

# Capacity of Noisy Permutation Channels

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12/11/2024

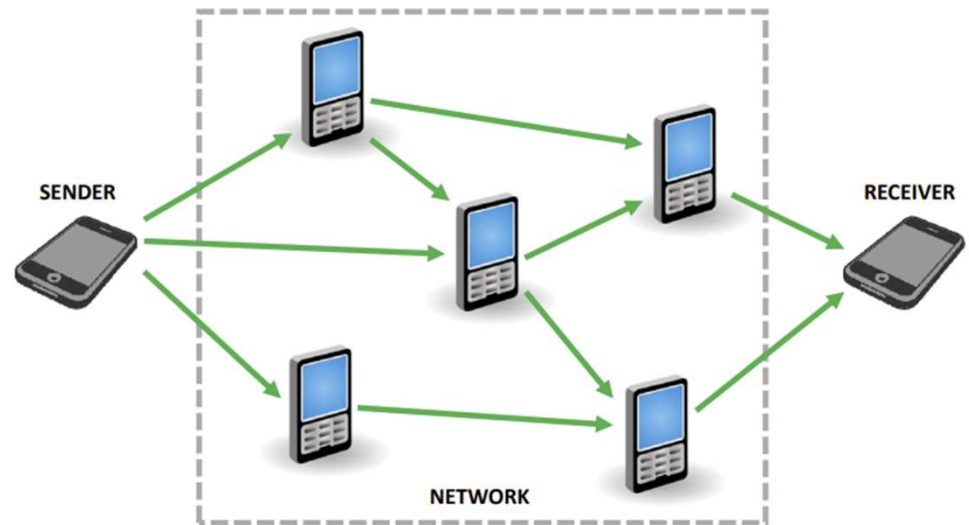
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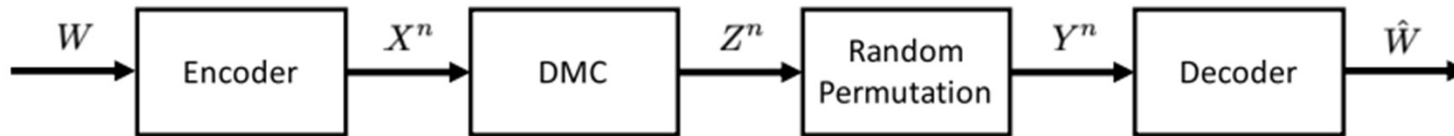
# 1. Motivation and Problem Statement

- Some channels do not preserve order (abc  $\rightarrow$  cba)
  - DNA storage
  - Packets in a network

Note: while mentioned in literature, these applications are not very compelling



# 1. Motivation and Problem Statement



- A DMC with an extra random permutation step
  - Draw one of the many possible permutations of  $n$  symbols uniformly
- The order of the received symbols is uninformative
  - Instead of sequences into/out of a channel, we care about types
  - A type tells us everything about a sequence except the ordering

# 1. Motivation and Problem Statement

## Definitions: Channel

- Message  $W \in \{1, 2, \dots, M\}$
- Encoded into codeword  $X^n = f_n(W)$  where  $X \in \mathcal{X} = \{1, 2, \dots, q\}$
- DMC channel input  $X^n$  yields output  $Z^n$ ,  $Z \in \mathcal{Y} = \{1, 2, \dots, k\}$
- Random permutation yields  $Y^n \in \mathcal{Y}^n$
- Decoded into output  $\hat{W} = g_n(Y^n)$
- $P_{\text{error}}^{(n)} = \Pr\{W \neq \hat{W}\}$

Rate is  $R = \frac{\log M}{\log n}$  to account for loss of information from random permutation

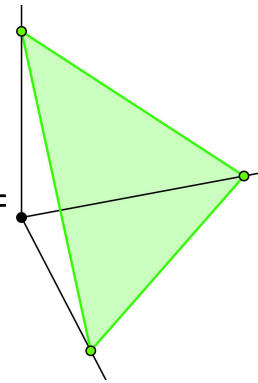
- The number of distinguishable **types** is polynomial in  $n$
- The number of distinguishable **sequences** is exponential in  $n$

# 1. Motivation and Problem Statement

**Definitions (continued):** General

- Probability simplex:

$$\Delta_{q-1} = \left\{ (\pi_1, \pi_2, \dots, \pi_n) : \sum_{i=1}^q \pi_i = 1 \right\}$$



