

Probabilistic Counting and Morris Counter

Lecture 04

September 1, 2022

Part I

Counting Events

Streaming model

- The input consists of m objects/items/tokens e_1, e_2, \dots, e_m that are seen one by one by the algorithm.
- The algorithm has “limited” memory say for B tokens where $B < m$ (often $B \ll m$) and hence cannot store all the input
- Want to compute interesting functions over input

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Yes, with randomization.

“Counting large numbers of events in small registers” by Rober Morris (Bell Labs), Communications of the ACM (CACM), 1978

Probabilistic Counting Algorithm

PROBABILISTIC COUNTING:

$X \leftarrow 0$

While (a new event arrives)

 Toss a biased coin that is heads with probability $1/2^X$

 If (coin turns up heads)

$X \leftarrow X + 1$

endWhile

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Theorem

Let $Y = 2^X$. Then $E[Y] - 1 = n$, the number of events seen.

$\log n$ vs $\log \log n$

Morris's motivation:

- Had 8 bit registers. Can count only up to $2^8 = 256$ events using deterministic counter. Had many counters for keeping track of different events and using 16 bits (2 registers) was infeasible.
- If only $\log \log n$ bits then can count to $2^{2^8} = 2^{256}$ events! In practice overhead due to error control etc. Morris reports counting up to 130,000 events using 8 bits while controlling error.

See 2 page paper for more details.

Analysis of Expectation

Induction on n . For $i \geq 0$, let X_i be the counter value after i events. Let $Y_i = 2^{X_i}$. Both are random variables.

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Base case: $n = 0, 1$ easy to check: $X_i, Y_i - 1$ deterministically equal to $0, 1$.

Analysis of Expectation

$$\begin{aligned} E[\mathbf{Y}_n] &= E[2^{\mathbf{X}_n}] = \sum_{j=0}^{\infty} 2^j \Pr[\mathbf{X}_n = j] \\ &= \sum_{j=0}^{\infty} 2^j \left(\Pr[\mathbf{X}_{n-1} = j] \cdot \left(1 - \frac{1}{2^j}\right) + \Pr[\mathbf{X}_{n-1} = j-1] \cdot \frac{1}{2^{j-1}} \right) \\ &= \sum_{j=0}^{\infty} 2^j \Pr[\mathbf{X}_{n-1} = j] \\ &\quad + \sum_{j=0}^{\infty} (2 \Pr[\mathbf{X}_{n-1} = j-1] - \Pr[\mathbf{X}_{n-1} = j]) \\ &= E[\mathbf{Y}_{n-1}] + 1 \quad (\text{by applying induction}) \\ &= n + 1 \end{aligned}$$

Jensen's Inequality

Definition

A real-valued function $f : \mathbb{R} \rightarrow \mathbb{R}$ is *convex* if

$f((a + b)/2) \leq (f(a) + f(b))/2$ for all a, b . Equivalently,
 $f(\lambda a + (1 - \lambda)b) \leq \lambda f(a) + (1 - \lambda)f(b)$ for all $\lambda \in [0, 1]$.

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Theorem (Jensen's inequality)

Let Z be random variable with $E[Z] < \infty$. If f is convex then
 $f(E[Z]) \leq E[f(Z)]$.

Implication for counter size

We have $Y_n = 2^{X_n}$. The function $f(z) = 2^z$ is convex. Hence

$$2^{E[X_n]} \leq E[Y_n] \leq n + 1$$

which implies

$$E[X_n] \leq \log(n + 1)$$

Hence expected number of bits in counter is $\lceil \log \log(n + 1) \rceil$.

Variance calculation

Question: Is the random variable Y_n well behaved even though expectation is right? What is its variance? Is it concentrated around expectation?

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Lemma

$E[Y_n^2] = \frac{3}{2}n^2 + \frac{3}{2}n + 1$ and hence $\text{Var}[Y_n] = n(n-1)/2$.

