#### CS 498ABD: Algorithms for Big Data

## **Topics in Streaming**

Lecture 18 and 19 October 27 and 29, 2020

1

#### **Topics in Streaming**

- *F<sub>p</sub>* estimation for *p* ∈ (0, 2] via *p*-stable distributions and pseudorandom generators
- Priority Sampling
- $\bullet$  Precision Sampling and Applications to  $\ell_2$  sampling in streams
- $\ell_0$  Sampling

2

### Part III

# Sampling according to frequency moments

**Sampling problem**: given  $x \in \mathbb{R}^n$  in (strict) turnstile setting, at the end output random (I, R) where  $I \in [n]$  and  $R \in \mathbb{R}$  such that  $\Pr[I=i] \simeq \frac{|x_i|^p}{\sum_i |x_i|^p} \text{ and } R = x_i \text{ if } I = i.$  $\frac{\left|X_{i}\right|^{2}}{\left\|\chi\right\|_{2}^{2}}$  $X = (0, 0, \dots, 0)$ +2 $\begin{pmatrix} l_{\nu} \\ l_{\rho} \\ l_{\rho} \end{pmatrix} \not \sim (0, \nu)$ -10 7 100 X = (-1, 0, 0, 01, ..., )

**Sampling problem**: given  $x \in \mathbb{R}^n$  in (strict) turnstile setting, at the end output random (I, R) where  $I \in [n]$  and  $R \in \mathbb{R}$  such that  $\Pr[I = i] \simeq \frac{|x_i|^p}{\sum_j |x_j|^p}$  and  $R = x_i$  if I = i.

Sampling is generally a more challenging problem than estimation

**Sampling problem**: given  $x \in \mathbb{R}^n$  in (strict) turnstile setting, at the end output random (I, R) where  $I \in [n]$  and  $R \in \mathbb{R}$  such that  $\Pr[I = i] \simeq \frac{|x_i|^p}{\sum_j |x_j|^p}$  and  $R = x_i$  if I = i.

Sampling is generally a more challenging problem than estimation

Approximation:  $\Pr[I = i] = (1 \pm \epsilon) \frac{|x_i|^p}{\sum_j |x_j|^p} + 1/\text{poly}(n)$  for some small  $\epsilon$  and  $R = (1 \pm \epsilon)x_i$ .

**Sampling problem**: given  $x \in \mathbb{R}^n$  in (strict) turnstile setting, at the end output random (I, R) where  $I \in [n]$  and  $R \in \mathbb{R}$  such that  $\Pr[I = i] \simeq \frac{|x_i|^p}{\sum_j |x_j|^p}$  and  $R = x_i$  if I = i.

Sampling is generally a more challenging problem than estimation

Approximation:  $\Pr[I = i] = (1 \pm \epsilon) \frac{|x_i|^p}{\sum_j |x_j|^p} + 1/\text{poly}(n)$  for some small  $\epsilon$  and  $R = (1 \pm \epsilon) x_i$ .

Can do  $\ell_0$ ,  $\ell_2$  and  $\ell_p$  for 0 in polylog space using ideas from sketching. Works in (strict) turnstile models.

Several important applications

### Part IV

### $\ell_2$ Sampling

$$\begin{array}{c} x = (1, -3, 10, 5, 0, 3) \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & & & \\ & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & & & & & \\ & &$$

### **ℓ**<sub>2</sub> Sampling

Based on precision sampling which has similarities to priority sampling.

High-level Algorithm:

•  $x = (x_1, x_2, \dots, x_n)$  is the vector being updated

- Can estimate ||x||<sub>2</sub> using F<sub>2</sub> estimation. Assume ||x||<sub>2</sub> = 1 for normalization purposes/simplicity
- Consider  $y = (y_1, y_2, \dots, y_n)$  where  $y_i = x_i / \sqrt{u_i}$  where

 $u_1, u_2, \ldots, u_n$  are independent random variables from [0, 1].

• For some threshold t to be chosen, return  $(i, x_i^2)$  if i is the unique index such that  $y_i^2 \ge t$ .  $x_i^2 = u_i^2$   $y_i^2 > l^2$ 

#### Questions:

- How should we choose t? Why does it work?
- How do we implement in streaming setting?