- Method of Types

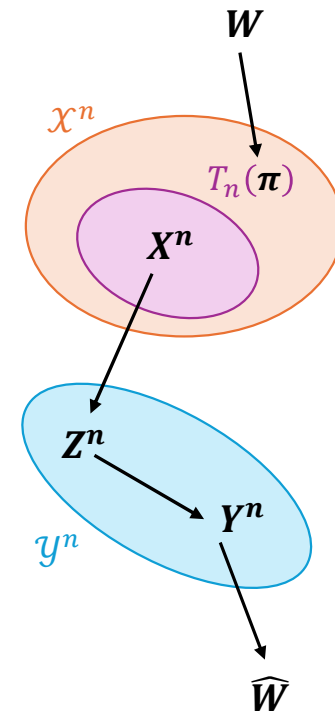
$$\mathcal{P}_n = \left\{ P \in \Delta_{k-1} : P = \left( \frac{a_1}{n}, \frac{a_2}{n}, \dots, \frac{a_n}{n} \right) \text{ where } a_1, a_2, \dots, a_n \in \{0, 1, \dots, n\} \right\}$$

$$T_n(P) = \left\{ x^n \in \mathcal{X}^n : P = \left( \frac{N(a_1|x)}{n}, \frac{N(a_2|x)}{n}, \dots, \frac{N(a_n|x)}{n} \right) \right\}$$

# 1. Motivation and Problem Statement

## Definitions (continued): Channel Distributions

- $P_{Z|X}$  is a  $q \times k$  DMC matrix
  - $(P_{Z|X})_{b,j} = P_{Z|X}(j|b) = p_{bj}$  for  $b \in \mathcal{X}$  and  $j \in \mathcal{Y}$
  - Rows sum to 1
  - $P_{Z|X}$  s.p. (strictly positive) when  $p_{bj} > 0 \forall b \in \mathcal{X}, j \in \mathcal{Y}$
- $P_{Y^n|\pi} = P_{Y|X}^n \circ U$  is the distribution under fixed type  $\pi \in \mathcal{P}_n$ 
  - Where we draw  $U$  uniformly from  $T_n(\pi)$  (A)
  - Then pass independently through the DMC  $P_{Y|X}$
  - Note that  $P_{Z|X}$  and  $P_{Y|X}$  are interchangeable
- Marginal  $P_Y$  for any  $Y_t \in Y^n \sim P_{Y^n|\pi}$ 
  - Marginal is independent of index  $t$  due to (A)



# 1. Motivation and Problem Statement

**Definitions (continued):** Other Distributions:

- $Q_Y$  is any distribution over  $Y$
- $Q_Y^n(y^n) = \prod_{i=1}^n Q_Y(y_i)$  is the product distribution
- Categorical distribution  $Q_{Y|\mu}$  with  $\mu = (\mu_1, \dots, \mu_k) \in \Delta_{k-1}$  with  $Q_{Y|\mu}(j) = \mu_j$
- Multinomial  $Q_{Y|\mu}^n$  has no direct relevance to a noisy permutation, but simplifies analysis

Nats are used, not bits



# 1. Motivation and Problem Statement

## Theorem 1: Main result

For s.p. (strictly positive)  $P_{Y|X}$ , the capacity of the noisy permutation channel is

$$C_{\text{perm}}(P_{Z|X}) = \frac{\text{rank}(P_{Z|X}) - 1}{2}$$

From [Makur, 2020] we know that (for s.p.  $P_{Y|X}$ )

$$C_{\text{perm}}(P_{Z|X}) \geq \frac{\text{rank}(P_{Z|X}) - 1}{2}$$

Thus, we seek a converse (upper) bound to prove theorem 1

# 1. Motivation and Problem Statement

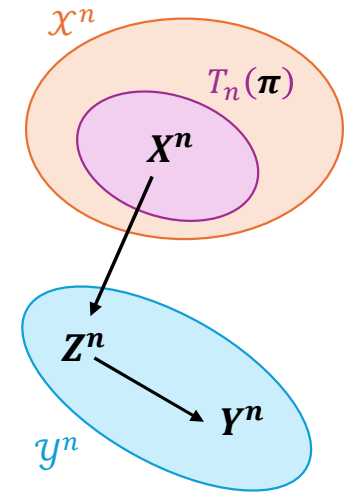
Markov chain is

$$\pi \rightarrow X^n \rightarrow Z^n \rightarrow Y^n$$

for some n-type  $\pi \in \mathcal{P}_n$  chosen by the encoder. Thus, should find the upper bound of  $I(\pi; Y^n)$  to get capacity.

