

1. (a) No #

Consider the following example:

$X, Y_1,$  and  $Y_2$  are 3 RVS on  $\{0,1\}$  with joint pmf

$$\begin{cases} P[(X, Y_1, Y_2) = (1, 0, 0)] = \frac{1}{4} \\ P[(X, Y_1, Y_2) = (0, 1, 0)] = \frac{1}{4} \\ P[(X, Y_1, Y_2) = (0, 0, 1)] = \frac{1}{4} \\ P[(X, Y_1, Y_2) = (1, 1, 1)] = \frac{1}{4} \end{cases}$$

Or equivalently,  $(X, Y_1, Y_2) \sim \text{Unif}(\{(1,0,0), (0,1,0), (0,0,1), (1,1,1)\})$

It immediately follows that  $H(X, Y_1, Y_2) = \log_2 4 = 2$ .

Note that  $X, Y_1,$  and  $Y_2$  have a marginal distribution  $\text{Ber}(\frac{1}{2})$ , and thus  $H(X) = H(Y_1) = H(Y_2) = 1$ .

In addition, note that  $(X, Y_1), (X, Y_2), (Y_1, Y_2)$  have the same distribution  $\text{Unif}(\{(0,0), (0,1), (1,0), (1,1)\})$ .

Thus  $H(X, Y_1) = H(X, Y_2) = H(Y_1, Y_2) = \log_2 4 = 2$ .

Now we have

$$I(X; Y_1) = H(X) + H(Y_1) - H(X, Y_1) = 1 + 1 - 2 = 0,$$

$$I(X; Y_2) = H(X) + H(Y_2) - H(X, Y_2) = 1 + 1 - 2 = 0, \text{ but}$$

$$\begin{aligned} I(X; Y_1, Y_2) &= H(X) + H(Y_1, Y_2) - H(X, Y_1, Y_2) \\ &= 1 + 2 - 2 = 1 \neq 0, \end{aligned}$$

2. No #

We can take arbitrary iid  $X$  and  $Y_1$ , say  $X, Y_1 \stackrel{\text{iid}}{\sim} \text{Ber}(\frac{1}{2})$ ,  
and then set  $Y_2 = Y_1$ .

Then  $I(X; Y_1) = I(X; Y_2) = 0$  since  $X$  and  $Y_i$  are independent  
and  $Y_2 = Y_1$ .

$$\begin{aligned} \text{But } I(Y_1, Y_2) &= I(Y_1, Y_1) = H(Y_1) - H(Y_1|Y_1) \\ &= H(Y_1) = 1 \neq 0. \end{aligned}$$

2-1

2-(a) We calculate  $H(X, Z|Y)$  in two different ways:

① By chain rule of conditional entropy, we have

$$H(X, Z|Y) = H(Z|Y) + H(X|Y, Z) \dots (2-1)$$

② Since  $X \rightarrow Y \rightarrow Z$ , we have  $X$  and  $Z$  are conditionally independent given  $Y$ . Thus,

$$\begin{aligned} H(X, Z|Y) &= \sum_Y P_Y(y) H(X, Z|Y=y) \\ &= \sum_Y P_Y(y) (H(X|Y=y) + H(Z|Y=y)) \\ &= H(X|Y) + H(Z|Y), \dots (2-2) \end{aligned}$$

Comparing (2-1) and (2-2), we have  $H(X|Y) = H(X|Y, Z)$ .

Q.E.D.

(b) From (a) and the fact that conditioning reduces entropy, we have

$$H(X|Y) = H(X|Y, Z) \leq H(X|Z). \quad \text{Q.E.D.}$$

(c) From (b), we have

$$\begin{aligned} I(X; Y) - I(X; Z) &= (H(X) - H(X|Y)) - (H(X) - H(X|Z)) \\ &= H(X|Z) - H(X|Y) \geq 0. \end{aligned}$$

Thus  $I(X; Y) \geq I(X; Z)$ .

QED.

(d) From (a), we have

$$I(X; Z|Y) = H(X|Y) - H(X|Y, Z) = 0. \quad \text{QED.}$$

3. Following the terminology in class, we seek to prove that for any given collection of RVs  $X_1, \dots, X_n$ , the following set function  $g$  is submodular:

$$g: 2^N \rightarrow \mathbb{R}_{\geq 0}$$

$$g(\{i_1, i_2, \dots, i_\ell\}) = H(X_{i_1}, X_{i_2}, \dots, X_{i_\ell}).$$

More precisely, we seek to show for any  $S_1$  and  $S_2$  such that

$$S_1 = \{i_1, \dots, i_\ell\} \subseteq S_2 = \{i_1, \dots, i_\ell, i_{\ell+1}, \dots, i_m\} \\ \subseteq \{1, \dots, n\}$$

and for any  $j \in \{1, \dots, n\} \setminus S_2$ , we have

$$g(S_1 \cup \{j\}) - g(S_1) \geq g(S_2 \cup \{j\}) - g(S_2),$$

or equivalently

$$H(X_{i_1}, \dots, X_{i_\ell}, X_j) - H(X_{i_1}, \dots, X_{i_\ell})$$

$$\geq H(X_{i_1}, \dots, X_{i_m}, X_j) - H(X_{i_1}, \dots, X_{i_m}).$$

... (3-1)

3-2

Note that by the chain rule of entropy, we have

$$\begin{aligned} H(X_{i_1}, \dots, X_{i_\ell}, X_j) &= H(X_{i_1}, \dots, X_{i_\ell}) \\ &+ H(X_j | X_{i_1}, \dots, X_{i_\ell}). \quad \dots (3-2) \end{aligned}$$

And similarly,

$$\begin{aligned} H(X_{i_1}, \dots, X_{i_m}, X_j) &= H(X_{i_1}, \dots, X_{i_m}) \\ &+ H(X_j | X_{i_1}, \dots, X_{i_m}). \quad \dots (3-3). \end{aligned}$$

Furthermore, since conditioning reduces entropy, we have

$$\begin{aligned} H(X_j | X_{i_1}, \dots, X_{i_m}) \\ &= H(X_j | X_{i_1}, \dots, X_{i_\ell}, X_{i_{\ell+1}}, \dots, X_{i_m}) \\ &\leq H(X_j | X_{i_1}, \dots, X_{i_\ell}). \quad \dots (3-4) \end{aligned}$$

Combining (3-2), (3-3) and (3-4) proves (3-1). QED

4. (c) I write  $\mathbb{1}_{\{x\}}$  for the indicator function instead to avoid confusion with mutual information.