Variance analysis

Analyze $E[Y_n^2]$ via induction.

Base cases: $n = 0, 1$ are easy to verify since Y_n is deterministic.

$$\begin{aligned}E[Y_n^2] &= E[2^{2X_n}] = \sum_{j \geq 0} 2^{2j} \cdot \Pr[X_n = j] \\&= \sum_{j \geq 0} 2^{2j} \cdot \left(\Pr[X_{n-1} = j] \left(1 - \frac{1}{2^j}\right) + \Pr[X_{n-1} = j-1] \frac{1}{2^{j-1}} \right) \\&= \sum_{j \geq 0} 2^{2j} \cdot \Pr[X_{n-1} = j] \\&\quad + \sum_{j \geq 0} \left(-2^j \Pr[X_{n-1} = j-1] + 4 \cdot 2^{j-1} \Pr[X_{n-1} = j-1] \right) \\&= E[Y_{n-1}^2] + 3E[Y_{n-1}] \\&= \frac{3}{2}(n-1)^2 + \frac{3}{2}(n-1) + 1 + 3n = \frac{3}{2}n^2 + \frac{3}{2}n + 1.\end{aligned}$$

Error analysis via Chebyshev inequality

We have $E[Y_n] = n$ and $\text{Var}(Y_n) = n(n-1)/2$ implies $\sigma_{Y_n} = \sqrt{n(n-1)/2} \leq n$.

Applying Cheybshev's inequality:

$$\Pr[|Y_n - E[Y_n]| \geq tn] \leq 1/(2t^2).$$

Hence constant factor approximation with constant probability (for instance set $t = 1/2$).

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Question: Want estimate to be tighter. For any given $\epsilon > 0$ want estimate to have error at most ϵn with say constant probability or with probability at least $(1 - \delta)$ for a given $\delta > 0$.

Part II

Improving Estimators

Probabilistic Estimation

Setting: want to compute some real-value function f of a given input I

Probabilistic estimator: a randomized algorithm that given I outputs a random answer X such that $E[X] \simeq f(I)$. Estimator is *exact* if $E[X] = f(I)$ for all inputs I .

Additive approximation: $|E[X] - f(I)| \leq \epsilon$

Multiplicative approximation: $(1 - \epsilon)f(I) \leq E[X] \leq (1 + \epsilon)f(I)$

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Question: Estimator only gives expectation. Bound on $Var[X]$ allows Chebyshev. Sometimes Chernoff applies. How do we improve estimator?

Variance reduction via averaging

- Run h parallel copies of algorithm with *independent* randomness
- Let $Y^{(1)}, Y^{(2)}, \dots, Y^{(h)}$ be estimators from the h parallel copies
- Output $Z = \frac{1}{h} \sum_{i=1}^h Y^{(i)}$

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To run h copies need $O(\frac{1}{\epsilon^2} \log \log n)$ bits for the counters.

Error reduction via median trick

We have:

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Want:

$$\Pr[|Z_n - E[Z_n]| \geq \epsilon n] \leq \delta$$

for some given parameter δ .

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Can set $h = \frac{1}{2\epsilon^2\delta}$ and apply Chebyshev. Better dependence on δ ?

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Algorithm: Output median of $Z^{(1)}, Z^{(2)}, \dots, Z^{(\ell)}$.

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Let Z' be median of the $\ell = c \log(1/\delta)$ independent estimators.

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- For median estimate to be bad, more than half of A_i 's have to be bad.
- Using Chernoff bounds: probability of bad median is at most $2^{-c'\ell}$ for some constant c' .

Summarizing

Using variance reduction and median trick: with $O(\frac{1}{\epsilon^2} \log(1/\delta) \log \log n)$ bits one can maintain a $(1 - \epsilon)$ -factor estimate of the number of events with probability $(1 - \delta)$. This is a *generic* scheme that we will repeatedly use.

For counter one can do (much) better by changing algorithm and better analysis. See homework and references in notes.

