Marginal Probability Distributions (discrete)

For a discrete joint PDF, there are marginal distributions for each random variable, formed by summing the joint PMF over the other variable.

$$f_X(x) = \sum_y f_{XY}(x, y)$$
$$f_Y(y) = \sum_x f_{XY}(x, y)$$

Called marginal because they are written in the margins

Figure 5-6 From the prior example, the joint PMF is shown in green while the two marginal PMFs are shown in purple.

y = number of	x = number of bars of			
times city name	signal strength			
is stated	1	2	3	$f_{Y}(y) =$
1	0.01	0.02	0.25	0.28
2	0.02	0.03	0.20	0.25
3	0.02	0.10	0.05	0.17
4	0.15	0.10	0.05	0.30
$f_X(x) =$	0.20	0.25	0.55	1.00

Conditional Probability Distributions

Recall that $P(B|A) = \frac{P(A \cap B)}{P(A)}$

P(Y=y | X=x)=P(X=x,Y=y)/P(X=x)== $f(x,y)/f_X(x)$

From Example 5-1 x = number of bars of y = number of signal strength P(Y=1|X=3) = 0.25/0.55 = 0.455times city name 1 2 3 $|f_{Y}(y)| =$ is stated P(Y=2|X=3) = 0.20/0.55 = 0.3640.01 0.02 0.25 0.28 1 0.20 0.03 0.25 2 0.02 P(Y=3|X=3) = 0.05/0.55 = 0.0910.05 3 0.02 0.10 0.17 P(Y=4|X=3) = 0.05/0.55 = 0.0910.05 4 0.15 0.10 0.30 $f_X(x) =$ 0.20 0.25 0.55 1.00 Sum = 1.00

Note that there are 12 probabilities conditional on X, and 12 more probabilities conditional upon Y.

X and Y are Bernoulli variables

	Y=0	Y=1
X=0	2/6	1/6
X=1	2/6	1/6

What is the marginal $P_{\gamma}(Y=0)$?

- A. 1/6B. 2/6
- C. 3/6
- D. 4/6

E. I don't know

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X and Y are Bernoulli variables

	Y=0	Y=1
X=0	2/6	1/6
X=1	2/6	1/6

What is the conditional P(X=0|Y=1)?

E. I don't know

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Reminder

Statistically independent events Always true: $P(A \cap B) = P(A \mid B) \cdot P(B) = P(B \mid A) \cdot P(A)$

Two events

Two events are **independent** if any one of the following equivalent statements is true:

$$(1) \quad P(A|B) = P(A)$$

$$(2) \quad P(B|A) = P(B)$$

$$(3) \quad P(A \cap B) = P(A)P(B)$$

Multiple events

The events E_1, E_2, \ldots, E_n are independent if and only if for any subset of these events $E_{i_1}, E_{i_2}, \ldots, E_{i_k}$,

 $P(E_{i_1} \cap E_{i_2} \cap \cdots \cap E_{i_k}) = P(E_{i_1}) \times P(E_{i_2}) \times \cdots \times P(E_{i_k})$

Independence of Random Variables X and Y

 Random variable independence means that knowledge of any of the values of X does not change probabilities of any of the values of Y

 Opposite: Dependence implies that some values of X influence the probability of some values of Y Independence for Discrete Random Variables

- Remember independence of events (slide 13 lecture 4) : Events are independent if any one of the three conditions are met:
 1) P(A | B)=P(A ∩ B)/P(B)=P(A) or
 2) P(B | A)= P(A ∩ B)/P(A)=P(B) or
 3) P(A ∩ B)=P(A) · P(B)
- Random variables independent if <u>all events</u> *A* that *Y*=*y* and *B* that *X*=*x* are independent if any one of these conditions is met: 1) P(Y=y | X=x)=P(Y=y) for any *x* or 2) P(X=x | Y=y)=P(X=x) for any *y* or 3) P(X=x, Y=y)=P(X=x)·P(Y=y) <u>for every pair x and y</u>

X and Y are Bernoulli variables

	Y=0	Y=1
X=0	2/6	1/6
X=1	2/6	1/6

Are they independent?