Following the hint, define  $P_n'$  by

$$\hat{P}_n'(x) = \frac{1}{n} \sum_{j=n+1}^{2n} \mathbb{1}_{\{X_j = x\}} \quad \text{for each } x \in \mathcal{X}.$$

Then we have for each  $x \in \mathcal{X}$  that

$$\begin{aligned} \hat{P}_{2n}(x) &= \frac{1}{2n} \sum_{i=1}^{2n} \mathbb{1}_{\{X_i = x\}} \\ &= \frac{1}{2n} \left( \sum_{i=1}^n \mathbb{1}_{\{X_i = x\}} + \sum_{j=n+1}^{2n} \mathbb{1}_{\{X_j = x\}} \right) \\ &= \frac{1}{2} \hat{P}_n + \frac{1}{2} \hat{P}_n'. \end{aligned}$$

Thus, by the convexity of KL divergence, we have

$$\begin{aligned} D(\hat{P}_{2n} \| P) &= D\left(\frac{1}{2} \hat{P}_n + \frac{1}{2} \hat{P}_n' \| P\right) \\ &\leq \frac{1}{2} D(\hat{P}_n \| P) + \frac{1}{2} D(\hat{P}_n' \| P). \end{aligned}$$

... (4-1)

4-2

Taking expectation on both sides of (4-1), we have

$$\mathbb{E}[D(\hat{P}_n \| P)] \leq \frac{1}{2} \mathbb{E}[D(\hat{P}_n \| P)] + \frac{1}{2} \mathbb{E}[D(\hat{P}'_n \| P)]. \quad \dots (4-2)$$

Finally, note that since  $(X_i)_{i=1}^n$  are iid,  $\hat{P}'_n$  has the same distribution as  $\hat{P}_n$ , and thus (4-2) becomes

$$\mathbb{E}[D(\hat{P}_n \| P)] \leq \mathbb{E}[D(\hat{P}_n \| P)]. \quad \text{QED.}$$

(b) Following the hint, define for each  $k \in \{1, \dots, n\}$  and for each  $x \in \mathcal{X}$  that

$$\hat{P}_n^{(k)}(x) = \frac{1}{n-1} \sum_{\substack{i=1 \\ i \neq k}}^n \mathbb{1}_{\{X_i = x\}}.$$

Then, we have for each  $x \in \mathcal{X}$  that

$$\begin{aligned} \sum_{k=1}^n \hat{P}_n^{(k)}(x) &= \frac{1}{n-1} \sum_{k=1}^n \sum_{\substack{i=1 \\ i \neq k}}^n \mathbb{1}_{\{X_i = x\}} \\ &\stackrel{\text{(switch summation)}}{=} \frac{1}{n-1} \sum_{i=1}^n \sum_{\substack{k=1 \\ k \neq i}}^n \mathbb{1}_{\{X_i = x\}} \end{aligned}$$



4-3

$$\begin{aligned}
 &= \frac{1}{n-1} \sum_{i=1}^n (n-1) \mathbb{1}_{\{X_i = x\}} \\
 &= n \hat{P}_n(x).
 \end{aligned}$$

That is, we have

$$\hat{P}_n = \sum_{k=1}^n \frac{1}{n} \hat{P}_n^{(k)}.$$

Then, by convexity of  $D(\cdot \| P)$  again, we have

$$\begin{aligned}
 D(\hat{P}_n \| P) &= D\left(\sum_{k=1}^n \frac{1}{n} \hat{P}_n^{(k)} \| P\right) \\
 &\leq \sum_{k=1}^n \frac{1}{n} D(\hat{P}_n^{(k)} \| P) \\
 &= \frac{1}{n} \sum_{k=1}^n D(\hat{P}_n^{(k)} \| P). \dots (4-3).
 \end{aligned}$$

Finally, by the iid property of  $(X_i)_{i=1}^n$  again, all the  $\hat{P}_n^{(k)}$  have the same distribution as  $\hat{P}_n^{(n)} = \hat{P}_{n-1}$ , and thus taking expectation on both sides of (4-3) yields

$$\begin{aligned}
 \mathbb{E}[D(\hat{P}_n \| P)] &\leq \frac{1}{n} \sum_{k=1}^n \mathbb{E}[D(\hat{P}_n^{(k)} \| P)] \\
 &= \mathbb{E}[D(\hat{P}_{n-1} \| P)]. \quad \text{QED.}
 \end{aligned}$$

5-1

5. (a)  $H(X) = -0.6 \log 0.6 - 0.4 \log 0.4 \approx \underline{0.9710} \#$

(b) A sequence  $x^n$  falls in  $A_{\epsilon}^{(n)}$  if and only if

$$H(X) - \epsilon \leq -\frac{1}{n} \log P(x^n) \leq H(X) + \epsilon.$$

Putting  $H(X) = 0.9710$  and  $\epsilon = 0.1$  gives

$$0.8710 \leq -\frac{1}{n} \log P(x^n) \leq 1.0710. \dots (5-1)$$

According to the table, the sequences with  $k=11, \dots, 19$  ones satisfy (5-1).

Thus,  $A_{0.1}^{(0.8)}$  contains the sequences with  $k=11, \dots, 19$

ones. #

The probability of this set can be found by

$$\sum_{k=11}^{19} \binom{n}{k} p^k (1-p)^{n-k}, \text{ where } p=0.6 \text{ and } n=25.$$

By table, this value is approximately 0.9362 #.

The cardinality of this set can be found by

$$\sum_{k=11}^{19} \binom{n}{k}, \text{ which is } \underline{26366510} \#.$$

(c) Note that  $P(X=1) = 0.6 > P(X=0) = 0.4$ .

Therefore, this smallest set with prob. 0.9, denoted as B, should contain as many 1's as possible. More explicitly, B should contain all sequences with 25 1's, and then 24 1's, and so on, until the probability of B is no less than 0.9.

That is, one seek to find  $k_0$  s.t.

$$\sum_{k=k_0}^{25} \binom{25}{k} p^k (1-p)^{n-k} < 0.9 \text{ but}$$

$$\sum_{k=k_0-1}^{25} \binom{25}{k} p^k (1-p)^{n-k} \geq 0.9.$$

A numerical computation shows  $k_0 = 13$ , and thus

B contains all the sequences with 13, 14, ..., 25 1's.