Further, we know

$$I(\pi; Y^n) \leq \max_{\pi} D(P_{Y^n|\pi} \| Q_{Y^n})$$



## Proposition 1: Covering for Noisy Permutation Channel

For a noisy permutation channel with DMC  $P_{Y|X}$  and any n-type  $\pi \in \mathcal{P}_n$ , we **assume** that

$$D(P_{Y^n|\pi} \| Q_{Y^n|\bar{\mu}}) \leq nD(P_Y \| Q_Y) + f(n)$$

for any distribution  $Q_{Y^n}$ . Then,

$$C_{\text{perm}}(P_{Z|X}) \leq \frac{\text{rank}(P_{Z|X}) - 1}{2} + \lim_{n \rightarrow \infty} \frac{f(n)}{\log n}$$

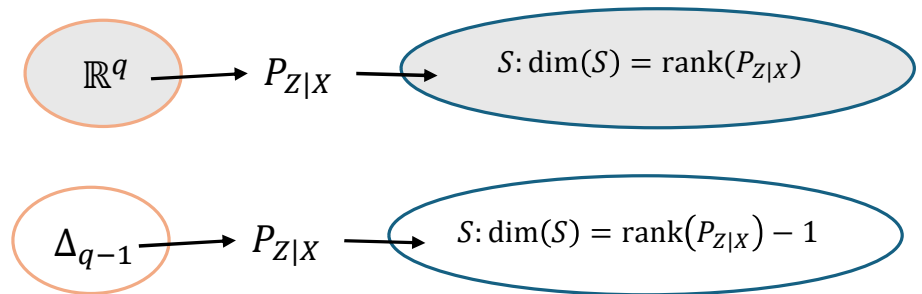
## 2. Covering Converse

### Covering a space of distributions

Consider the space of possible marginals  $P_Y$

$$\mathcal{L}(P_{Y|X}) = \bigcup_{\pi \in \Delta_{k-1}} \left( \sum_i \pi_i p_{i1}, \sum_i \pi_i p_{i2}, \dots, \sum_i \pi_i p_{ik} \right)$$

with dimension  $l = \underbrace{\text{rank}(P_{Z|X})}_{\text{DMC}} - 1$   
Probability vec



## 2. Covering Converse

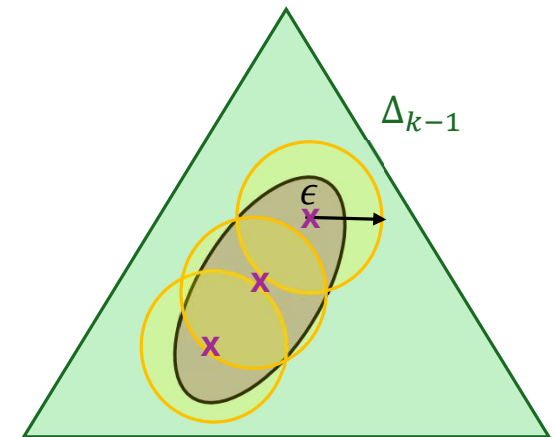
Background:  $\epsilon$ -net covering are a popular tool in CS

### Covering a space of distributions (continued)

Define the  $\epsilon$ -net covering of  $\mathcal{L}(P_{Y|X})$  as  $N_n$  (also, let  $\epsilon = 1/n$ )

$N_n$  is a set of “covering centers”  $\bar{\mu} = (\mu_1, \mu_2, \dots, \mu_k) \in \Delta_{k-1}$ ,  
and not necessarily  $\bar{\mu} \in \mathcal{L}(P_{Y|X})$

**Idea:** use a few vectors to describe a space of vectors



1. Every point  $\pi$  in space  $\mathcal{L}(P_{Y|X})$  is  $\epsilon$ -close to some center  $\bar{\mu}$ , in terms of KLD

$$\max_{\pi \in \mathcal{L}(P_{Y|X})} \min_{\bar{\mu} \in N_n} D(Q_{Y|\mu} \parallel Q_{Y|\bar{\mu}}) \leq \epsilon = \frac{1}{n}$$

## 2. Covering Converse

### Covering a space of distributions (continued)

#### 2. Cardinality of covering has convenient form

$$|N_n| \leq C(q, l)(nl)^{l/2}$$

----- Theorem 4

- $C(q, l)$  is given in the text, but is independent of  $n$  and will be made to vanish
- Reminder:  $\epsilon = 1/n$

Why bother with coverings?

$$I(\pi; Y^n) \leq \log|N_n| + \max_{\pi \in \mathcal{P}_n} \min_{\bar{\mu} \in N_n} D(P_{Y^n|\pi} \| Q_{Y|\bar{\mu}}^n)$$

----- [Yang 1999]

## 2. Covering Converse

$$I(\pi; Y^n) \leq \log|N_n| + \max_{\pi \in \mathcal{P}_n} \min_{\bar{\mu} \in N_n} D(P_{Y^n|\pi} \| Q_{Y|\bar{\mu}}^n) \quad \text{[Yang 1999]}$$

$$\leq \log|N_n| + \max_{\pi \in \mathcal{P}_n} \min_{\bar{\mu} \in N_n} nD(P_Y \| Q_Y) + f(n) \quad \text{Assume (for now)}$$

$$\leq \log|N_n| + f(n) + \max_{\pi \in \mathcal{P}_n} \min_{\bar{\mu} \in N_n} nD(P_Y \| Q_Y)$$