- B. no
- C. I don't know

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X and Y are Bernoulli variables

	Y=0	Y=1
X=0	1/2	0
X=1	0	1/2

Are they independent?

A. yesB. noC. I don't know

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Joint Probability Density Function Defined

The joint probability density function for the continuous random variables X and Y, denotes as $f_{XY}(x,y)$, satisfies the following properties:

(1)
$$f_{XY}(x,y) \ge 0$$
 for all x, y
(2) $\int_{-\infty} \int_{-\infty} \int_{-\infty} f_{XY}(x,y) dx dy = 1$
(3) $P((X,Y) \subset R) = \iint_{R} f_{XY}(x,y) dx dy$ (5-2)



Figure 5-2 Joint probability density function for the random variables X and Y. Probability that (X, Y) is in the region R is determined by the volume of $f_{XY}(x,y)$ over the region R.

Joint Probability Density Function Graph



Figure 5-3 Joint probability density function for the continuous random variables *X* and *Y* of expression levels of two different genes. Note the asymmetric, narrow ridge shape of the PDF – indicating that small values in the *X* dimension are more likely to occur when small values in the *Y* dimension occur.

Marginal Probability Distributions (continuous)

- Rather than summing a discrete joint PMF, we integrate a continuous joint PDF.
- The marginal PDFs are used to make probability statements about one variable.
- If the joint probability density function of random variables X and Y is f_{XY}(x,y), the marginal probability density functions of X and Y are:

$$f_X(x) = \int_{y} f_{XY}(x, y) \, dy \qquad f_X(x) = \sum_{y} f_{XY}(x, y)$$
$$f_Y(y) = \int_{x} f_{XY}(x, y) \, dx \qquad (5-3) \qquad f_Y(y) = \sum_{x} f_{XY}(x, y)$$

Conditional Probability Density Function Defined

Given continuous random variables *X* and *Y* with joint probability density function $f_{XY}(x, y)$, the conditional probability densiy function of *Y* given *X*=x is $f_{Y|x}(y) = \frac{f_{XY}(x, y)}{f_X(x)} = \frac{f_{XY}(x, y)}{\int_{Y} f_{XY}(x, y) \, dy}$ if $f_X(x) > 0$ (5-4)

which satifies the following properties:

(1)
$$f_{Y|x}(y) \ge 0$$

(2) $\int f_{Y|x}(y)dy = 1$
(3) $P(Y \subset B|X = x) = \int_{B} f_{Y|x}(y)dy$ for any set B in the range of Y

Compare to discrete: $P(Y=y|X=x)=f_{XY}(x,y)/f_X(x)$

Conditional Probability Distributions

- Conditional probability distributions can be developed for multiple random variables by extension of the ideas used for two random variables.
- Suppose p = 5 and we wish to find the distribution of X_1, X_2 and X_3 conditional on $X_4 = x_4$ and $X_5 = x_5$.

$$f_{X_1X_2X_3|x_4x_5}(x_1, x_2, x_3) = \frac{f_{X_1X_2X_3X_4X_5}(x_1, x_2, x_3, x_4, x_5)}{f_{X_4X_5}(x_4, x_5)}$$

for $f_{X_4X_5}(x_4, x_5) > 0$.

Independence for Continuous Random Variables

For random variables X and Y, if any one of the following properties is true, the others are also true. Then X and Y are independent.

(1) $f_{XY}(x, y) = f_X(x) \cdot f_Y(y)$ (2) $f_{Y|x}(y) = f_Y(y)$ for all x and y with $f_X(x) > 0$ (3) $f_{X|y}(y) = f_X(x)$ for all x and y with $f_Y(y) > 0$ (4) $P(X \subset A, Y \subset B) = P(X \subset A) \cdot P(Y \subset B)$ for any sets *A* and *B* in the range of *X* and *Y*, respectively. (5–7)

 $P(Y=y | X=x)=P(Y=y) \frac{\text{for any } x}{\text{for any } y} \text{ or}$ $P(X=x | Y=y)=P(X=x) \frac{\text{for any } y}{\text{for any } y} \text{ or}$ $P(X=x, Y=y)=P(X=x) \cdot P(Y=y) \frac{\text{for any } x \text{ and } y}{\text{for any } x \text{ and } y}$

Covariation, Correlations

Quick and dirty check for <u>linear</u> (in)dependence between variables

Covariance Defined

Covariance is a number quantifying the average *linear* dependence between two random variables.