The probability of these sequences is

5-3

$$\sum_{k=13}^{25} \binom{25}{k} p^k (1-p)^{n-k} \approx 0,8462.$$

The rest probability  $p_0 = 0,9 - 0,8462 = 0,0538$  can be fulfilled by collecting  $n_0 = \left\lceil \frac{0,0538}{p^{12}(1-p)^{13}} \right\rceil \approx 3680693$

sequences with 12 1's into B.

Therefore, the number of elements in B is

$$\sum_{k=13}^{25} \binom{25}{k} + n_0 \approx \underline{20457889} \#.$$

(d) The intersection contains all the sequences of 13, ..., 19 1's and  $n_0$  sequences of 12 1's.

Thus there are

$$\sum_{k=13}^{19} \binom{25}{k} + n_0 \approx \underline{20389483} \# \text{ elements in}$$

the intersection.

5-4

The probability is

$$\sum_{k=13}^{19} \binom{25}{k} p^k (1-p)^{25-k} + p_0 \approx \underline{0.8706} \#.$$

6-1

6. Let  $t_n := \frac{x_n}{y_n}$ .

By assumption,  $\sum_{n=1}^{\infty} t_n$  converges, and denote this limit to be  $A \in \mathbb{R}$ .

Now we seek to prove

$$\frac{1}{y_n} \sum_{i=1}^n x_i = \frac{1}{y_n} \sum_{i=1}^n y_i t_i \rightarrow 0 \text{ as } n \rightarrow \infty. \quad \dots (6-1).$$

Define  $S_n = \begin{cases} \sum_{i=1}^n t_i, & n \geq 1 \\ 0, & n = 0 \end{cases}$ , and then

$$t_n = S_n - S_{n-1} \text{ for } n \geq 1.$$

Now, by summation by part, we have

$$\begin{aligned} \frac{1}{y_n} \sum_{i=1}^n y_i t_i &= \frac{1}{y_n} \sum_{i=1}^n y_i (S_i - S_{i-1}) \\ &= \frac{1}{y_n} \left( \sum_{i=1}^n y_i S_i - \sum_{i=1}^n y_i S_{i-1} \right) \\ &= \frac{1}{y_n} \left( y_n S_n + \sum_{i=1}^{n-1} y_i S_i - \sum_{i=2}^n y_i S_{i-1} \right) \end{aligned}$$

6-2

$$= \frac{1}{y_n} \left( y_n S_n + \sum_{i=1}^{n-1} y_i S_i - \sum_{i=1}^{n-1} y_{i+1} S_i \right)$$

$$= S_n + \frac{1}{y_n} \sum_{i=1}^{n-1} S_i (y_i - y_{i+1})$$

$$= S_n - \frac{1}{y_n} \sum_{i=1}^{n-1} S_i (y_{i+1} - y_i) \quad \dots (6-2)$$

Taking absolute value on (6-2), by triangle inequality we have

$$\left| \frac{1}{y_n} \sum_{i=1}^n t_i y_i \right| = \left| S_n - \frac{1}{y_n} \sum_{i=1}^{n-1} S_i (y_{i+1} - y_i) \right|$$

$$\leq |S_n - A| + \left| \frac{1}{y_n} \sum_{i=1}^{n-1} S_i (y_{i+1} - y_i) - A \right|$$

$$= |S_n - A| + \left| \frac{1}{y_n} \sum_{i=1}^{n-1} S_i (y_{i+1} - y_i) - \frac{1}{y_n} \sum_{i=1}^{n-1} A (y_{i+1} - y_i) \right|$$

$$= |S_n - A| + \left| \frac{1}{y_n} \sum_{i=1}^{n-1} (S_i - A) (y_{i+1} - y_i) \right|$$

$$\leq |S_n - A| + \frac{1}{|y_n|} \sum_{i=1}^{n-1} |S_i - A| |y_{i+1} - y_i|$$

6-3

$$= |S_n - A| + \frac{1}{|y_n|} \sum_{i=1}^{n-1} |S_i - A| (y_{i+1} - y_i) \quad (6-3)$$

where the last equality holds since  $(y_i)_{i=1}^n$  is increasing.

Now let  $\epsilon > 0$ .

Since  $S_n = \sum_{i=1}^n t_i \xrightarrow{n \rightarrow \infty} A$  by assumption,  $\exists N_1 \in \mathbb{N}$  s.t.

$\forall n \geq N_1$ , we have  $|S_n - A| < \frac{\epsilon}{3}$

In addition, since  $y_n \xrightarrow{n \rightarrow \infty} \infty$  by assumption,  $\exists N_2 \in \mathbb{N}$

s.t.  $\forall n \geq N_2$ , we have  $y_n > 0$  and

$$\frac{1}{y_n} \sum_{i=1}^{n-1} |S_i - A| (y_{i+1} - y_i) < \frac{\epsilon}{3}$$

Now for  $n \geq N_3 := \max(N_1, N_2)$ , we have from (6-3) that

$$\left| \frac{1}{y_n} \sum_{i=1}^n t_i y_i \right| \leq \frac{\epsilon}{3} + \frac{1}{y_n} \sum_{i=1}^{n-1} |S_i - A| (y_{i+1} - y_i)$$

$$+ \frac{1}{y_n} \sum_{i=N_1}^{n-1} |S_i - A| (y_{i+1} - y_i)$$

$$\leq \frac{\epsilon}{3} + \frac{\epsilon}{3} + \frac{1}{y_n} \sum_{i=N_1}^{n-1} \frac{\epsilon}{3} (y_{i+1} - y_i)$$



6-4

$$= \frac{2}{3} \epsilon + \frac{\epsilon}{3} \left(1 - \frac{y_{N_1}}{y_n}\right)$$

$$\leq \frac{2}{3} \epsilon + \frac{\epsilon}{3} = \epsilon. \quad \dots (6-4)$$

Since  $\epsilon$  is arbitrary, from (6-4) we deduce that

$$\left| \frac{1}{y_n} \sum_{i=1}^n x_i \right| = \left| \frac{1}{y_n} \sum_{i=1}^n t_i y_i \right| \xrightarrow{n \rightarrow \infty} 0, \text{ and}$$

$$\text{thus } \frac{1}{y_n} \sum_{i=1}^n x_i \xrightarrow{n \rightarrow \infty} 0.$$

QED

17. (a) Define

$$(P_1, \dots, P_7) = (0.05, 0.08, 0.13, 0.09, 0.30, 0.20, 0.15)$$

$$\text{Then } H(X) = - \sum_{i=1}^7 P_i \log P_i \approx \underline{2.5989} \#$$

(b) If there is no probability known, encoding 7 symbols needs

A minimum of  $\lceil \log_2 7 \rceil = 3$  bits.

