EM: Expectation Maximization

- 🎯 Goal
 - We have a generative process as follows: heta o Z o X
 - Here, X is observed and Z is latent
 - Estimate $\boldsymbol{\theta}$ that maximizes the likelihood of observed data.

▼ Look at an example to better understand the problem

- Coin toss example:
- Someone has K coins with biases $\theta_{1:K}$... the person picks a coin z_i at random and records the outcome of the toss as x_i , where x_i is either H or T
- N such identical and independent coin toss results are published: $[x_1 \;\; x_2 \;\; ... \;\; x_n]$
- Your goal is to estimate $heta_{1:k}$ that maximizes the likelihood of the published observations, i.e., $argmax \ p(X| heta)$
- In fact, I would like you to also estimate $Z = [z_1, z_2, ..., z_n]$ in the process?
- ▼ Why is this not just a maximum likelihood estimation problem?

Consider the term $\mathop{argmax}\limits_{ heta} p(x_1, x_2, ..., x_n | heta_{1:K}) = \mathop{argmax}\limits_{ heta} \prod_i p(x_i | heta_{1:K})$

- But how do you compute this individual term?
- You cannot because you don't know which coin this x_i came from.
- One possibility is to pretend x_i came from coin θ_j ... and compute the likelihood.
- But then it becomes a huge exponentially growing combinatorics problem since there are K^n possible assignments for n tosses with K coins.
- This is crazy.

But observe that if you can select some possible assignments ... and associate probabilities to them ... you can compute the expectation of the likelihood.

- In other words, since you don't know how to compute the likelihood function, you are setting up the average likelihood over multiple possible assignments.
- Now, optimize that average likelihood.
- That's what EM wants to go towards.
- ▼ The insight in plain language

1. Write out what you need, i.e., $L = p(X, Z|\theta)$... and this breaks into: $L = p(X|Z, \theta)p(Z|\theta)$, i.e., the likelihood and prior. 2. If you pretend to know Z = k, $k \in [1, 2, ..., K]$... then you can compute the likelihood as a function of (Z, θ) ... and the prior can be assumed as equal or some other distribution from domain knowledge. 3. You can simply perform MLE by differentiating w.r.t. θ

4. But you don't know $Z \ \dots$ so what do you do?

5. What if you pretend to know the distribution for Z? Then, you can sample some $Z = z_k$, plug that z_k into the likelihood and prior, and since picking z_k is associated with some probability, you multiply that L with $p(Z = z_k)$ 6. If you do this for each z_k , you are essentially taking the expectation of L over Z. 7. In other words, L can be viewed as the function of the random variable, g(Z), and $p(Z = z_k)$ is the distribution on Z ... and

now, the equation $\sum g(Z)p(Z = z_k)$ forms the expectation of the function of the RV. 8. Since g(Z) is the likelihood, we are actually computing the expectation of the likelihood.

9. Now, how do you get this distribution for Z?

10. Why not compute the posterior distribution for Z based on the observed data and some initial guess of $heta_{1:K}$

11. This won't be right initially ... but it can improve the likelihood L which can then improve the posterior ... which can then improve L ... until convergence.

12. Convergence happens when $\hat{ heta}(t+1) - \hat{ heta}(t) \leq \epsilon$

▼ Set up the likelihood function

- $L = p(X, Z|\theta) = p(X|Z, \theta)p(Z|\theta) = p(x_1, x_2, ..., x_n|z_1, z_2, ..., z_n; \theta_{1:K})p(z_1, z_2, ..., z_n|\theta_{1:K}) = \prod_i p(x_i|z_i; \theta_i)p(z_i|\theta_i)$ - Since θ_i does not influence the choice of $z_i \dots p(z_i|\theta_i) = p(z_i) \dots$ which is the prior for picking z_i . Let's denote this prior with $\pi_{z_i} = p(z_i)$

- The log likelihood is hence: $\log p(X, Z|\theta_{1:K}) = \sum_{i=1}^n \{\log p(x_i|\theta_{z_i}) + \log \pi_{z_i}\}$

- Now, z_i can take any one value between [1:K] ...