Chandra (	UIU	JC

30

#### **Choosing threshold**

Let  $w_i = x_i^2$  and hence we have  $w_1, w_2, \dots, w_n$  and  $W = \sum_i w_i = ||x||_2^2$ . Normalize such that W = 1

Recall priority sampling where we pick  $u_1, \ldots, u_n \in [0, 1]$ independently and store the largest k amongst  $w_i/u_i$  values. Here we think of storing only largest. Also  $y_i^2 = x_i^2/u_i = w_i/u_i$ 

#### **Choosing threshold**

Let  $w_i = x_i^2$  and hence we have  $w_1, w_2, \ldots, w_n$  and  $W = \sum_i w_i = ||x||_2^2$ . Normalize such that W = 1

Recall priority sampling where we pick  $u_1, \ldots, u_n \in [0, 1]$ independently and store the largest k amongst  $w_i/u_i$  values. Here we think of storing only largest. Also  $y_i^2 = x_i^2/u_i = w_i/u_i$  $\rightarrow >_i W$ .

Fix threshold *t*. What is probability that *i* is returned?

 $\frac{\chi_{i}^{2}}{u_{i}} + \Pr[y_{i}^{2} \ge t] \prod_{j \neq i} \Pr[y_{j}^{2} < t] = \underbrace{\frac{\chi_{i}^{2}}{t}}_{t} \underbrace{\prod_{j \neq i} (1 - \frac{\chi_{j}^{2}}{t})}_{t} + \frac{\chi_{i}^{2}}{t} + \frac{\chi_$ 

Chandra (UIUC)	CS498ABD	31	Fall 2020	31 / 44
----------------	----------	----	-----------	---------

\_

$$X = (-, -, -)$$

$$Y = ( \frac{x_i}{y_i}, ...)$$

$$X_i \in X_i + (X_i)$$

$$Y_i \in Y_i + \frac{D_i}{y_{ii}}$$

*t* should be large compared to  $\sum_i x_i^2 = ||x||_2^2$ . Probability of output is 1/t so need *t* attempts. Thus choose  $t = O(\log n) ||x||_2^2$ .

*t* should be large compared to  $\sum_i x_i^2 = ||x||_2^2$ . Probability of output is 1/t so need *t* attempts. Thus choose  $t = O(\log n) ||x||_2^2$ .

Need to store  $y_1^2, y_2^2, ..., y_n^2$ ?

*t* should be large compared to  $\sum_i x_i^2 = ||x||_2^2$ . Probability of output is 1/t so need *t* attempts. Thus choose  $t = O(\log n) ||x||_2^2$ .

Need to store  $y_1^2, y_2^2, \ldots, y_n^2$ ? But we only need the <u>two largest</u> to decide if largest is above threshold. Hence can use Count Sketch on y to store only heavy hitters.

*t* should be large compared to  $\sum_i x_i^2 = ||x||_2^2$ . Probability of output is 1/t so need *t* attempts. Thus choose  $t = O(\log n) ||x||_2^2$ .

Need to store  $y_1^2, y_2^2, \ldots, y_n^2$ ? But we only need the two largest to decide if largest is above threshold. Hence can use Count Sketch on y to store only heavy hitters.



Note that  $y_i^2 \ge x_i^2$  for all *i*, hence  $||y||_2^2 \ge ||x||_2^2$ .

#### Lemma

With probability  $\geq (1 - \delta)$  we have  $||y||_2^2 \leq \frac{1}{\delta}c \ln n ||x||_2^2$  for some fixed c.

Prove above as exercise. Thus  $||y||_2$  is not much larger than  $||x||_2$ .

$$y = \begin{pmatrix} x_{1} \\ y_{1} \end{pmatrix}, \begin{pmatrix} x_{1} \\ y_{n} \end{pmatrix}, \dots, \begin{pmatrix} x_{n} \\ \sqrt{u_{n}} \end{pmatrix}, \dots, \begin{pmatrix} x_{n} \\ \sqrt{u_{n}} \end{pmatrix}, \dots, \begin{pmatrix} y_{n} \\ y_{n} \end{pmatrix}, \begin{pmatrix} x_{n} \\ y_{n} \end{pmatrix}, \begin{pmatrix} x_{n} \\ y_{n} \end{pmatrix}, \begin{pmatrix} x_{n} \\ y_{n} \end{pmatrix}, \dots, \begin{pmatrix} x_{n} \\ y_{n} \end{pmatrix}, \dots,$$

Note that  $y_i^2 \ge x_i^2$  for all *i*, hence  $||y||_2^2 \ge ||x||_2^2$ .