A fixed-length code can be constructed as follows:

A  $\rightarrow$  000

B  $\rightarrow$  001

C  $\rightarrow$  010

D  $\rightarrow$  011

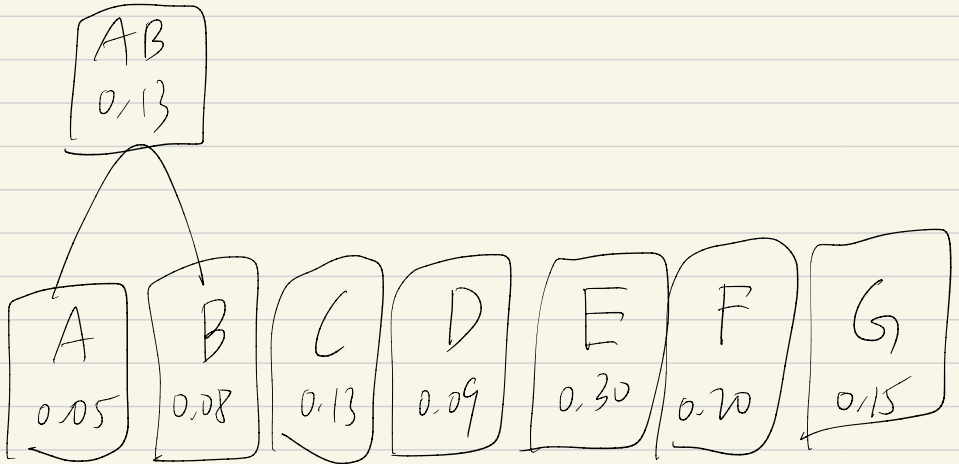
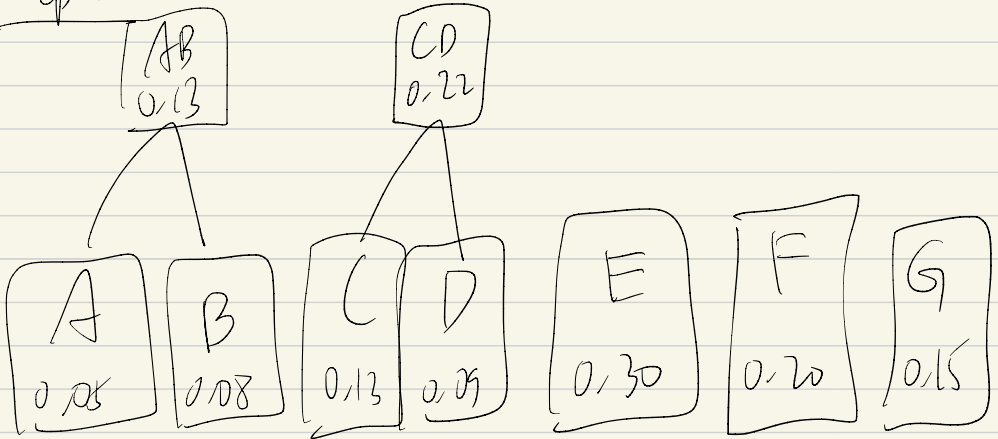
E  $\rightarrow$  100