- so let's model that by summing over all K possible values but picking only one term from the sum using a DELTA function \rightarrow nice trick

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$$\log p(X, Z|\theta_{1:K}) = \sum_{i=1}^{n} \sum_{k=1}^{K} \delta_k(z_i) \{\log p(x_i|\theta_k) + \log \pi_k\}$$

- Here,
$$\delta_k(z_i) = 1$$
 when $z_i = k$... otherwise, $\delta_k(z_i) = 0$

- Importantly, z_i is the only variable we don't know above.
- ▼ Expectation step → create the posterior distribution

• The posterior is:
$$p(Z|X;\theta) = \prod_{i} p(z_{i}|x_{i};\theta)$$

• Using Bayes rule, we have: $p(Z|X;\theta) = \prod_{i} p(z_{i}|x_{i};\theta) = \prod_{i} \{ p(x_{i}|z_{i};\theta)p(z_{i};\theta) / p(x_{i}) \}$
 $= \prod_{i} \{ p(x_{i}|z_{i};\theta)p(z_{i};\theta) / \sum_{i=1}^{K} p(x_{i}|z_{i};\theta)p(z_{i};\theta)$
• Now, we can shorten this by picking the correct θ based on z_{i} , so we can write this equation as:
 $= \prod_{i} \{ p(x_{i}|\theta_{z_{i}})\pi_{z_{i}} / \sum_{k=1}^{K} p(x_{i}|\theta_{k})\pi_{k} \}$
• Let's call this posterior distribution on variable Z as $q(Z)$... we have $q(Z) = \prod_{i} q_{i}(z_{i})$

ullet Maximization step $\ensuremath{
ightarrow}$ take the expectation of the likelihood over the posterior on Z

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- Recall the log likelihood: $log \ p(X, Z|\theta_{1:K}) = \sum_{i=1}^{n} \sum_{k=1}^{K} \delta_k(z_i) \{log \ p(x_i|\theta_k) + log \ \pi_k\}$... where z_i is the only missing variable.

- Also, we have the posterior on z_i ... and that posterior does not have any missing variable ... so it's a complete distribution. - This means we can get a probability value for $z_i = 1, z_i = 2, ... z_i = K$

- So we can compute the expectation of the log likelihood ... where the expectation is computed over the posterior distribution of z_i , i.e., $q_i(z_i)$

$$E_{q_i}ig[log \ p(X, Z| heta_{1:K})ig] = \sum_{i=1}^n \sum_{k=1}^K \ E_{q_i}ig[\delta_k(z_i) \{ log \ p(x_i| heta_k) + log \ \pi_k \}ig]$$

- Observe that $q_i(z_i)$ is essentially K probabilities values, one for each $z_i, i \in [1,K]$

- Hence, the expectation can be written as the expectation of only the δ functions, since that's the only one with z_i in it

$$egin{aligned} &= \sum_{i=1}^n \sum_{k=1}^K ~~ E_{q_i}ig[\delta_k(z_i)ig]~~ \{log~p(x_i| heta_k)+log~\pi_k\} \ &~~ E_{q_i}ig[\delta_k(z_i)ig] \end{aligned}$$

- Inside the double summation, take the first term, i = 1 and k = 1. This first term's expectation can be written as:

 $K)\delta_1(z_1=K)$

- Now, consider this term:

$$egin{array}{rcl} E_{q_i}ig[\delta_k(z_i)ig] &=& p(z_1=1)1+p(z_1=2)0+...&+p(z_1=K)0\ E_{q_i}ig[\delta_k(z_i)ig] &=& p(z_1=1) \end{array}$$

 $E_{q_i}[\delta_k(z_i)] = p(z_1 = 1)\delta_1(z_1 = 1) + p(z_1 = 2)\delta_1(z_1 = 2) + ... + p(z_1 = 2)$

- See how the expectation for the first term reduced to only a single probability ... since the delta functions zero-forced all other probabilities.