The covariance between the random variables *X* and *Y*, denoted as cov(X, Y) or σ_{XY} is

 $\sigma_{XY} = E[(X - \mu_X)(Y - \mu_Y)] = E(XY) - \mu_X\mu_Y$ Montgomery, Runger 5th edition Eq. (5–14)

The units of σ_{XY} are the units of X times the units of Y.

Unlike the range of the variance, covariance can be negative: $-\infty < \sigma_{XY} < \infty$.

Covariance - 1 humber to measure dependance Between random variables Cov(X,Y) or Bxy $G_{XY} = E\left[\left(X - \mu_X\right), \left(Y - \mu_Y\right)\right] =$ = E(X.Y) - Mx - My• Var(X) = Cov(X,X)· If X& Y are independent $Cov(X,Y) = E[X-Mx] \cdot E[Y-My] = 0$ $-\infty < Cov(X,Y) < +\infty$ Can be negative.

Covariance and PMF tables

y = number of times city	x = number of bars of signal strength		
name is stated	1	2	3
1	0.01	0.02	0.25
2	0.02	0.03	0.20
3	0.02	0.10	0.05
4	0.15	0.10	0.05

The probability distribution of Example 5-1 is shown.

By inspection, note that the larger probabilities occur as X and Y move in opposite directions. This indicates a negative covariance.

Covariance and Scatter Patterns



Figure 5-13 Joint probability distributions and the sign of cov(X, Y). Note that covariance is a measure of linear relationship. Variables with non-zero covariance are correlated. Independence Implies $\sigma = \rho = 0$ but <u>not vice versa</u>

• If X and Y are independent random variables,

$$\sigma_{XY} = \rho_{XY} = 0 \tag{5-17}$$

• $\rho_{XY} = 0$ is necessary, but not a sufficient condition for independence.



Correlation is "normalized covariance"

Also called:
 Pearson correlation
 coefficient

 $\rho_{XY} = \sigma_{XY} / \sigma_X \sigma_Y$ is the covariance normalized to be $-1 \le \rho_{XY} \le 1$



Karl Pearson (1852–1936) English mathematician and biostatistician

Prove that f_{xy} is in [-1, 1] $Z_x = \frac{X - M_x}{6_x}; \quad Z_y = \frac{Y - M_y}{6_y}$ $0 \le E((2_x - 2_y)^2) = E(2_x^2) + E(2_y^2) -2E(2_{x}\cdot Z_{y})=2-2\frac{1}{b_{x}b_{y}}E((x-\mu_{x})(y-\mu_{1}))=$ $2 - 2 f_{XX} \longrightarrow f_{XY} \leq 2$ $0 \leq E((z_{x} + z_{y})^{2}) = E(z_{x}^{2}) + E(z_{y}^{2}) +$ $+2E(2_{x},2_{y}) = 2+2p_{xy} = 3$ $\implies \int x_{Y} \ge -1$

Spearman rank correlation

- Pearson correlation tests for linear relationship between X and Y
- Unlikely for variables with broad distributions → nonlinear effects dominate
- <u>Spearman correlation</u> tests for any <u>monotonic relationship</u> between X and Y
- Calculate ranks (1 to n), r_x(i) and r_y(i) of variables in both samples. Calculate Pearson correlation between ranks:
 Spearman(X,Y) = Pearson(r_x, r_y)
- Ties: convert to fractions, e.g. tie for 6s and 7s place both get 6.5. This can lead to artefacts.
- If lots of ties: use Kendall rank correlation (Kendall tau)



