# Binomial Distribution

- Binomially-distributed random variable *X*  equals sum (number of successes) of n independent Bernoulli trials
- The probability mass function is:  $f(x) = C_x^n p^x (1-p)^{n-x}$  for  $x = 0,1,...n$  (3-7)
- Based on the binomial expansion:



Binomial mean, variance and standard deviation

Let *X* be a binomial random variable with parameters *p* and *n*

- Mean:

*μ=np*

- Variance:
- σ 2 $2 = V(X) = np(1-p)$
- Standard deviation:

σ <sup>=</sup>*np*(1 −*p* )

- Standard deviation to mean ratio

$$
\sigma/\mu = \sqrt{np(1-p)}/np = \frac{\sqrt{(1-p)/p}}{\sqrt{n}}
$$

# Poisson Distribution

• Limit of the binomial distribution when  $\mathcal{L}_{\mathcal{A}}$  , where  $\mathcal{L}_{\mathcal{A}}$  is the set of the *n* , the number of attempts, is very large  $\mathcal{L}_{\mathcal{A}}$  , and the set of  $\mathcal{L}_{\mathcal{A}}$  *p , the probability of success* is very small –*E(X)= n p =λ* is O(1)

The annual numbers of deaths from horse kicks in 14 Prussian army corps between 1875 and 1894



#### and physicist From von Bortkiewicz <sup>1898</sup>



Siméon Denis Poisson (1781–1840) French mathematician

Let 
$$
\lambda = np = E(x)
$$
, so  $p = \frac{\lambda}{n}$   
\n
$$
P(X = x) = {n \choose x} p^{x} (1-p)^{n-x}
$$
\n
$$
= \frac{n(n-1)...(n-x+1)}{x!} \left(\frac{\lambda}{n}\right)^{x} \left(1-\frac{\lambda}{n}\right)^{n-x} \sim \frac{n^{x}}{x!} \left(\frac{\lambda}{n}\right)^{x} = \frac{\lambda^{x}}{x!};
$$
\n
$$
\sum_{x} \frac{\lambda^{x}}{x!} = e^{\lambda}.
$$
\nNormalization requires

\n
$$
\sum_{x} P(Y = x) = 1
$$

Normalization requires  $\sum_{x} P(X = x) = 1$ .<br>Thus  $P(X = x) = \frac{\lambda^x}{x!} e^{-\lambda}$ 

# Poisson Mean & Variance

If X is a Poisson random variable, then:

- Mean:  $\mu = E(X) = \lambda$
- Variance: σ  $2 = V(X) = λ$
- Standard deviation:  $\sigma = \lambda^{1/2}$

Note: Variance = MeanNote: Standard deviation/Mean =  $\lambda$ <sup>-1/2</sup> decreases with λ

## Matlab exercise: Poisson distribution

- Generate a sample of size 100,000 for Poissondistributed random variable X with λ =2
- Plot the **approximation** to the Probability Mass Function based on this sample
- Calculate the mean and variance of this sample and compare it to theoretical calculations:

E[X]= λ and V[X]= λ

## Matlab exercise: Poisson distribution

- **Stats=100000; lambda=2;**
- **r2=random('Poisson',lambda,Stats,1);**
- •**mu\_p=sum(r2)./Stats;**
- **disp(mu\_p);**
- •**var\_p=sum((r2-mu\_p).^2)./Stats;**
- **disp(var\_p);**
- **std\_p=sqrt(var\_p)**
- **[a,b]=hist(r2, 0:max(r2));**
- •**p\_p=a./sum(a);**
- **figure; stem(b,p\_p);**
- **figure; semilogy(b,p\_p,'ko-');**



Poisson Distribution in Genome Assembly

**Cost per Raw Megabase of DNA Sequence** 



## Poisson Example: Genome Assembly

- Goal: DNA sequence of the entire genome of an organism
- Problem: Sequencers generate short reads of random portions of a genome
- Solution: assemble genome from short reads using computers
- Whole Genome Shogun Assembly pioneered by Craig Venter in 1990s
- The human genome was jointly announced in 2001 by the Human Genome Project (public) and Celera Genomics (Craig Venter's company)

# Short Reads assemble into Contigs



Figure 5.1.



# Current sequencing technologies





MinION, a palm-sized gene sequencer made by UK-based Oxford Nanopore Technologies

## Promise of Genomics



Drew Sheneman, New Jersey -- The Newark Star Ledger, E-mail Drew.

### I think I found the corner piece!

### How many short reads do we need?**Input Output Low coverage:** A few pieces to many contigs, assemble many gaps **High coverage:** a few contigs, a many pieces to assemble few gaps

### Genome Assembly

Whole-genome "shotgun" sequencing starts by copying and fragmenting the DNA

("Shotgun" refers to the random fragmentation of the whole genome; like it was fired from ashotgun)

Input: GGCGTCTATATCTCGGCTCTAGGCCCTCATTTTTT 35bp

- **Copy** GGCGTCTATATCTCGGCTCTAGGCCCTCATTTTTT
- by GGCGTCTATATCTCGGCTCTAGGCCCTCATTTTTT
- PCR:GGCGTCTATATCTCGGCTCTAGGCCCTCATTTTTTGGCGTCTATATCTCGGCTCTAGGCCCTCATTTTTT

Fragment: GGCGTCTA TATCTCGG CTCTAGGCCCTC ATTTTTT GGC GTCTATAT CTCGGCTCTAGGCCCTCA GGCGTC TATATCT CGGCTCTAGGCCCT CATTTTTTGGCGTCTAT ATCTCGGCTCTAG GCCCTCA

Courtesy of Ben Langmead. Used with permission.