$$\leq \log|N_n| + f(n) + n \frac{1}{n} \quad \text{By covering}$$

$$\leq \log(C(q, l)(nl)^{l/2}) + f(n) + n \frac{1}{n} \quad \text{By covering}$$

$$\leq \log C(q, l) + \frac{l}{2}(\log l + \log n) + f(n) + n \frac{1}{n}$$

$$\leq \frac{l}{2} \log n + c' + f(n)$$

$c'$  contains all terms not depending on  $n$

## 2. Covering Converse

### Proof of proposition 1

$$\log M \leq I(\pi; Y^n) \leq \frac{l}{2} \log n + c' + f(n)$$

By  $R = \frac{\log M}{\log n}$  we have

$$R \leq \frac{l}{2} + \frac{c'}{\log n} + \frac{f(n)}{\log n}$$
$$\lim_{n \rightarrow \infty} R = \frac{l}{2} + \lim_{n \rightarrow \infty} \frac{f(n)}{\log n}$$

Thus

$$C_{\text{perm}}(P_{Z|X}) \leq \frac{\text{rank}(P_{Z|X}) - 1}{2} + \lim_{n \rightarrow \infty} \frac{f(n)}{\log n}$$

QED (Proposition 1)

## 2. Covering Converse

**Proposition 1:** Covering for Noisy Permutation Channel

For a noisy permutation channel with DMC  $P_{Y|X}$  and any  $n$ -type  $\pi \in \mathcal{P}_n$ , we **assume** that

$$D(P_{Y^n|\pi} \| Q_{Y|\bar{\mu}}^n) \leq nD(P_Y \| Q_Y) + f(n) \text{ ----- Next (Theorem 2)}$$

For any distribution  $\tilde{Q}_{Y^n}$ . Then,

$$C_{\text{perm}}(P_{Z|X}) \leq \frac{\text{rank}(P_{Z|X}) - 1}{2} + \lim_{n \rightarrow \infty} \frac{f(n)}{\log n}$$

**Theorem 1:** Main result

For s.p. (strictly positive)  $P_{Y|X}$ , the capacity of the noisy permutation channel is

$$C_{\text{perm}}(P_{Z|X}) = \frac{\text{rank}(P_{Z|X}) - 1}{2}$$



### 3. Divergence under Fixed Types

- Recap **Theorem 2**: When  $P_{Y|X}$  is s.p. (strictly positive),  $\exists c = c(P_{Y|X})$ ,  $\forall P \in \mathcal{P}_n$ , with  $U \sim T_n(P)$  in uniform,  $\forall Q_Y$ :

$$nD(P_Y \parallel Q_Y) \leq D(P_{Y|X}^n \circ U \parallel Q_Y^n) \leq nD(P_Y \parallel Q_Y) + c \quad \leftarrow \text{Final goal}$$

### 3. Divergence under Fixed Types

- **Proposition 3:** Consider  $(X, Y)^n \sim (P \times P_{Y|X})$  in i.i.d, with  $A = 1\{X^n \in T_n(P)\}$ :

$$D(P_{Y|X}^n \circ U \parallel Q_Y^n) = nD(P_Y \parallel Q_Y) + \underbrace{\sum_{y^n \in Y^n} P[Y^n = y^n | A = 1] \cdot \log \frac{P[A = 1 | Y^n = y^n]}{P(A = 1)}}_{\text{②}}$$

- **Proof (1/2):**

- with:  $(P_{Y|X}^n \circ U)(y^n) = P[Y^n = y^n | A = 1]$ ;

- $D(P_{Y|X}^n \circ U \parallel Q_Y^n)$

$$= \sum_{y^n} P[Y^n = y^n | A = 1] \cdot \log \frac{P[Y^n = y^n | A = 1]}{Q_Y^n(y^n)} \quad \text{----- KL-Divergence}$$

$$= \sum_{y^n} P[Y^n = y^n | A = 1] \cdot \log \frac{P[A = 1 | Y^n = y^n] P(Y^n = y^n)}{P(A = 1) Q_Y^n(y^n)} \quad \text{----- Baye's Theorem}$$

$$= \underbrace{\mathbb{E} \left[ \log \frac{P_Y^n(Y^n)}{Q_Y^n(Y^n)} \mid A = 1 \right]}_{\text{①}} + \underbrace{\mathbb{E} \left[ \log \frac{P[\tilde{A} = 1 | \tilde{Y}^n = Y^n]}{P(\tilde{A} = 1)} \mid A = 1 \right]}_{\text{②}}$$

①

②

### 3. Divergence under Fixed Types

- **Proposition 3:** Consider  $(X, Y)^n \sim (P \times P_{Y|X})$  in i.i.d, with  $A = 1\{X^n \in T_n(P)\}$ :

$$D(P_{Y|X}^n \circ U \parallel Q_Y^n) = \underbrace{nD(P_Y \parallel Q_Y)}_{\text{①}} + \underbrace{\sum_{y^n \in Y^n} P[Y^n = y^n \mid A = 1] \cdot \log \frac{P[A = 1 \mid Y^n = y^n]}{P(A = 1)}}_{\text{②}}$$