Example 3.10. Let an experiment consist of drawing a card at random from a standard deck of 52 playing cards. Define events *A* and *B* as "the card is a **\bigstar**" and "the card is a queen." Are the events *A* and *B* independent? By definition, $P(A \cdot B) = P(Q \spadesuit) = \frac{1}{52}$. This is the product of $P(\spadesuit) = \frac{13}{52}$ and $P(Q) = \frac{4}{52}$, and events *A* and *B* in question are independent. In this situation, intuition provides no help. Now, pretend that the 2 \heartsuit is drawn and excluded from the deck prior to the experiment. Events *A* and *B* become dependent since

$$\mathbb{P}(A) \cdot \mathbb{P}(B) = \frac{13}{51} \cdot \frac{4}{51} \neq \frac{1}{51} = \mathbb{P}(A \cdot B).$$

Example 1: Uniform distribution in the square -1 < X < 1, -1 < 4 < 1 $\begin{aligned} \int f(xy) = c \quad if \quad -1 < x < 1 \quad and \quad -1 < y < 1 \\ 0 \quad - \quad otherwise \end{aligned}$ $1 = \int dx \, dy \, f_{XY}(x, y) = (Area = (A - C) - \frac{1}{4})$ Square

Are X and Y independent? Yes they are Let's test if $-f_{XY}(x,y) = f_X(x) \cdot -f_Y(y)$ $f_{X}(z) = \int f_{XY}(x,y) dy =$ $f_{X}(z) = \int +x_{Y}(x,y)dy =$ $= \int \frac{1}{4}dy = \frac{1}{2}i_{f} - 1 < x < 1$ $\int \frac{1}{4}dy = \frac{1}{2}i_{f} - 1 < x < 1$ $\int \frac{1}{2}\int x$ Some for $f_{Y}(y) = \frac{1}{2}i_{f} - 1 < y < 1$

 $\frac{1}{4} = \int_{XY} (x, y) = \frac{1}{2} \cdot \frac{1}{2} = \int_{X} (x) \cdot f_{Y}(y)$ 0 otherwise if both x e y are in (-1,1)

X and Y are uniformly distributed in the disc x²+y²≤1



Are they independent?



C. I could not figure it out

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Joint PDF $\int xy(x,y) = \frac{1}{\alpha roa} = \frac{1}{\pi}$ if x,y in the disc Marginal distributions: D-otherwise $f_X(x) = \int \int xy(x,y) dy = \int \frac{dy}{\pi} = \frac{2\sqrt{1-x^2}}{\pi}$ Same for $f_Y(y) = \frac{2\sqrt{1-x^2}}{\pi}$ $\frac{1}{\pi} = f_{xy}(x,y) \neq \frac{2}{\pi} \int -x^2 \cdot \frac{2}{\pi} \int -y^2 = f_x(y) \cdot f_y(y)$ Variables are NOT independent



Matlab exercise: Correlation/Covariation

- Generate a sample with Stats=100,000 of two Gaussian random variables r1 and r2 which have mean 0 and standard deviation 2 and are:
 - Uncorrelated
 - Correlated with correlation coefficient 0.9
 - Correlated with correlation coefficient -0.5
 - Trick: first make uncorrelated r1 and r2. Then make anew variable: r1mix=mix.*r2+(1-mix.^2)^0.5.*r1; where mix= corr. coeff.
- For each value of mix calculate covariance and correlation coefficient between r1mix and r2
- In each case make a scatter plot: plot(r1mix,r2,'k.');

Matlab exercise: Correlation/Covariation

- 1. Stats=100000;
- 2. r1=2.*randn(Stats,1);
- 3. r2=2.*randn(Stats,1);
- 4. disp('Covariance matrix='); disp(cov(r1,r2));
- 5. disp('Correlation=');disp(corr(r1,r2));
- 6. figure; plot(r1,r2,'k.');
- 7. mix=0.9; %Mixes r2 to r1 but keeps same variance
- 8. r1mix=mix.*r2+sqrt(1-mix.^2).*r1;
- 9. disp('Covariance matrix='); disp(cov(r1mix,r2));
- 10.disp('Correlation=');disp(corr(r1mix,r2));
- 11.figure; plot(r1mix,r2,'k.');
- 12.mix=-0.5; %REDO LINES 8-11



