITION. RANDOM VARIABLES X and Y ARE INDEP. IF ANY EVENT OF THE FORM $X \in A$ IS INEP. OF ANY EVENT OF THE FORM $Y \in B$.

$$X \in A : \{ \omega : X(\omega) \in A \} \in \mathcal{F}$$

$$Y \in B : \{ \omega : Y(\omega) \in B \} \in \mathcal{F}$$

$$P\{X \in A, Y \in B\} = P\{X \in A\} \cdot P\{Y \in B\} \quad \forall A, B \subseteq \mathbb{R}$$

PROPERTY : IF X and Y ARE INDEP. THEN

$$\underbrace{P\{X=i, Y=j\}}_{\text{JOINT PMF: } P_{XY}(i,j)} = \underbrace{P\{X=i\}}_{P_X(i)} \cdot \underbrace{P\{Y=j\}}_{P_Y(j)} \quad i, j \in \mathbb{R}$$

PMFS OF X and Y

PROOF: GIVEN THAT X AND Y ARE INDEP.

$$P\{X \in A, Y \in B\} = P\{X \in A\} \cdot P\{Y \in B\} \quad \forall A, B \subseteq \mathbb{R}$$

$$\text{TAKE } A = \{i\}, B = \{j\} \quad i, j \in \mathbb{R}$$

$$P\{X=i, Y=j\} = P\{X=i\} \cdot P\{Y=j\}$$

PROPERTY : IF

$$\underbrace{P\{X=i, Y=j\}}_{\text{JOINT PMF: } P_{XY}(i,j)} = \underbrace{P\{X=i\}}_{P_X(i)} \cdot \underbrace{P\{Y=j\}}_{P_Y(j)} \quad i, j \in \mathbb{R}$$

PMFS OF X and Y

THEN X and Y ARE INDEP.

PROOF: WE NEED TO SHOW:

$$P\{X \in A, Y \in B\} = P\{X \in A\} \cdot P\{Y \in B\} \quad \forall A, B \subseteq \mathbb{R}$$

CONSIDER:

$$P\{X \in A, Y \in B\}$$

AXIOM P.2

$$= \sum_{(i,j): i \in A, j \in B} P\{X=i, Y=j\}$$

GIVEN

$$= \sum_{(i,j): i \in A, j \in B} P\{X=i\} P\{Y=j\}$$

$$= \sum_{i: i \in A} P\{X=i\} \cdot \sum_{j: j \in B} P\{Y=j\}$$

$$\Rightarrow P\{X \in A, Y \in B\} = P\{X \in A\} \cdot P\{Y \in B\}$$