F  $\rightarrow$  101

G  $\rightarrow$  110

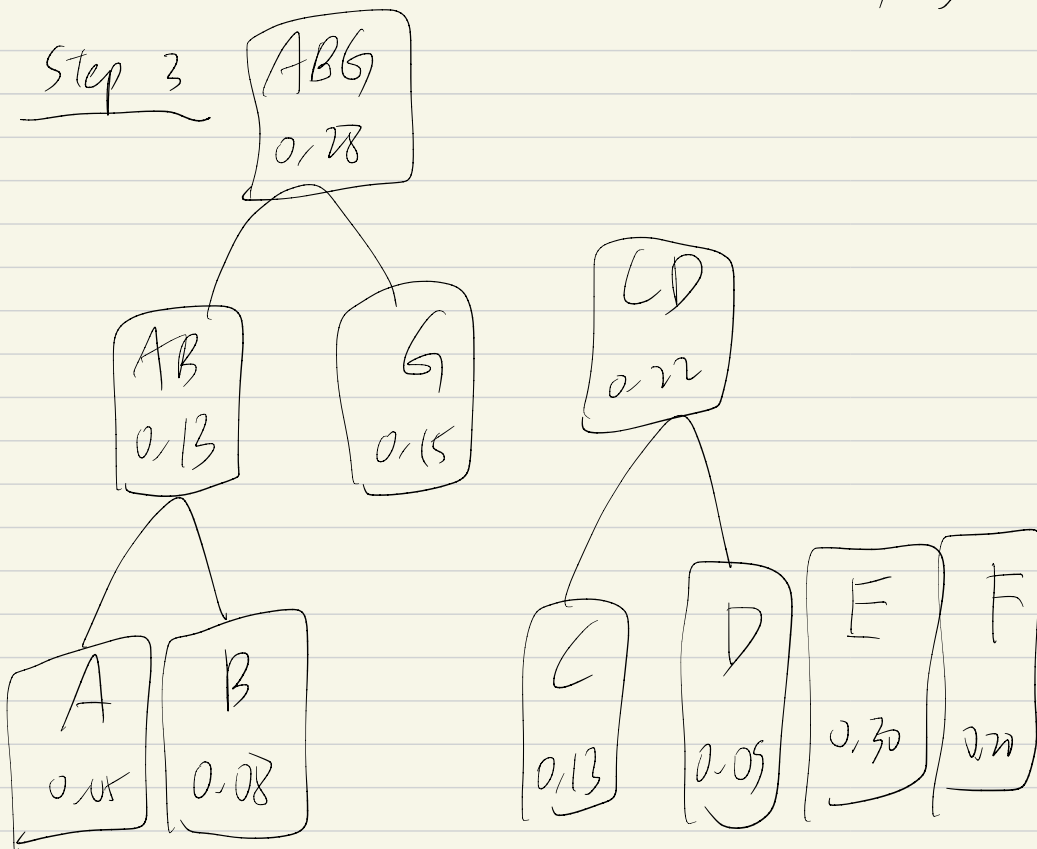
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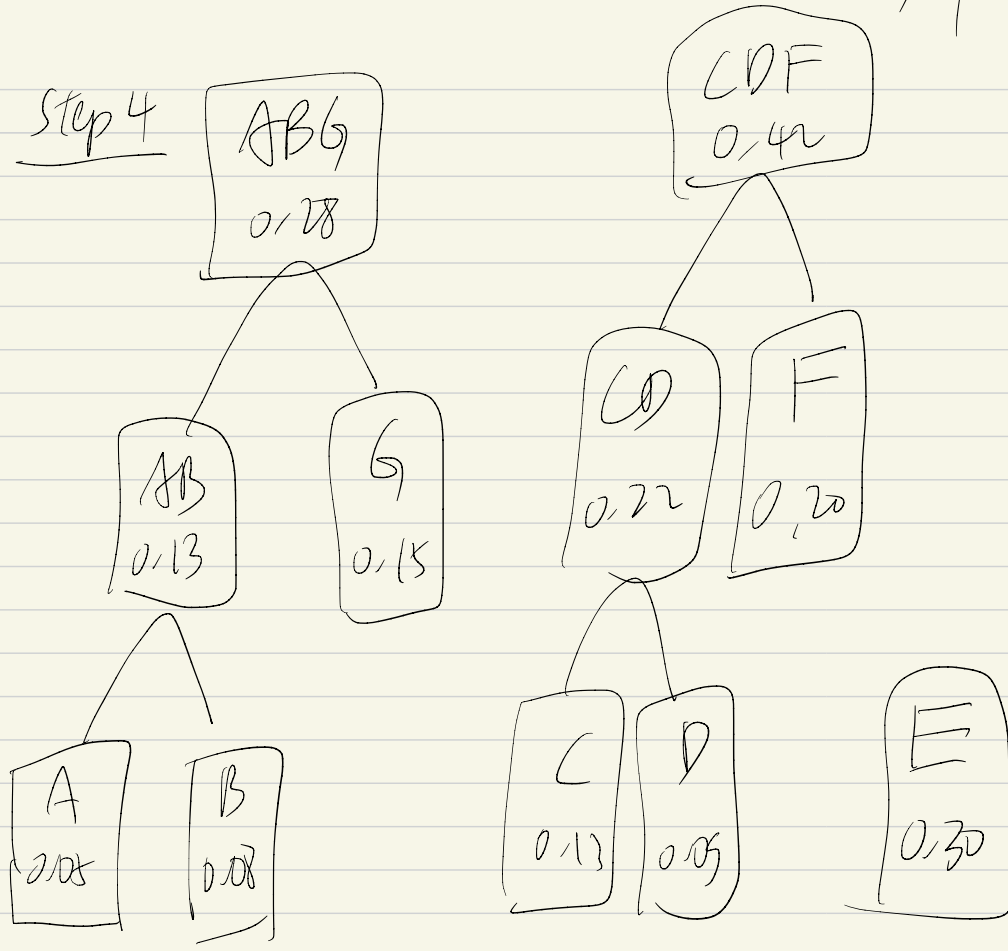
(c) Step 1Step 2