#### Lemma

With probability  $\geq (1 - \delta)$  we have  $||y||_2^2 \leq \frac{1}{\delta}c \ln n ||x||_2^2$  for some fixed c.

Prove above as exercise. Thus  $||y||_2$  is not much larger than  $||x||_2$ .

Recall Count Sketch for y gives estimate  $|\tilde{y}_i|$  for each *i* such that  $|\tilde{y}_i - y_i|^2 \le \epsilon^2 ||y||_2^2$  and space is  $O(\frac{1}{\epsilon^2} \log n)$ . Choose  $\epsilon = \epsilon' / \log n$  and hence we have  $|\tilde{y}_i - y_i|^2 \le \frac{\epsilon'}{\log n} ||x||_2^2$  $|\tilde{y}_i - \tilde{y}_i|^2 \le (\frac{\epsilon'}{\log n})^2$ .  $||\tilde{y}_i|_2^2 \le (\frac{\epsilon'}{\log n})^2$  form.  $||x||_2^2$ 

Note that  $y_i^2 \ge x_i^2$  for all *i*, hence  $||y||_2^2 \ge ||x||_2^2$ .

#### Lemma

With probability  $\geq (1 - \delta)$  we have  $||y||_2^2 \leq \frac{1}{\delta}c \ln n ||x||_2^2$  for some fixed c.

Prove above as exercise. Thus  $||y||_2$  is not much larger than  $||x||_2$ .

Recall Count Sketch for y gives estimate  $\tilde{y}_i$  for each i such that  $|\tilde{y}_i - y_i|^2 \le \epsilon^2 ||y||_2^2$  and space is  $O(\frac{1}{\epsilon^2} \log n)$ . Choose  $\epsilon = \epsilon' / \log n$  and hence we have  $|\tilde{y}_i - y_i|^2 \le \frac{\epsilon'}{\log n} ||x||_2^2$ 

Above implies that  $\tilde{y}_i$  is a close mutiplicative approximation of  $y_i$  if  $y_i$  is sufficiently large compared to  $||x||_2^2$ 

Recall threshold  $t = c \log n ||x||_2^2$ . Implies that

- Sufficient to keep track of small number of heavy hitters in y hence Count Sketch for y needs only poly( $\log n/\epsilon^2$ ) space.
- Can keep track of  $||x||_2$  and  $||y||_2$  to check if heavy hitters are sufficiently large and hence estimates are accurate even if additive error
- Output *i* if  $\tilde{y}_i^2 \ge t$  and is unique.

Recall threshold  $t = c \log n ||x||_2^2$ . Implies that

- Sufficient to keep track of small number of heavy hitters in y hence Count Sketch for y needs only poly( $\log n/\epsilon^2$ ) space.
- Can keep track of  $||x||_2$  and  $||y||_2$  to check if heavy hitters are sufficiently large and hence estimates are accurate even if additive error
- Output *i* if  $\tilde{y}_i^2 \ge t$  and is unique.

Since we use  $\tilde{y}_i$  which is an estimate of  $y_i$ , the probability of *i* being output is proportional to  $\frac{(1\pm\epsilon)x_i^2}{\|x\|_2^2}$ .

How do we sketch y without storing  $u_1, \ldots, u_n$ ? Recall analysis crucially relied on independence.

How do we sketch y without storing  $u_1, \ldots, u_n$ ? Recall analysis crucially relied on independence.

- Use *k*-wise independence for sufficiently large *k* and redo analysis
- Use hammer of pseudorandom generators

#### Algorithm again

- x is vector being updated. Keep track of  $||x||_2$
- Use Count Sketch to sketch y where  $y_i = x_i/\sqrt{u_i}$  with  $u_i$  drawn independently from [0, 1]. Use sketch to obtain estimates  $\tilde{y}_i$  for heavy hitters in y
- Output *i* if  $\tilde{y}_i^2$  is the unique heavy hitter that is above threshold
  - t where  $t = c \log n ||x||_2^2$ . If no such *i* then declare FAIL.