- **Proof (2/2):**

$$\mathbb{E} \left[ \log \frac{P_Y^n(Y^n)}{Q_Y^n(Y^n)} \mid A = 1 \right]$$

$$= \sum_{y^n} P[Y^n = y^n \mid A = 1] \log \frac{P_Y^n(Y^n)}{Q_Y^n(Y^n)}$$

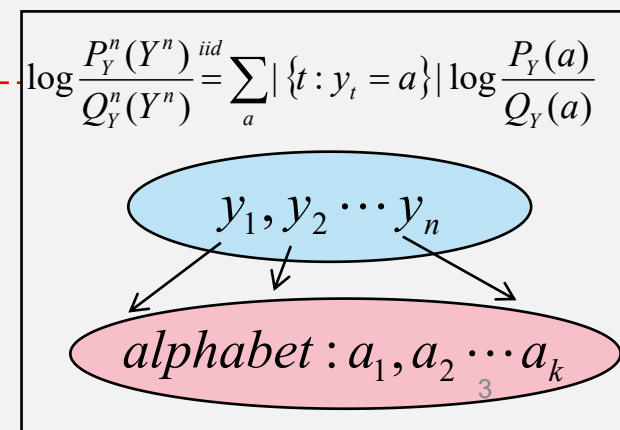
$$= \sum_{y^n} P[Y^n = y^n \mid A = 1] \sum_a n \frac{|\{t : y_t = a\}|}{n} \log \frac{P_Y(a)}{Q_Y(a)}$$

$$= n \sum_a P_Y(a) \log \frac{P_Y(a)}{Q_Y(a)}$$

$$= nD(P_Y \parallel Q_Y)$$

②

----- Expectation



### 3. Divergence under Fixed Types

- Recap **Theorem 2**: When  $P_{Y|X}$  is s.p. (strictly positive),  $\exists c = c(P_{Y|X})$ ,  $\forall P \in \mathcal{P}_n$ , with  $U \sim T_n(P)$  in uniform,  $\forall Q_Y$ :

$$nD(P_Y \parallel Q_Y) \leq D(P_{Y|X}^n \circ U \parallel Q_Y^n) \leq nD(P_Y \parallel Q_Y) + c \quad \leftarrow \text{Final goal}$$

- Recap **Proposition 3**: Consider a noisy permutation channel with DMC  $P_{Y|X}$ , for any  $\pi \in \mathcal{P}_n$ :

$$D(P_{Y|X}^n \circ U \parallel Q_Y^n) = nD(P_Y \parallel Q_Y) + \underbrace{\sum_{y^n \in Y^n} P[Y^n = y^n \mid A = 1] \cdot \log \frac{P[A = 1 \mid Y^n = y^n]}{P(A = 1)}}_{\textcircled{2}}$$

- What's next?
- Equals to proof:

$$\boxed{0 \leq \sum_{y^n \in Y^n} P[Y^n = y^n \mid A = 1] \cdot \log \frac{P[A = 1 \mid Y^n = y^n]}{P(A = 1)} \leq c} \quad \leftarrow \text{Final goal}$$

### 3. Divergence under Fixed Types

- Equalized **Theorem 2**: When  $P_{Y|X}$  is s.p. (strictly positive),  $\exists c = c(P_{Y|X})$ ,  $\forall P \in \mathcal{P}_n$ , with  $U \sim T_n(P)$  in uniform,  $\forall Q_Y$ :

$$0 \leq \sum_{y^n \in Y^n} P[Y^n = y^n | A = 1] \cdot \log \frac{P[A = 1 | Y^n = y^n]}{P(A = 1)} \leq c \quad \leftarrow \text{Final goal}$$

- Let's **prove** the lower bound:

$$\begin{aligned} & \mathbb{E} \left[ \log \frac{P[\tilde{A} = 1 | \tilde{Y}^n = y^n]}{P(\tilde{A} = 1)} \mid A = 1 \right] \\ &= \sum_{y^n} P[Y^n = y^n | A = 1] \cdot \log \frac{P[A = 1 | Y^n = y^n]}{P(A = 1)} \\ &= \sum_{y^n} P[Y^n = y^n | A = 1] \cdot \log \frac{P[Y^n = y^n | A = 1]}{P[Y^n = y^n]} \quad \text{----- Baye's Theorem} \\ &= D(P[Y^n | A = 1] \parallel P[Y^n]) \geq 0 \quad \text{----- KL-Divergence} \end{aligned}$$

- We only need to prove the upper bound for remaining part.