7-3

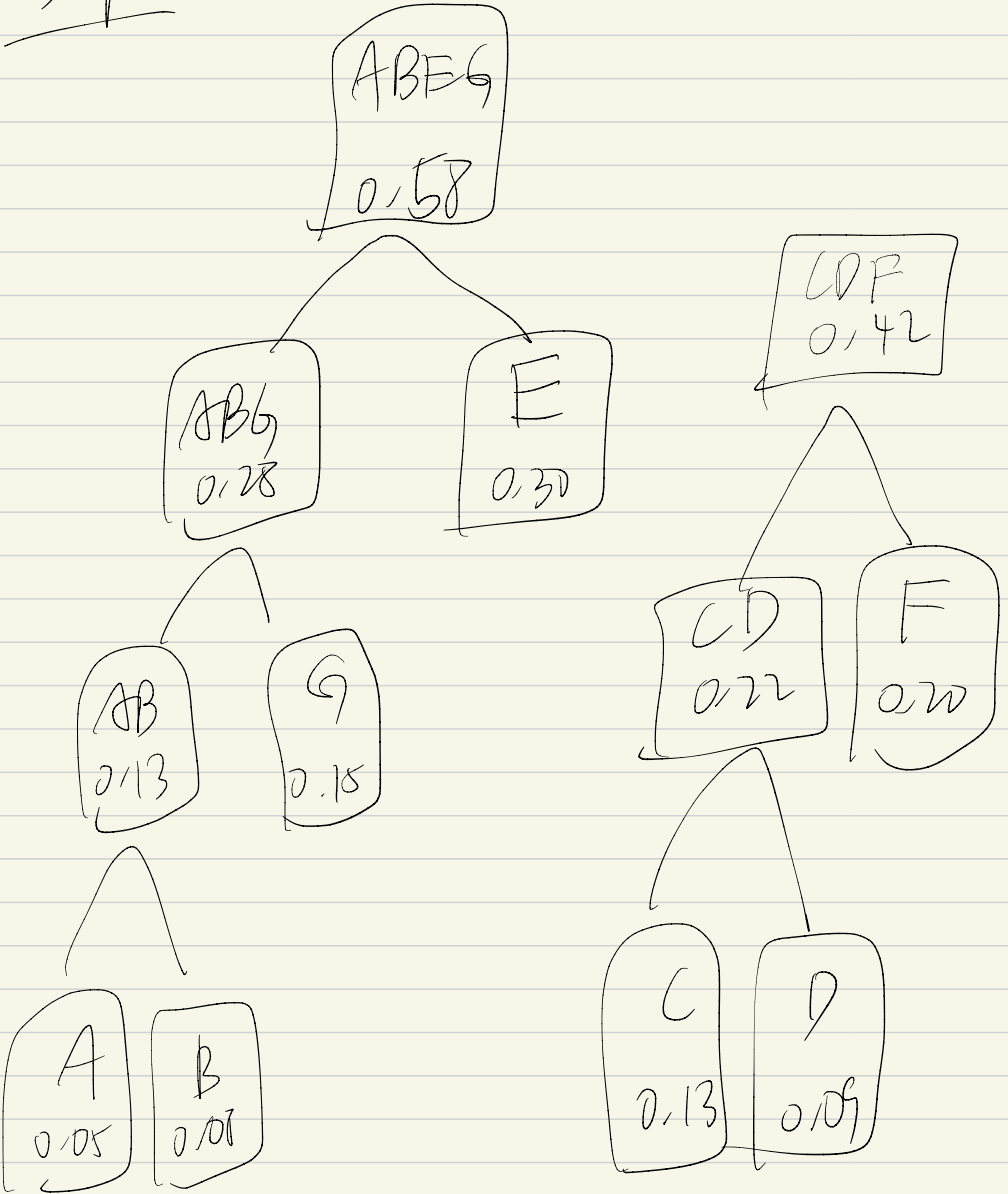
Step 3



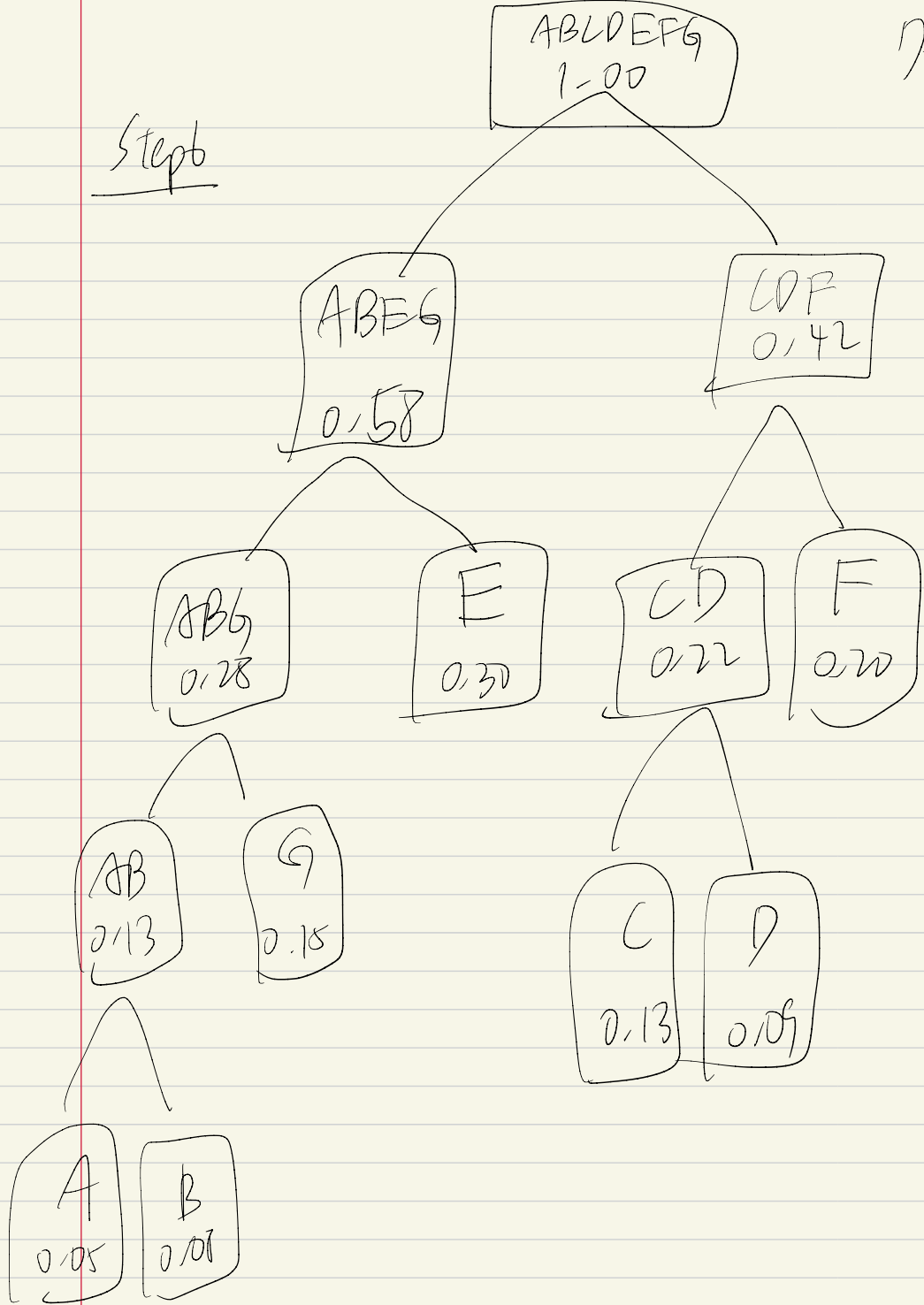
Step 4



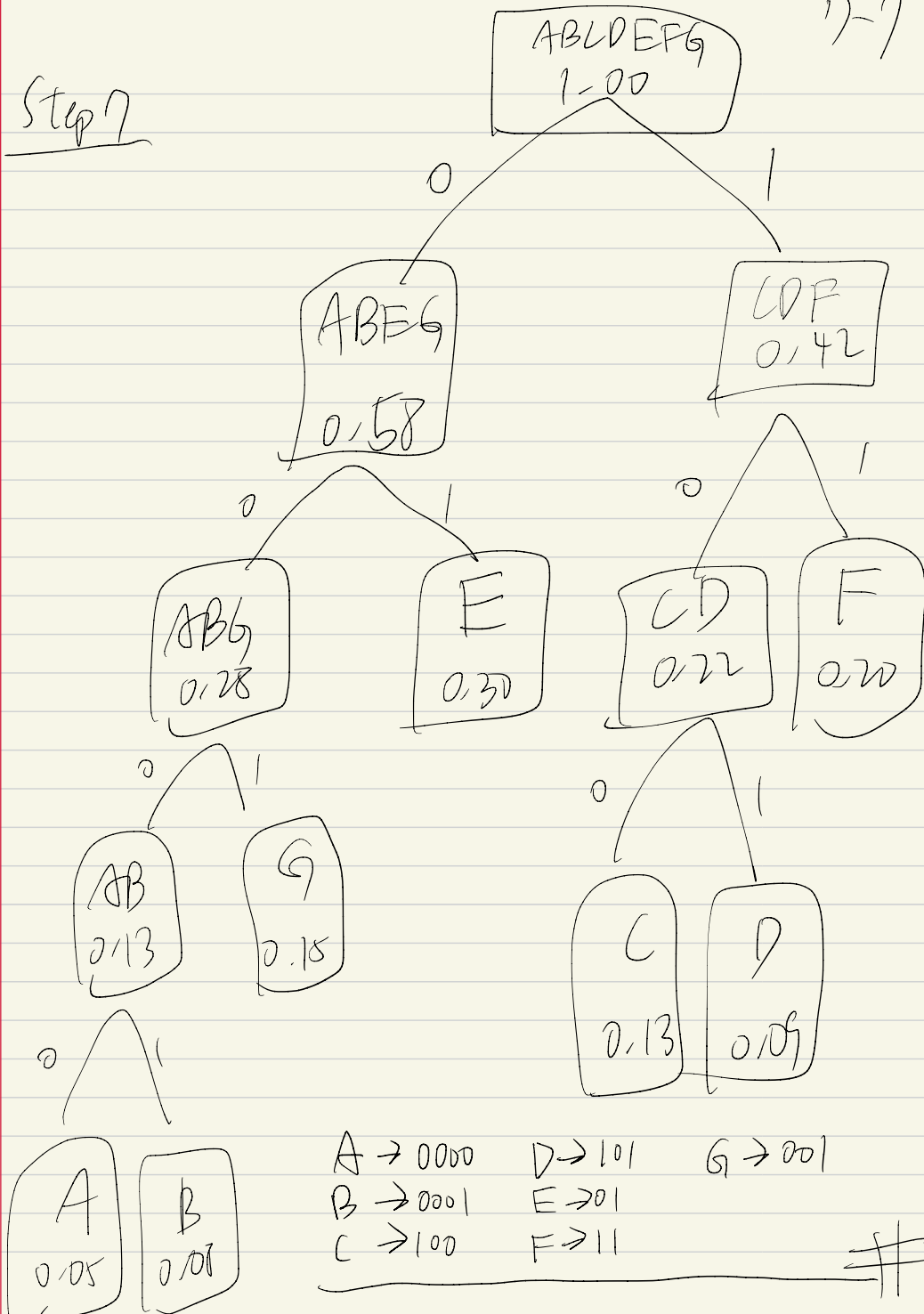
Step 5



Steps



Step 7





We have  $(l_1, l_2, l_3, l_4, l_5, l_6, l_7) = (4, 4, 3, 3, 2, 2, 3)$ .

Thus, the average code length is

$$\sum_{i=1}^7 p_i l_i = \underline{2.63} \#$$

(d) We compute

$$\sum_{i=1}^7 2^{-l_i} = 1 \leq 1, \text{ and thus Kraft's inequality}$$

is satisfied #

(e) The minimum code length satisfies

$$H(x_1, \dots, x_{10}) \leq \lceil 10L \rceil < H(x_1, \dots, x_{10}) + 1, \text{ where}$$

$$x_i \stackrel{\text{i.i.d.}}{\sim} X.$$

$$\text{Thus } \lceil 10H(X) \rceil \leq 10L < \lceil 10H(X) \rceil + 1$$

$$\Rightarrow \underline{H(X) \leq L < H(X) + \frac{1}{10}} \#$$

# I ILLINOIS ECE

7. (P)(1)

sample average  $\mu$ :  $\frac{1}{n} \sum_{i=1}^n \log_2 p(x_i)$  → using source statistics

$$= \frac{1}{8} (3 \log_2 0.3 + 2 \log_2 0.2 + 1 \log_2 0.05 + \log_2 0.08 + \log_2 0.13)$$

$$= -2.5955 \quad \bar{X}_n$$

(this gives us a running sample average of 2.5955 bits)

2) 2.5984 bits