Repeat above in parallel  $O(\log^2 n)$  times to guarantee high probability of obtaining a good sample.

#### Algorithm again

- x is vector being updated. Keep track of  $||x||_2$
- Use Count Sketch to sketch y where  $y_i = x_i/\sqrt{u_i}$  with  $u_i$  drawn independently from [0, 1]. Use sketch to obtain estimates  $\tilde{y}_i$  for heavy hitters in y
- Output *i* if  $\tilde{y}_i^2$  is the unique heavy hitter that is above threshold t where  $t = c \log n ||x||_2^2$ . If no such *i* then declare FAIL.
- t where  $t = c \log n ||x||_2^2$ . If no such *I* then declare FAIL.

Repeat above in parallel  $O(\log^2 n)$  times to guarantee high probability of obtaining a good sample.

Space is for Count Sketch and to store generate  $u_i$  values pseudorandomly.

#### Algorithm again

- x is vector being updated. Keep track of  $||x||_2$
- Use Count Sketch to sketch y where  $y_i = x_i/\sqrt{u_i}$  with  $u_i$  drawn independently from [0, 1]. Use sketch to obtain estimates  $\tilde{y}_i$  for heavy hitters in y
- Output *i* if  $\tilde{y}_i^2$  is the unique heavy hitter that is above threshold t where  $t = c \log n ||x||_2^2$ . If no such *i* then declare FAIL.
- Repeat above in parallel  $O(\log^2 n)$  times to guarantee high

probability of obtaining a good sample.

Space is for Count Sketch and to store generate  $u_i$  values pseudorandomly.

Algorithm uses  $poly(\log n/\epsilon)$  space and with high probability outputs  $i \in [n]$  such that  $Pr[i \text{ is output}] = (1 \pm \epsilon)x_i^2/||x||_2^2 + 1/n^c$ .

Chandra (UIUC)	CS498ABD	36	Fall 2020	36 / 44
· · · ·				

#### Application of $\ell_2$ sampling to $F_p$ estimation

For p > 2 AMS-Sampling gives algorithm to estimate  $F_p$  using  $\tilde{O}(n^{1-1/p})$  space. Optimal space is  $\tilde{O}(n^{1-2/p})$ . Can estimat Ep & lp & p =0 and p = (0,2] in polylog space  $\mathcal{J}(n^{l-\frac{2}{p}})$ p>2

### Application of $\ell_2$ sampling to $F_p$ estimation

For p > 2 AMS-Sampling gives algorithm to estimate  $F_p$  using  $\tilde{O}(n^{1-1/p})$  space. Optimal space is  $\tilde{O}(n^{1-2/p})$ .

Use l<sub>2</sub> sampling algorithm to generate (i, |x̃<sub>i</sub>|)
Estimate ||x||<sub>2</sub><sup>2</sup>.

• Output 
$$\underline{T} = ||x_2||^2 |\tilde{x_i}|^{p-2}$$
 as estimate

To simplify analysis/notation assume sampling is exact.

$$\mathsf{E}[\mathsf{T}] = \frac{||\mathsf{x}||_2^2}{|\mathsf{x}|_2} \sum_i \frac{x_i^2}{||\mathsf{x}||_2^2} |x_i|^{p-2} = \sum_i |x_i|^p$$

### Application of $\ell_2$ sampling to $F_p$ estimation

For p > 2 AMS-Sampling gives algorithm to estimate  $F_p$  using  $\tilde{O}(n^{1-1/p})$  space. Optimal space is  $\tilde{O}(n^{1-2/p})$ .