### 3. Divergence under Fixed Types

- Equalized **Theorem 2**: When  $P_{Y|X}$  is s.p. (strictly positive),  $\exists c = c(P_{Y|X})$ ,  $\forall P \in \mathcal{P}_n$ , with  $U \sim T_n(P)$  in uniform,  $\forall Q_Y$ :

$$0 \leq \sum_{y^n \in Y^n} P[Y^n = y^n | A = 1] \cdot \log \frac{P[A = 1 | Y^n = y^n]}{P(A = 1)} \leq c \quad \leftarrow \text{Final goal}$$

- The lower bound has been proved in previous slides:
- To prove the upper bound:

$$\mathbb{E} \left[ \log \frac{P[\tilde{A} = 1 | \tilde{Y}^n = y^n]}{P(\tilde{A} = 1)} \mid A = 1 \right] = \underbrace{\mathbb{E} \left[ \log P[\tilde{A} = 1 | \tilde{Y}^n = y^n] \mid \tilde{A} = 1 \right]}_{\textcircled{1}} - \underbrace{\log P[A = 1]}_{\textcircled{2}}$$

### 3. Divergence under Fixed Types

- **Lemma 2:** Let  $(X, Y)^n \sim (P \times P_{Y|X})$  in i.i.d, with  $A = 1\{X^n \in T_n(P)\}$ , and  $P = (p_1, \dots, p_q) \in \mathcal{P}_n$ :

$$\log \frac{1}{P(A=1)} \leq -\frac{1}{2} \log n + \sum_{i:p_i>0} \frac{1}{2} \log p_i n + \frac{q-1}{2} \log 2\pi + \frac{1}{12n}$$

- **Proof (1/1):**

- Stirling approximation:  $\sqrt{2n\pi} \left(\frac{n}{e}\right)^n e^{\frac{1}{12n+1}} \leq n! \leq \sqrt{2n\pi} \left(\frac{n}{e}\right)^n e^{\frac{1}{12n}}$ ;

$$-\log P(A=1)$$

$$= -\log \left[ \left( \frac{n}{p_1 n, \dots, p_q n} \right) \prod_{i=1}^q p_i^{p_i n} \right]$$

----- Method of Types

$$= -\log \left( \frac{n!}{n^n} \right) - \log \left( \prod_{i=1}^q \frac{(p_i n)^{p_i n}}{(p_i n)!} \right)$$

$$\leq \cancel{n} - \frac{1}{2} \log n - \frac{1}{2} \log 2\pi + \sum_{i=1}^q \left( \cancel{-p_i n} + \frac{1}{2} \log p_i n + \frac{1}{2} \log 2\pi + \frac{q}{12n} \right)$$

----- Stirling approx.

### 3. Divergence under Fixed Types

- Recap equalized **Theorem 2**: When  $P_{Y|X}$  is s.p. (strictly positive),  $\exists c = c(P_{Y|X})$ ,  $\forall P \in \mathcal{P}_n$ , with  $U \sim T_n(P)$  in uniform,  $\forall Q_Y$ :

$$0 \leq \sum_{y^n \in Y^n} P[Y^n = y^n | A = 1] \cdot \log \frac{P[A = 1 | Y^n = y^n]}{P(A = 1)} \leq c \quad \leftarrow \text{Final goal}$$

- From Lemma 2 which we have proved:

$$\log \frac{1}{P(A = 1)} \leq -\frac{1}{2} \log n + \sum_{i: p_i > 0} \frac{1}{2} \log p_i n + \frac{q-1}{2} \log 2\pi + \frac{1}{12n}$$

- We can get:

$$\begin{aligned} c &= \mathbb{E} \left[ \log \frac{P[\tilde{A} = 1 | \tilde{Y}^n = Y^n]}{P(\tilde{A} = 1)} \mid A = 1 \right] \\ &\leq \frac{q-1}{2} \log n + c' + \mathbb{E} \left[ \log P[\tilde{A} = 1 | \tilde{Y}^n = Y^n] \mid A = 1 \right] \\ &\leq \frac{q-1}{2} \log n + c' \end{aligned}$$



### 3. Divergence under Fixed Types

- **Theorem 5 (Petrov 2012):** When  $P(W_i - a_i \leq -\lambda_i/2) \geq b_i, P(W_i - a_i \geq \lambda_i/2) \geq b_i, i = 1, \dots, n, \exists \alpha:$

$$Q(S_n; \lambda) = \sup_z P[z \leq S_n \leq z + \lambda] \leq \alpha \lambda \left( \sum_{i=1}^n \lambda_i^2 b_i \right)^{-1/2}$$