3) using problem 3.1 in T&C (and Chelyshev)

$$P\{|\bar{X}_n - \mu| < \epsilon\} \leq \frac{\sigma^2}{n\epsilon^2} \quad \begin{array}{l} \epsilon = 1 \\ n = 8 \quad \sigma^2 = 2 \end{array}$$

$$P\{|\bar{X}_n - 2.5984| < 1\} \geq \frac{0.75}{2} = 0.375$$

$$P\{|\bar{X}_n - 2.5984| \geq \boxed{0.75}\}$$

# I ILLINOIS ECE

$$4) P\{\sum |\bar{X}_n - \mu| < \epsilon\} \geq 1 - \frac{\sigma^2}{n\epsilon}$$

$$1 - \frac{\sigma^2}{n\epsilon} \geq 0.99$$

$$n = \left\lceil \frac{\sigma^2}{0.01\epsilon} \right\rceil = \boxed{200}$$

~~or~~

$$5) n = \left\lceil \frac{100 \cdot 0.25}{1^2} \right\rceil = 25$$

$$\text{possible \# tuples} = 7^{25}$$

6) as  $n$  gets large

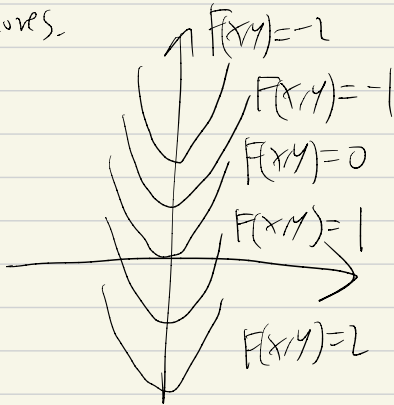
$$p_0 = \frac{2.5989^n}{7^n} \approx \frac{6.054^n}{7^n}$$

$$7.) \left\lceil \log_2 2^{2.5989n} \right\rceil = \left\lceil 2.5989n \right\rceil$$

per tuple

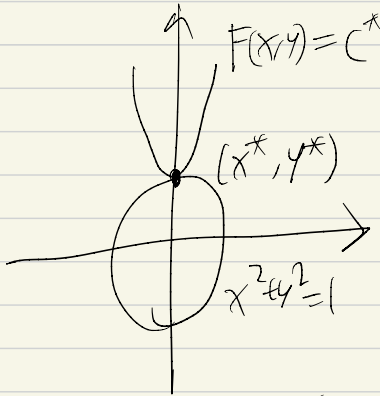
Max

8. For each  $c \in \mathbb{R}$ , the equation  $F(x,y) = 8x^2 - 4y = c$  defines a curve on  $\mathbb{R}^2$ . As  $c$  slightly changes, the curve slightly moves.