### Assembly

Assume sequencing produces such <sup>a</sup> large # fragments that almost all genome positions are *covered* by manyfragments...

...but we don't know what came fromwhere

Reconstruct this

CTAGGCCCTCAATTTTTGGCGTCTATATCTCTCTAGGCCCTCAATTTTTTCTATATCTCGGCTCTAGGGGCTCTAGGCCCTCATTTTTTCTCGGCTCTAGCCCCTCATTTTTATCTCGACTCTAGGCCCTCAGGCGTCGATATCTTATCTCGACTCTAGGCCGGCGTCTATATCTCG

From these

GGCGTCTATATCTCGGCTCTAGGCCCTCATTTTTT

Courtesy of Ben Langmead. Used with permission.

### Assembly

Overlaps between short reads help to put them together

177 nucleotidesCTAGGCCCTCAATTTTTCTCTAGGCCCTCAATTTTTGGCTCTAGGCCCTCATTTTTTCTCGGCTCTAGCCCCTCATTTTTATCTCGACTCTAGGCCCTCATATCTCGACTCTAGGCCTCTATATCTCGGCTCTAGG GGCGTCTATATCTCGGGCGTCGATATCTGGCGTCTATATCTGGCGTCTATATCTCGGCTCTAGGCCCTCATTTTTT35 nucleotides

Courtesy of **Ben Langmead**. Used with permission. The endotermies of the end of

# Where is the Poisson?

- •*G - genome length (in bp)*
- •*L - short read average length*
- •*N – number of short read sequenced*
- •*λ – sequencing coverage redundancy = LN/G*
- •*x- number of short reads covering a given site on the genome*

$$
P(x) = \frac{\lambda^x e^{-\lambda}}{x!}
$$

Poisson as a limit of Binomial: For a given site on the genome for each short read Prob(site covered): p=L/G is very small. Number of attempts (short reads): N is very large. Their product (sequencing redundancy): λ = NL/G is O(1).



### What fraction of the genome is missing?

- 
- 
- 
- 
- -
- -

# What fraction of genome is covered?

• Coverage: *λ=NL/G, X – random variable equal to the number of times a given site is covered by short reads. Poisson: P(X=x)= λ xexp(- λ)/x! P(X=0)=exp(- λ), P(X>0)=1- exp(- λ )*

• *Total length covered: G\*[1- exp(- λ)]*

Mean proportion   .864665 .981684 .997521 .999665 .999955 .999994 of genome covered			

Table 5.1. The mean proportion of the genome covered for different values of  $\lambda$ 



**G**

If DNA was a random chain with  $p_{A}$ =  $p_{C}$ =  $p_{G}$ =  $p_{T}$ =1/4  $L_{\alpha}$ <sup>~</sup>16-20 would be enough  $2\cdot \mathsf{G}\cdot4^{\mathsf{\textrm{-}low}}$  =2  $\cdot$  3x10 $^{\textrm{9}}\cdot4^{\mathsf{\textrm{-}16}}$ =1.4  $2\cdot 3$ x $10^9\cdot 4^{\text{-}20}$ =0.0055<<1



G

P(short read can be extended by another short read) =  $\frac{L - L_o}{G}$  = p P(short read cannot be extended by any short reads)= $e^{-pN} \approx Ne^{-\lambda}$ 

number of contigs=
$$
Ne^{-pN} \approx Ne^{-\lambda}
$$



## How many contigs?

- A given short read is the right end of a contig if and only if no left ends of other short reads fall within it.
- The left end of another short read has the probability *p=(L-1)/G* to fall within a given read. There are *N-1* other reads. Hence the expected number of left ends inside a given shot read is *p· (N-1)=(N-1) ·(L-1)/G ≈λ*
- If significant overlap required to merge two short reads is *Lov*, modified *λ* is given by (*N-1) ·(L- Lov)/G*
- Probability that no left ends fall inside a short read is *exp(- λ).* Thus the Number of contigs is  $N_{contigs}$ =Ne<sup>- λ</sup>:

	$\begin{array}{ccccccccc} \n0.5 & 0.75 & 1 & 1.5 & 2 & 3 & 4 & 5 & 6 & 7\n\end{array}$					
Mean number 60.7 70.8 73.6 66.9 54.1 29.9 14.7 6.7 3.0 1.3 of contigs						

Table 5.2. The mean number of contigs for different levels of coverage, with  $G = 100,000$  and  $L = 500$ .

# Average length of a contig?

- Length of a genome covered: *Gcovered=G· P(X>0)=G · (1- exp(- λ))*
- Number of contigs N<sub>contigs</sub>=N  $\cdot$  e<sup>-λ</sup>
- Average length of a contig <sup>=</sup>

$$
=\sum_{i}L_{i}/N_{contigs}=G_{covered}/N_{contigs}=
$$

*G · (1- exp(- λ))/ N · e <sup>λ</sup>=L · (1- exp(- λ))/ λ · e - λ*

Mean contig 1,600 6,700 33,500 186,000 1,100,000 size			

Table 5.3. The mean contig size for different values of a for the case  $L = 500$ .