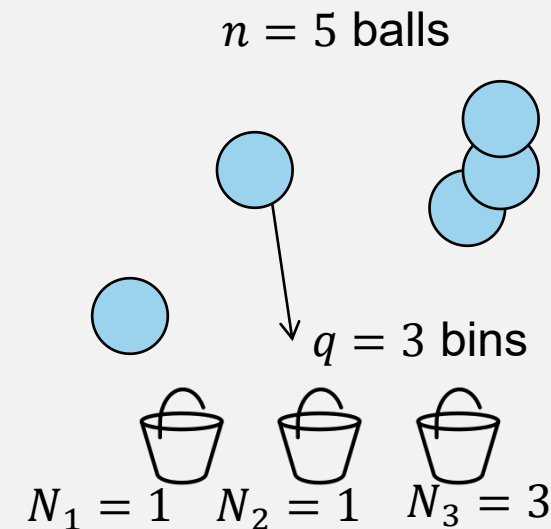
- This provides us a conclusion: Let  $W_i \sim \text{Bernoulli}(p_i), S_n = \sum_{i=1}^n W_i,$

$$P[S_n = z] \leq Q(S_n; 1/2) \leq \frac{\alpha}{\sqrt{\sum_{i=1}^n \min\{p_i, 1 - p_i\}}}$$

- **Lemma 3:** Independently throw  $n$  balls into one of  $q$  bins, we get:

$$P[N_1 = n\pi_1, \dots, N_q = n\pi_q] \leq \frac{\alpha^{q-1}}{n^{(q-1)/2} \sqrt{B}}$$

$$B = c_*^{q-1} \frac{\prod_b \pi_b}{\pi_{\max}} \quad c_* = \min_i \frac{c_-(i)}{c_+(i)} = \min_i \frac{\min_b \{p_{i,b} / \pi_b\}}{\max_b \{p_{i,b} / \pi_b\}}$$



$p_{i,b}$ : Probability of  $i$ -th ball thrown into  $b$ -th bin.

Increasing order

### 3. Divergence under Fixed Types

- A useful bound:

$$\min \left\{ \frac{p_{i,b}}{\sum_{a=b}^q p_{i,a}}, 1 - \frac{p_{i,b}}{\sum_{a=b}^q p_{i,a}} \right\}$$

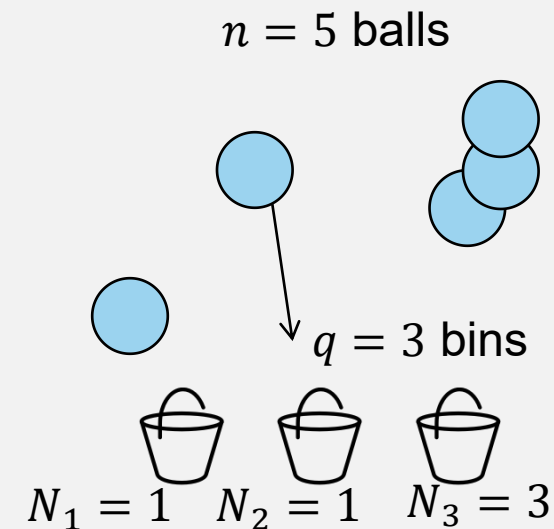
$$= \min \left\{ \frac{\pi_b \frac{p_{i,b}}{\pi_b}}{\sum_{a=b}^q \pi_a \frac{p_{i,a}}{\pi_a}}, \frac{\sum_{a>b}^q \pi_a \frac{p_{i,a}}{\pi_a}}{\sum_{a=b}^q \pi_a \frac{p_{i,a}}{\pi_a}} \right\}$$

$$\geq \frac{\min_a \left\{ \frac{p_{i,a}}{\pi_a} \right\}}{\max_a \left\{ \frac{p_{i,a}}{\pi_a} \right\}} \min \left\{ \frac{\pi_b}{\sum_{a=b}^q \pi_a}, \frac{\sum_{a>b}^q \pi_a}{\sum_{a=b}^q \pi_a} \right\}$$

$$\geq c_* \frac{\pi_b}{\sum_{a=b}^q \pi_a}$$

Increasing order:  $\pi_b \leq \pi_{b+1} \leq \dots$

$$c_* = \min_i \frac{c_-(i)}{c_+(i)} = \min_i \frac{\min_b \{p_{i,b} / \pi_b\}}{\max_b \{p_{i,b} / \pi_b\}}$$



$p_{i,b}$ : Probability of  $i$ -th ball throw into  $b$ -th bin.