- Use  $\ell_2$  sampling algorithm to generate  $(i, |\tilde{x}_i|)$  • Estimate  $||x||_2^2$  or estimate  $(i, |\tilde{x}_i|)$  (parce -

To simplify analysis/notation assume sampling is exact.  $\mathbf{E}[\mathbf{T}] = \|\mathbf{x}\|_2^2 \sum_i \frac{|\mathbf{x}_i|^2}{\|\mathbf{x}\|_2^2} |\mathbf{x}_i|^{p-2} = \sum_i |\mathbf{x}_i|^p$  $Var[T] \le ||x||_2^4 \sum_i \frac{x_i^2}{||x||_2^2} x_i^{2(p-2)} \le ||x||_2^2 \sum_i x_i^{2p-2} \le \frac{n^{1-2/p}}{\sum_i |x_i|^p} \sum_i x_i^{2p-2} \sum_i x_i^{2p-2} \ge \frac{n^{1-2/p}}{\sum_i |x_i|^p} \sum_i x_i^{2p-2} \sum_i x_i^$ Now do average plus median hith

### $\mathsf{Part}\ \mathsf{V}$

### $\ell_0$ Sampling

Chandra (UIUC)

### $\ell_0$ Sampling

Turnstile stream: x updated with positive and negative entries

At end of stream want to sample uniformly a coordinate i among all *non-zero* coordinates in x

Special case: sampling a uniform distinct element in cash register model

$$\begin{array}{l} \chi_{-} & (0, 0, 0, 0) \\ & (0, -1, 1, 0) \\ & (0, 0, 2, 0) \\ & (1, 0, -1, -1) \\ & (-1, 0, 0, 3) \end{array}$$

### $\ell_0$ Sampling

Turnstile stream: x updated with positive and negative entries

At end of stream want to sample uniformly a coordinate i among all *non-zero* coordinates in x

Special case: sampling a uniform distinct element in cash register model

Goal: illustrate a simple algorithm via two powerful hammers

#### Sparse Recovery

Recall sparse recovery using Count Sketch.

#### Theorem

There is a linear sketch with size  $O(\frac{k}{\epsilon^2} \text{polylog}(n))$  that returns z such that  $||z||_0 \le k$  and with high probability  $||x - z||_2 \le (1 + \epsilon) \operatorname{err}_2^k(x)$ .

$$\operatorname{err}_{2}^{k}(x) = \min_{z:||z||_{0} \leq k} ||x - z||_{2}$$

Hence space is proportional to desired output. Assumption k is typically quite small compared to n, the dimension of x.

Note that if x is k-sparse vector is exactly reconstructed

Chandra (U	IUC	)
------------	-----	---

X

x is updated in turnstile streaming fashion. Let J be the non-zero indices of x  $(-1, 0, 0, 3, 0) = \int = \{1, 4\}$ .

Suppose we knew |J| is small, say  $\leq s$ . Then can use sparse recovering with  $\tilde{O}(s)$  space to completely recover x and can then sample uniformly.  $\leq shum(n)$ 

 $\boldsymbol{x}$  is updated in turnstile streaming fashion. Let  $\boldsymbol{J}$  be the non-zero indices of  $\boldsymbol{x}$ 

Suppose we knew |J| is small, say  $\leq s$ . Then can use sparse recovering with  $\tilde{O}(s)$  space to completely recover x and can then sample uniformly.

What if **J** is large?

- Guess |J| to within factor of **2**.
- More formally, for j = 0 to log n let l<sub>j</sub> be n/2<sup>j</sup> coordinates of [n] sampled uniformly at random. Note l<sub>0</sub> = [n].
- Let  $y^j$  be vector obtained by restricting x to coordinates in  $I_j$ .  $y^0 = x$ .

$$\begin{array}{l} \chi_{=} \left( \chi_{1}, \chi_{2}, -\cdot, \chi_{e} \right) & \\ y'' = \left( \chi_{1}, \chi_{2}, \ldots, \chi_{e} \right) & \\ y'' = \left( \chi_{1}, \chi_{2}, \chi_{4}, \chi_{e} \right) & \\ y'' = \left( \chi_{1}, \chi_{2}, \chi_{4}, \chi_{e} \right) & \\ y'' = \left( \chi_{2}, \chi_{4} \right) & \\ y'' = \left( \chi_{2}, \chi_{4} \right) & \\ y'' = \left( \chi_{2}, \chi_{4} \right) & \\ y'' = \left( \chi_{2} \right) & \\ y'' = \left( \chi_{2} \right) & \\ \end{array}$$

Choose 
$$s = \Omega(\log(1/\delta))$$
.