- Now, across all the terms of the double summation (which is a n imes K matrix), we have the following summation:

	k=1	K=2		K = K
i=l	$P(z_1 = 1) \{ f_{i = 1} \} + k_{k = 1}$	$p(\mathcal{Z}_{1}=2)\left\{\begin{array}{c}f_{i=1}\\\kappa^{\leq 2}\end{array}\right\} +$	• • •	+ p(z,=K) { f _{i=1} K=K}
i=2	+ $P(z_{2}=1) \begin{cases} f_{1}=2 \\ k=1 \end{cases}$ +	$p(\overline{z}_{2}, 2) \begin{cases} f_{i=2} \\ k=2 \end{cases} $	•••	+ $p(z_2 = K) \begin{cases} f_{i=2} \\ k = K \end{cases}$
1				
i=n	+ p(zn=1) { f _{i=n} } +	$P(z_{n}=2) \begin{cases} f_{i=n} & 2 \\ k=2 \end{cases}$		+ $p(z_n \in K) \begin{cases} f_{i=n} \\ k \in K \end{cases}$

- Now we can rewrite the expectation of the whole log likelihood function as:

$$E_{q_i}ig[log \ p(X,Z| heta_{1:K})ig] = \sum_{i=1}^n \sum_{k=1}^K \ q_i(k) \ \{ log \ p(x_i| heta_k) + log \ \pi_k \ \}$$

The final step: let's maximize the expected likelihood function.

- Observe we need to maximize the first portion w.r.t. $heta_k, \, k \in [1,K]$... and the second portion w.r.t. π_k
- At the end, we want 2 vectors: $\hat{ heta_{1:K}}$ and prior probabilities, $\pi_{1:K}$

Let's optimize the first term:

- Note that $heta_k$ only occurs in the k^{th} column of the matrix above, so when we optimize for $heta_k$, we sum over all data points i=1:n

$$heta_k = rg\max_{i=1} \, q_i(k) \, log \, p(x_i | heta_k)$$
 .

- This is a standard MLE problem, except that each log term is weighted with a probability value. But differentiation will solve this.

Let's now optimize the second term:

- We basically need to find $[\pi_1, \ \pi_2, \ ... \ \pi_K]$ that maximizes $\ \sum_{i=1}^n \sum_{k=1}^K \ q_i(k) \ log \ \pi_k$

- Note that we can group by column and write this as: $\sum_{i=1}^n q_i(1) \log \pi_1 + \sum_{i=1}^n q_i(2) \log \pi_2 \dots + \sum_{i=1}^n q_i(K) \log \pi_K$

- And since all terms are negative (log of a probability is negative), we can optimize each term individually while satisfying the constraints that:

$$\sum_{k=1}^{\kappa} \pi_k = 1$$
 and $\pi_k \geq 0, \,\, orall k = 1, 2, ..., K$

- So finding the optimal $\pi_{1:K}$ is equivalent to:

$$\hat{\pi}_{1:K} = rg\max \ \sum_{i=1}^n \ q_i(k) \ log \ \pi_k$$
 s.t. $\sum_{k=1}^K \ \pi_k = 1, \ \pi_k \geq 0, \ orall k = 1, 2, ..., K$

- This is a straight application of Lagrange multipliers \rightarrow and we get the optimal π_k as the average of the $q_i(k)$ coefficients:

$$\hat{\pi}_k = rac{1}{n} \sum_{i=1}^n q_i(k)$$

▼ The final iteration

Putting everything together:

- E step: The posterior $q_i(z_i)$ is a function of a single data point x_i ... all the $\theta_{1:K}$ calculated till now ... and all the $\pi_{1:K}$ estimated till now

- M step: The expected log-likelihood needs to be calculated ... where the log-likelihood depends on all data samples $x_{1:n}$... all the $\theta_{1:K}$ calculated till now ... and all the $\pi_{1:K}$
- estimated till now

and the expectation over the posterior also needs $q_i(z_i)$

- So, start the E^0 step with an initial guess on $[heta_{1:K}^0, \ \pi_{1:K}^0]$... and compute $q_i^0(z_i)$
- Then, compute $~~[heta_{1:K}^1,~\pi_{1:K}^1]$ using $~q_i^0(z_i)$ and $[heta_{1:K}^0,~\pi_{1:K}^0]$

Then iterate as:
$$[\theta_{1:K}^1, \pi_{1:K}^1] \rightarrow q_i^1 \rightarrow [\theta_{1:K}^2, \pi_{1:K}^2] \rightarrow q_i^2 \dots [\theta_{1:K}^t, \pi_{1:K}^t] \rightarrow q_i^{t+1}$$

- Terminate when $| heta^t - heta^{t-1}| \leq \epsilon$

Tutorial by Dahua Lin (MIT): here