Increasing order

$$B = c_*^{q-1} \frac{\prod_b \pi_b}{\pi_{\max}} \quad c_* = \min_i \frac{c_-(i)}{c_+(i)} = \min_i \frac{\min_b \{p_{i,b} / \pi_b\}}{\max_b \{p_{i,b} / \pi_b\}}$$

### 3. Divergence under Fixed Types

- **Lemma 3:** Independently throw  $n$  balls into one of  $q$  bins, we get:

$$P[N_1 = n\pi_1, \dots, N_q = n\pi_q] \leq \frac{\alpha^{q-1}}{n^{(q-1)/2} \sqrt{B}}$$

- **Proof:**  $P[N_1 = n\pi_1, \dots, N_q = n\pi_q]$

$$= \prod_{b=1}^q P[N_b = n\pi_b \mid N_1 = n\pi_1, \dots, N_{b-1} = n\pi_{b-1}]$$

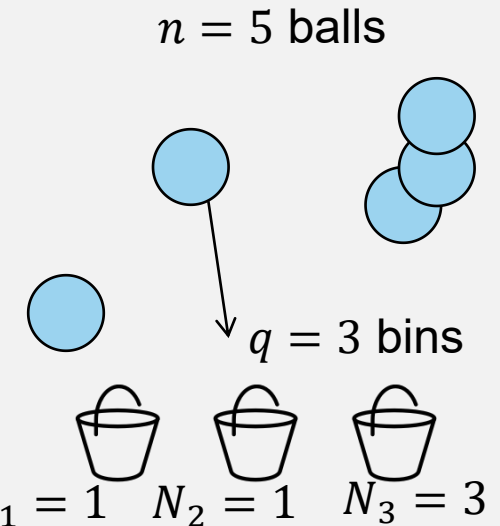
$$\leq \prod_{b=1}^q \frac{\alpha}{\sqrt{\sum_{a=b}^q \min\{p_{i,a}, 1 - p_{ia}\}}}$$

[Petrov 2012]

$$\leq \prod_{b=1}^q \frac{\alpha}{\sqrt{(n - \sum_{a=1}^{b-1} n\pi_a) c_* \frac{\pi_b}{\sum_{a=b}^q \pi_a}}}$$

useful bound

$$= \prod_{b=1}^q \frac{\alpha}{\sqrt{nc_* \pi_b}} = \frac{\alpha^{q-1}}{n^{(q-1)/2} \sqrt{B}}$$



$p_{i,b}$ : Probability of  $i$ -th ball throw into  $b$ -th bin.

Increasing order

### 3. Divergence under Fixed Types

- To prove the upper bound:

$$\begin{aligned} & \textcircled{1} \\ & \overline{\mathbb{E}[\log P[\tilde{A} = 1 \mid \tilde{Y}^n = y^n \mid \tilde{A} = 1]]} \\ & = \log P[N_1 = n\pi_1, \dots, N_{q'} = n\pi_{q'}] \end{aligned}$$

$$\leq \log \frac{\alpha^{q'-1}}{n^{(q'-1)/2} \sqrt{B}}$$

----- Lemma3

$$= \log \left( \frac{\alpha^{q'-1}}{c_*^{(q'-1)/2}} \sqrt{\frac{n\pi_{\max}}{\prod_{b=1}^{q'} n\pi_b}} \right)$$

$$= \frac{1}{2} \log n\pi_{\max} - \sum_{b:\pi_b > 0} \frac{1}{2} \log n\pi_b + (q'-1) \log \frac{\alpha}{\sqrt{c_*}}$$

-----  $\alpha \equiv \text{const}$ ,  $q'$  is number of symbols.

$$\leq \frac{1}{2} \log n - \sum_{b:\pi_b > 0} \frac{1}{2} \log n\pi_b + c'$$

$$c_* = \min_i \frac{c_-(i)}{c_+(i)} = \min_i \frac{\min_b \{p_{i,b} / \pi_b\}}{\max_b \{p_{i,b}^{12} / \pi_b\}}$$

### 3. Divergence under Fixed Types

- Recap **Theorem 2**: When  $P_{Y|X}$  is s.p. (strictly positive),  $\exists c = c(P_{Y|X})$ ,  $\forall P \in \mathcal{P}_n$ , with  $U \sim T_n(P)$  in uniform,  $\forall Q_Y$ :

$$nD(P_Y \parallel Q_Y) \leq D(P_{Y|X}^n \circ U \parallel Q_Y^n) \leq nD(P_Y \parallel Q_Y) + c \quad \leftarrow \text{Proved}$$

- Recap **Proposition 1**: Consider a noisy permutation channel with DMC  $P_{Y|X}$ , for any  $\pi \in \mathcal{P}_n$ , with  $D(P_{Y|X}^n \circ U \parallel Q_Y^n) \leq nD(P_Y \parallel Q_Y) + f(n)$ , we get:

$$C_{perm}(P_{Y|X}) \leq \frac{\text{rank}(P_{Y|X}) - 1}{2} + \lim_{n \rightarrow \infty} \frac{f(n)}{\log n} \quad \leftarrow \text{Proved}$$

- Recap [Makur 2020], that:

$$C_{perm}(P_{Y|X}) \geq \frac{\text{rank}(P_{Y|X}) - 1}{2} \quad \leftarrow \text{Reference}$$

- **Combine** them, we get the main result **Theorem 1**:

$$C_{perm}(P_{Y|X}) = \frac{\text{rank}(P_{Y|X}) - 1}{2}$$

Thank you for your attention!