- For  $j = 0, 1, \ldots, \log n$ 
  - Use *s*-sparse recovery on *y<sup>j</sup>*.
  - If y<sup>j</sup> is not s-sparse discard. Else pick a random non-zero coordinate in y<sup>j</sup> and output it. And stop.

Choose  $s = \Omega(\log(1/\delta))$ .

- For  $j = 0, 1, ..., \log n$ 
  - Use *s*-sparse recovery on *y<sup>j</sup>*.
  - If y<sup>j</sup> is not s-sparse discard. Else pick a random non-zero coordinate in y<sup>j</sup> and output it. And stop.

Uses  $O(\log n)$  s-sparse recovery data structures and hence space is poly-logarithmic assuming  $\delta$  is  $\Omega(n^{-c})$  for some fixed constant c.

Choose  $s = \Omega(\log(1/\delta))$ .

#### For $j = 0, 1, \ldots, \log n$

• Use *s*-sparse recovery on *y<sup>j</sup>*.

If y<sup>j</sup> is not s-sparse discard. Else pick a random non-zero coordinate in y<sup>j</sup> and output it. And stop.

Uses  $O(\log n)$  s-sparse recovery data structures and hence space is poly-logarithmic assuming  $\delta$  is  $\Omega(n^{-c})$  for some fixed constant c.

How can we implement random coordinates of x? Cannot store them. So how can we run sparse recovery on  $y^j$ ?

Choose  $s = \Omega(\log(1/\delta))$ .

#### For $j = 0, 1, \ldots, \log n$

• Use *s*-sparse recovery on *y<sup>j</sup>*.

If y<sup>j</sup> is not s-sparse discard. Else pick a random non-zero coordinate in y<sup>j</sup> and output it. And stop.

Uses  $O(\log n)$  s-sparse recovery data structures and hence space is poly-logarithmic assuming  $\delta$  is  $\Omega(n^{-c})$  for some fixed constant c.

How can we implement random coordinates of x? Cannot store them. So how can we run sparse recovery on  $y^j$ ? Use Nisan's generator!

Question: Will algorithm output a random non-zero coordinate?

Question: Will algorithm output a random non-zero coordinate?

#### Lemma

Suppose  $|J| \leq s$  then algorithm outputs a uniform non-zero coordinate of x with high probability.

 $y^0 = x$  is *s*-sparse. Sparse recovery algorithm succeeds with high probability.

$$I_{J} \qquad J = I_{L} \qquad K \qquad I_{k}$$

Question: Will algorithm output a random non-zero coordinate?

#### Lemma

Suppose  $|J| \leq s$  then algorithm outputs a uniform non-zero coordinate of x with high probability.

 $y^0 = x$  is *s*-sparse. Sparse recovery algorithm succeeds with high probability.

#### Lemma

Assume |J| > s. There is an index k such that with probability  $(1 - \delta)$ ,  $y^k$  is s-sparse and has at least one non-zero coordinate.

Question: Will algorithm output a random non-zero coordinate?

#### Lemma

Suppose  $|J| \leq s$  then algorithm outputs a uniform non-zero coordinate of x with high probability.

 $y^0 = x$  is *s*-sparse. Sparse recovery algorithm succeeds with high probability.

#### Lemma

Assume |J| > s. There is an index k such that with probability  $(1 - \delta)$ ,  $y^k$  is s-sparse and has at least one non-zero coordinate.

Expected number of coordinates of J in  $y^j$  is  $|J|/2^j$ . Find j such that expected number is between s/4 and s and use Chernoff bound.

#### **Analysis continued**

#### Lemma

Assume |J| > s. There is an index k such that with probability  $(1 - \delta)$ ,  $y^k$  is s-sparse and has at least one non-zero coordinate.

*s*-sparse recovery of  $y^k$  will reconstruct it exactly.  $y^k$  has random sample of coordinates of x hence has random sample of non-zero coordinates as well. Output random non-zero coordinate of  $y^k$ .

Algorithm fails only if every  $y^{j}$  fails sparse recovery and |J| > 0 but we see that  $y^{k+1}$  succeeds with probability at least  $(1 - \delta)$ .

44