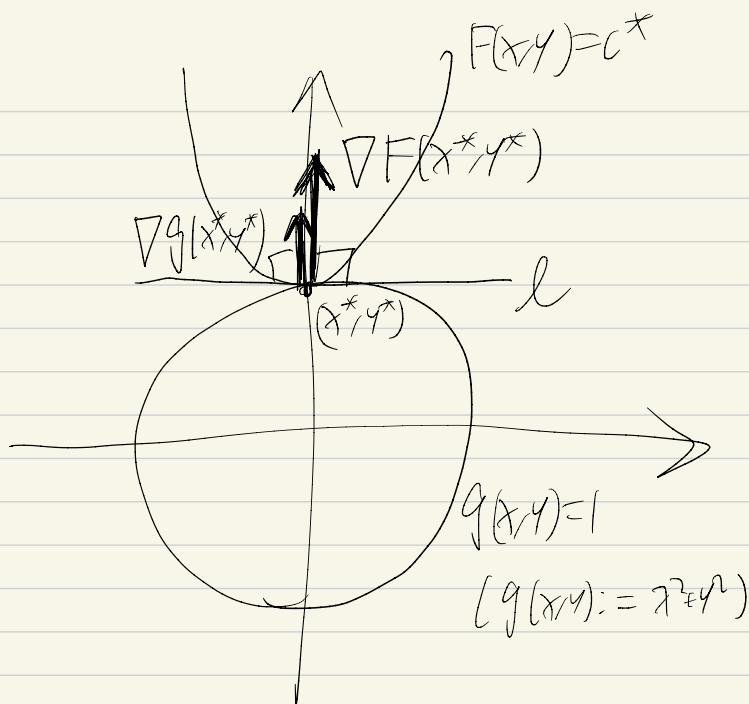
In this example, as  $c$  decreases, the curve moves upward. To minimize  $F(x,y)$  under the constraint  $x^2 + y^2 = 1$ , one seeks to move the curve upward as much as possible while keeping at least an intersection of the unit circle.

8-2



For the minimal  $C^*$ , it is the case that moving the curve upward any further will make the curve have no intersection of  $x^2 + y^2 = 1$ . Graphically, we know that in this case, the curve  $F(x,y) = C^*$  is tangent to  $x^2 + y^2 = 1$  at the intersecting point  $(x^*, y^*)$ .

8-3



Graphically, the normal vector of  $F(x, y) = c^*$  at  $(x^*, y^*)$ , i.e.  $\nabla F(x^*, y^*)$ , is perpendicular to the tangent line  $l$ .

At the same time, the normal vector of  $g(x, y) = 1$  at  $(x^*, y^*)$ , i.e.  $\nabla g(x^*, y^*)$ , is also perpendicular to  $l$ .

Since  $\nabla F(x^*, y^*)$  and  $\nabla g(x^*, y^*)$  are perpendicular to the same line, they are parallel  $\#$ .

That is,  $\exists -\lambda \in \mathbb{R}$  s.t.  $\nabla F(x^*, y^*) = -\lambda \nabla g(x^*, y^*)$ , here the role of  $-\lambda$  is the scaling between two parallel vectors  $\nabla F(x^*, y^*)$  and  $\nabla g(x^*, y^*)$ .

$\#$

8-4

Now come back to the Lagrange multiplier equations:

$$\mathcal{L}(x, y, \lambda) = F(x, y) + \lambda (g(x, y) - 1).$$

$$\Rightarrow \left\{ \begin{array}{l} \frac{\partial \mathcal{L}}{\partial x} = \frac{\partial F}{\partial x} + \lambda \frac{\partial g}{\partial x} = 0 \quad \dots (8-1) \\ \frac{\partial \mathcal{L}}{\partial y} = \frac{\partial F}{\partial y} + \lambda \frac{\partial g}{\partial y} = 0 \quad \dots (8-2) \\ \frac{\partial \mathcal{L}}{\partial \lambda} = g(x, y) - 1 = 0 \quad \dots (8-3) \end{array} \right.$$

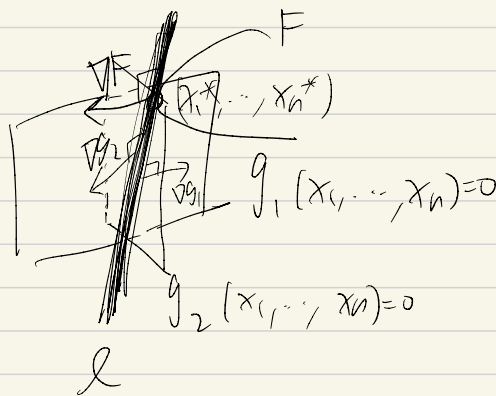
(8-1) & (8-2) mean exactly that  $\nabla F(x^*, y^*) = -\lambda \nabla g(x^*, y^*)$ ,  
and (8-3) is simply the constraint.

As a remark, having parallel gradient is just a necessary condition for being a minimal solution. That's why one have to calculate all the solutions satisfying (8-1), (8-2), and (8-3) and find the minimum among these solutions.

(say  $m$ ) 8-5

In the case where there are multiple<sup>v</sup> constraints, graphically,  $\nabla F(x_1, \dots, x_n)$  is tangent to the intersection of  $g_1(x_1, \dots, x_n) = 0$ ,  $g_2(x_1, \dots, x_n) = 0$ , ..., and  $g_m(x_1, \dots, x_n) = 0$ , and thus

$$\nabla F(x_1, \dots, x_n) \in \text{span}(\nabla g_1(x_1, \dots, x_n), \dots, \nabla g_m(x_1, \dots, x_n)).$$



$$\begin{cases} \ell \perp \nabla g_1(x_1^*, \dots, x_n^*) \\ \ell \perp \nabla g_2(x_1^*, \dots, x_n^*) \\ \ell \perp \nabla F(x_1^*, \dots, x_n^*) \end{cases}$$

That is,  $\exists -\lambda_1, \dots, -\lambda_m \in \mathbb{R}$  s.t.

$\nabla F(x_1^*, \dots, x_n^*) = \sum_{i=1}^m -\lambda_i \nabla g_i(x_1^*, \dots, x_n^*)$ , which explains the Lagrange multiplier with multiple constraints.